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**ECON 351\* -- NOTE 3**
**Desirable Statistical Properties of Estimators**
**1. Introduction**

We derived in *Note 2* the OLS (Ordinary Least Squares) estimators  $\hat{\beta}_j$  ( $j = 1, 2$ ) of the regression coefficients  $\beta_j$  ( $j = 1, 2$ ) in the simple linear regression model given by the **population regression equation**, or PRE

$$Y_i = \beta_1 + \beta_2 X_i + u_i \quad (1)$$

where  $u_i$  is an iid random error term.

The **OLS estimators of the regression coefficients  $\beta_1$  and  $\beta_2$**  are:

$$\hat{\beta}_2 = \frac{\sum_i x_i y_i}{\sum_i x_i^2} = \frac{\sum_i (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_i (X_i - \bar{X})^2} = \frac{\sum_i X_i Y_i - N\bar{X}\bar{Y}}{\sum_i X_i^2 - N\bar{X}^2}. \quad (2.2)$$

$$\hat{\beta}_1 = \bar{Y} - \hat{\beta}_2 \bar{X}. \quad (2.1)$$

where

$$y_i \equiv Y_i - \bar{Y} \quad (i = 1, \dots, N)$$

$$x_i \equiv X_i - \bar{X} \quad (i = 1, \dots, N)$$

$$\bar{Y} = \sum_i Y_i / N = \frac{\sum_i Y_i}{N} = \text{the sample mean of the } Y_i \text{ values.}$$

$$\bar{X} = \sum_i X_i / N = \frac{\sum_i X_i}{N} = \text{the sample mean of the } X_i \text{ values.}$$

The reason we use these **OLS coefficient estimators** is that, **under assumptions A1-A8 of the classical linear regression model**, they **have a number of desirable statistical properties**.

*Note 4* outlines these desirable statistical properties of the OLS coefficient estimators.

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**Why are statistical properties of estimators important?** These statistical properties are extremely important because they provide **criteria for choosing among alternative estimators**. Knowledge of these properties is therefore essential to understanding why we use the estimation procedures we do in econometric analysis.

## 2. Distinction Between an "Estimator" and an "Estimate"

Recall the distinction between an **estimator** and an **estimate**. Let some unknown population parameter be denoted in general as  $\theta$ , and let an *estimator of  $\theta$*  be denoted as  $\hat{\theta}$ .

## 3. Two Categories of Statistical Properties

There are *two categories of statistical properties*.

### (1) *Small-sample, or finite-sample, properties;*

The most fundamental *desirable small-sample properties* of an estimator are:

- S1. Unbiasedness;**
- S2. Minimum Variance;**
- S3. Efficiency.**

### (2) *Large-sample, or asymptotic, properties.*

The most important *desirable large-sample property* of an estimator is:

- L1. Consistency.**

Both sets of statistical properties refer to the **properties of the *sampling distribution, or probability distribution, of the estimator  $\hat{\theta}$  for different sample sizes.***

### 3.1 Small-Sample (Finite-Sample) Properties

- The *small-sample, or finite-sample, properties* of the estimator  $\hat{\theta}$  refer to the **properties of the sampling distribution of  $\hat{\theta}$  for any sample of fixed size  $n$** , where  $n$  is a *finite number* (i.e., a number less than infinity) denoting the number of observations in the sample.

*$n = \text{number of sample observations, where } n < \infty.$*

**Definition:** The **sampling distribution of  $\hat{\theta}$  for any finite sample size  $n < \infty$**  is called the *small-sample, or finite-sample, distribution of the estimator  $\hat{\theta}$* . In fact, there is a family of finite-sample distributions for the estimator  $\hat{\theta}$ , one for each finite value of  $n$ .

- The **sampling distribution of  $\hat{\theta}$**  is based on the **concept of repeated sampling**.
- Suppose a **large number of samples of size  $n$**  are randomly selected from some underlying population.
    - ◆ Each of these samples **contains  $n$  observations**.
    - ◆ Each of these samples in general **contains different sample values** of the observable random variables that enter the formula for the estimator  $\hat{\theta}$ .
  - **For each of these samples of  $n$  observations**, the **formula for  $\hat{\theta}$**  is used to **compute a numerical estimate** of the population parameter  $\theta$ .
    - ◆ Each sample yields a **different numerical estimate of the unknown parameter  $\theta$** .  
Why? Because each sample typically contains different sample values of the observable random variables that enter the formula for the estimator  $\hat{\theta}$ .
  - If we tabulate or plot these different sample estimates of the parameter  $\theta$  for a very large number of samples of size  $n$ , we obtain the *small-sample, or finite-sample, distribution of the estimator  $\hat{\theta}$* .

