

Similarly, when we test  $H_0$  against  $H_2$ , the NCP is

$$\begin{aligned} A_{21} &= \frac{\rho_0^2}{\sigma_0^2} \text{plim}_{n \rightarrow \infty} \left( \frac{1}{n} \mathbf{u}_{-1}^\top \mathbf{M}_X (\mathbf{X}_{-1} \boldsymbol{\beta}_0 + \mathbf{u}_{-1}) \right) \\ &\quad \times \text{plim}_{n \rightarrow \infty} \left( \frac{1}{n} (\mathbf{