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## 3.2 Large-Sample (Asymptotic) Properties

- The *large-sample, or asymptotic, properties of the estimator  $\hat{\theta}$*  refer to the **properties of the sampling distribution of  $\hat{\theta}$  as the sample size  $n$  becomes indefinitely large**, i.e., as sample size  $n$  approaches infinity (**as  $n \rightarrow \infty$** ).

**Definition:** The probability distribution to which the sampling distribution of  $\hat{\theta}$  converges *as sample size  $n$  becomes indefinitely large* (i.e., as  $n \rightarrow \infty$ ) is called the *asymptotic, or large-sample, distribution of the estimator  $\hat{\theta}$* .

- The **properties of the asymptotic distribution of  $\hat{\theta}$**  are what we call the *large-sample, or asymptotic, properties of the estimator  $\hat{\theta}$* .
- To define all the *large-sample properties* of an estimator, we need to distinguish between *two large-sample distributions* of an estimator:
  1. the *asymptotic distribution* of the estimator;
  2. the *ultimate, or final, distribution* of the estimator.

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## 4. Small-Sample Properties

### Nature of Small-Sample Properties

- The *small-sample, or finite-sample, distribution* of the estimator  $\hat{\theta}$  for any finite sample size  $n < \infty$  has
  1. a *mean*, or expectation, denoted as  $E(\hat{\theta})$ , and
  2. a *variance* denoted as  $\text{Var}(\hat{\theta})$ .
  
- The *small-sample properties of the estimator  $\hat{\theta}$*  are defined in terms of the *mean*  $E(\hat{\theta})$  and the *variance*  $\text{Var}(\hat{\theta})$  of the *finite-sample distribution of the estimator  $\hat{\theta}$* .

## **S1: Unbiasedness**

**Definition of Unbiasedness:** The estimator  $\hat{\theta}$  is an *unbiased estimator* of the population parameter  $\theta$  if the mean or expectation of the finite-sample distribution of  $\hat{\theta}$  is equal to the true  $\theta$ . That is,  $\hat{\theta}$  is an *unbiased estimator of  $\theta$*  if

$$E(\hat{\theta}) = \theta \quad \text{for any given finite sample size } n < \infty.$$

**Definition of the Bias of an Estimator:** The *bias of the estimator  $\hat{\theta}$*  is defined as

$$\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta = \text{the mean of } \hat{\theta} \text{ minus the true value of } \theta.$$

- The estimator  $\hat{\theta}$  is an *unbiased estimator* of the population parameter  $\theta$  if the bias of  $\hat{\theta}$  is equal to zero; i.e., if

$$\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta = 0 \quad \Leftrightarrow \quad E(\hat{\theta}) = \theta.$$

- Alternatively, the estimator  $\hat{\theta}$  is a *biased estimator* of the population parameter  $\theta$  if the bias of  $\hat{\theta}$  is non-zero; i.e., if

$$\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta \neq 0 \quad \Leftrightarrow \quad E(\hat{\theta}) \neq \theta.$$

1. The estimator  $\hat{\theta}$  is an *upward biased* (or *positively biased*) estimator of the population parameter  $\theta$  if the **bias of  $\hat{\theta}$  is greater than zero**; i.e., if

$$\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta > 0 \quad \Leftrightarrow \quad E(\hat{\theta}) > \theta.$$

2. The estimator  $\hat{\theta}$  is a *downward biased* (or *negatively biased*) estimator of the population parameter  $\theta$  if the **bias of  $\hat{\theta}$  is less than zero**; i.e., if

$$\text{Bias}(\hat{\theta}) = E(\hat{\theta}) - \theta < 0 \quad \Leftrightarrow \quad E(\hat{\theta}) < \theta.$$

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### Meaning of the Unbiasedness Property

- The estimator  $\hat{\theta}$  is an unbiased estimator of  $\theta$  if on average it equals the true parameter value  $\theta$ .
  - This means that on average the estimator  $\hat{\theta}$  is correct, even though any single estimate of  $\theta$  for a particular sample of data may not equal  $\theta$ .
  - More technically, it means that **the finite-sample distribution of the estimator  $\hat{\theta}$  is *centered* on the value  $\theta$** , not on some other real value.
- The ***bias of an estimator*** is an ***inverse measure*** of its **average accuracy**.
  - The smaller in absolute value is  $\text{Bias}(\hat{\theta})$ , the more accurate on average is the estimator  $\hat{\theta}$  in estimating the population parameter  $\theta$ .
  - Thus, an unbiased estimator for which  $\text{Bias}(\hat{\theta}) = 0$  -- that is, for which  $E(\hat{\theta}) = \theta$  -- is *on average* a perfectly accurate estimator of  $\theta$ .
- Given a choice between two estimators of the same population parameter  $\theta$ , of which one is biased and the other is unbiased, we prefer the unbiased estimator because it is more accurate on average than the biased estimator.

## **S2: Minimum Variance**

**Definition of Minimum Variance:** The estimator  $\hat{\theta}$  is a *minimum-variance estimator* of the population parameter  $\theta$  if the variance of the finite-sample distribution of  $\hat{\theta}$  is *less than or equal to* the variance of the finite-sample distribution of  $\tilde{\theta}$ , where  $\tilde{\theta}$  is *any other estimator* of the population parameter  $\theta$ ; i.e., if

$$\text{Var}(\hat{\theta}) \leq \text{Var}(\tilde{\theta}) \text{ for all finite sample sizes } n \text{ such that } 0 < n < \infty$$

where

$$\text{Var}(\hat{\theta}) = E[\hat{\theta} - E(\hat{\theta})]^2 = \text{the variance of the estimator } \hat{\theta};$$

$$\text{Var}(\tilde{\theta}) = E[\tilde{\theta} - E(\tilde{\theta})]^2 = \text{the variance of any other estimator } \tilde{\theta}.$$

**Note:** Either or both of the estimators  $\hat{\theta}$  and  $\tilde{\theta}$  may be biased. The minimum variance property implies nothing about whether the estimators are biased or unbiased.

### **Meaning of the Minimum Variance Property**

- The *variance of an estimator* is an *inverse measure* of its **statistical precision**, i.e., of its dispersion or spread around its mean.

The *smaller the variance* of an estimator, the *more statistically precise* it is.

- A *minimum variance estimator* is therefore the statistically *most precise estimator* of an unknown population parameter, although it may be biased or unbiased.

### **S3: Efficiency**

#### **A Necessary Condition for Efficiency -- Unbiasedness**

The small-sample property of efficiency is defined only for *unbiased estimators*.

Therefore, a *necessary condition for efficiency of the estimator*  $\hat{\theta}$  is that  $E(\hat{\theta}) = \theta$ , i.e.,  $\hat{\theta}$  must be an *unbiased estimator of the population parameter*  $\theta$ .

#### **Definition of Efficiency: Efficiency = Unbiasedness + Minimum Variance**

*Verbal Definition:* If  $\hat{\theta}$  and  $\tilde{\theta}$  are two unbiased estimators of the population parameter  $\theta$ , then the estimator  $\hat{\theta}$  is efficient relative to the estimator  $\tilde{\theta}$  if the variance of  $\hat{\theta}$  is smaller than the variance of  $\tilde{\theta}$  for any finite sample size  $n < \infty$ .

*Formal Definition:* Let  $\hat{\theta}$  and  $\tilde{\theta}$  be two *unbiased estimators of the population parameter*  $\theta$ , such that  $E(\hat{\theta}) = \theta$  and  $E(\tilde{\theta}) = \theta$ . Then **the estimator  $\hat{\theta}$  is efficient relative to the estimator  $\tilde{\theta}$**  if the variance of the finite-sample distribution of  $\hat{\theta}$  is less than or at most equal to the variance of the finite-sample distribution of  $\tilde{\theta}$ ; i.e. if

$$\text{Var}(\hat{\theta}) \leq \text{Var}(\tilde{\theta}) \text{ for all finite } n \text{ where } E(\hat{\theta}) = \theta \text{ and } E(\tilde{\theta}) = \theta.$$

*Note:* Both the estimators  $\hat{\theta}$  and  $\tilde{\theta}$  must be *unbiased*, since the efficiency property refers only to the variances of unbiased estimators.

#### **Meaning of the Efficiency Property**

- Efficiency is a desirable statistical property because of two unbiased estimators of the same population parameter, we prefer the one that has the smaller variance, i.e., the one that is statistically more precise.
- In the above definition of efficiency, **if  $\tilde{\theta}$  is any other unbiased estimator of the population parameter  $\theta$ , then the estimator  $\hat{\theta}$  is the best unbiased, or minimum-variance unbiased, estimator of  $\theta$ .**

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## 5. Large-Sample Properties

### Nature of Large-Sample Properties

- The large-sample properties of an estimator are the properties of the sampling distribution of that estimator **as sample size  $n$  becomes very large**, as  $n$  approaches infinity, **as  $n \rightarrow \infty$** .
- Recall that **the *sampling distribution* of an estimator differs for different sample sizes** -- i.e., **for different values of  $n$** .

The sampling distribution of a given estimator for one sample size is different from the sampling distribution of that same estimator for some other sample size.

Consider the estimator  $\hat{\theta}$  for **two different values of  $n$ ,  $n_1$  and  $n_2$** .

In general, **the sampling distributions of  $\hat{\theta}_{n_1}$  and  $\hat{\theta}_{n_2}$  are *different***: they can have

- ***different* means**
- ***different* variances**
- ***different* mathematical forms**

### Desirable Large-Sample Properties

- L1. Consistency;**
- L2. Asymptotic Unbiasedness;**
- L3. Asymptotic Efficiency.**

We need to distinguish between ***two large-sample distributions*** of an estimator:

- 1. the *asymptotic distribution* of the estimator;**
- 2. the *ultimate, or final, distribution* of the estimator.**

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## The Ultimate Distribution of an Estimator

- The asymptotic distribution of an estimator  $\hat{\theta}$  is not necessarily the final or ultimate form that the sampling distribution of the estimator takes as sample size  $n$  approaches infinity (as  $n \rightarrow \infty$ ).
- For many estimators, **the sampling distribution collapses to a single point as sample size  $n$  approaches infinity.**
  - More specifically, **as sample size  $n \rightarrow \infty$** , the sampling distribution of the estimator **collapses to a column of unit probability mass** (or unit probability density) **at a single point** on the real line.
  - Such an estimator is said to **converge in probability** to some value.
  - A distribution that is completely concentrated at one point (on the real line) is called a **degenerate distribution**. Graphically, a degenerate distribution is represented by **a perpendicular to the real line with height equal to one.**

## Distinguishing Between an Estimator's Asymptotic and Ultimate Distributions

- The **ultimate distribution** of an estimator is the distribution to which the sampling distribution of the estimator finally converges as sample size  $n$  "reaches" infinity.
- The **asymptotic distribution** of an estimator is the distribution to which the sampling distribution of the estimator converges just before it collapses (if indeed it does) as sample size  $n$  "approaches" infinity.
  - **If the ultimate distribution of an estimator is degenerate**, then its **asymptotic distribution is not identical** to its **ultimate (final) distribution.**
  - **But if the ultimate distribution of an estimator is non-degenerate**, then its **asymptotic distribution is identical** to its **ultimate (final) distribution.**

## **L1: Consistency**

### **A Necessary Condition for Consistency**

Let  $\hat{\theta}_n$  be an estimator of the population parameter  $\theta$  based on a sample of size  $n$  observations.

A *necessary condition for consistency* of the estimator  $\hat{\theta}_n$  is that **the ultimate distribution of  $\hat{\theta}_n$  be a degenerate distribution** at some point on the real line, meaning that it converges to a single point on the real line.

### **The Probability Limit of an Estimator** -- the **plim** of an estimator

**Short Definition:** The probability limit of an estimator  $\hat{\theta}_n$  is the value -- or point on the real line -- to which that estimator's sampling distribution converges as sample size  $n$  increases without limit.

**Long Definition:** If the estimator  $\hat{\theta}_n$  has a *degenerate ultimate distribution* -- i.e., if the sampling distribution of  $\hat{\theta}_n$  collapses on a single point as sample size  $n \rightarrow \infty$  -- then **that point on which the sampling distribution of  $\hat{\theta}_n$  converges** is called the *probability limit* of  $\hat{\theta}_n$ , and is denoted as  $\text{plim } \hat{\theta}_n$  or  $\text{plim}_{n \rightarrow \infty} \hat{\theta}_n$  where "plim" means "probability limit".

**Formal Definition of Probability Limit:** The point  $\theta_0$  on the real line is the **probability limit of the estimator  $\hat{\theta}_n$**  if the **ultimate sampling distribution of  $\hat{\theta}_n$**  is **degenerate at the point  $\theta_0$** , meaning that the sampling distribution of  $\hat{\theta}_n$  collapses to a column of unit density on the point  $\theta_0$  as sample size  $n \rightarrow \infty$ .

- This definition can be written concisely as

$$\text{plim } \hat{\theta} = \theta_0 \quad \text{or} \quad \text{plim}_{n \rightarrow \infty} \hat{\theta}_n = \theta_0.$$

- The three statements above are equivalent to the statement

$$\lim_{n \rightarrow \infty} \Pr(\theta_0 - \varepsilon \leq \hat{\theta}_n \leq \theta_0 + \varepsilon) = \lim_{n \rightarrow \infty} \Pr(-\varepsilon \leq \hat{\theta}_n - \theta_0 \leq +\varepsilon) = \lim_{n \rightarrow \infty} \Pr(|\hat{\theta}_n - \theta_0| \leq \varepsilon) = 1$$

where  $\varepsilon > 0$  is an arbitrarily small positive number and  $|\hat{\theta}_n - \theta_0|$  denotes the absolute value of the difference between the estimator  $\hat{\theta}_n$  and the point  $\theta_0$ .

1. This statement means that as sample size  $n \rightarrow \infty$ , the probability that the estimator  $\hat{\theta}_n$  is arbitrarily close to the point  $\theta_0$  approaches one.
2. The more technical way of saying this is that **the estimator  $\hat{\theta}_n$  converges in probability to the point  $\theta_0$** .

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## Definition of Consistency

**Verbal Definition:** The estimator  $\hat{\theta}_n$  is a *consistent estimator* of the population parameter  $\theta$  if its sampling distribution collapses on, or converges to, the value of the population parameter  $\theta$  as  $n \rightarrow \infty$ .

**Formal Definition:** The estimator  $\hat{\theta}_n$  is a *consistent estimator* of the population parameter  $\theta$  if the *probability limit of  $\hat{\theta}_n$  is  $\theta$* , i.e., if

$$\text{plim}_{n \rightarrow \infty} \hat{\theta}_n = \theta \quad \text{or} \quad \lim_{n \rightarrow \infty} \Pr\left(|\hat{\theta}_n - \theta| \leq \varepsilon\right) = 1 \quad \text{or} \quad \Pr\left(|\hat{\theta}_n - \theta| \leq \varepsilon\right) \rightarrow 1 \text{ as } n \rightarrow \infty.$$

The estimator  $\hat{\theta}_n$  is a *consistent estimator* of the population parameter  $\theta$

- if the probability that  $\hat{\theta}_n$  is arbitrarily close to  $\theta$  approaches 1 as the sample size  $n \rightarrow \infty$

or

- if the estimator  $\hat{\theta}_n$  converges in probability to the population parameter  $\theta$ .

## Meaning of the Consistency Property

- As sample size  $n$  becomes larger and larger, the sampling distribution of  $\hat{\theta}_n$  becomes more and more concentrated around  $\theta$ .
- As sample size  $n$  becomes larger and larger, the value of  $\hat{\theta}_n$  is more and more likely to be very close to  $\theta$ .

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### A Sufficient Condition for Consistency

One way of determining if the estimator  $\hat{\theta}_n$  is consistent is to trace the behavior of the sampling distribution of  $\hat{\theta}_n$  as sample size  $n$  becomes larger and larger.

- If as  $n \rightarrow \infty$  (sample size  $n$  approaches infinity) both the **bias** of  $\hat{\theta}_n$  and the **variance** of  $\hat{\theta}_n$  *approach zero*, then  $\hat{\theta}_n$  is a **consistent estimator of the parameter  $\theta$** .
- Recall that the bias of  $\hat{\theta}_n$  is defined as

$$\text{Bias}(\hat{\theta}_n) = E(\hat{\theta}_n) - \theta.$$

Thus, the bias of  $\hat{\theta}_n$  approaches zero as  $n \rightarrow \infty$  if and only if the mean or expectation of the sampling distribution of  $\hat{\theta}_n$  approaches  $\theta$  as  $n \rightarrow \infty$ :

$$\lim_{n \rightarrow \infty} \text{Bias}(\hat{\theta}_n) = \lim_{n \rightarrow \infty} E(\hat{\theta}_n) - \theta = 0 \quad \Leftrightarrow \quad \lim_{n \rightarrow \infty} E(\hat{\theta}_n) = \theta.$$

- **Result:** A *sufficient condition for consistency of the estimator  $\hat{\theta}_n$*  is that

$$\lim_{n \rightarrow \infty} \text{Bias}(\hat{\theta}_n) = 0 \quad \text{or} \quad \lim_{n \rightarrow \infty} E(\hat{\theta}_n) = \theta \quad \text{and} \quad \lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}_n) = 0.$$

This condition states that **if both the bias and variance of the estimator  $\hat{\theta}_n$  approach zero as sample size  $n \rightarrow \infty$ , then  $\hat{\theta}_n$  is a consistent estimator of  $\theta$** .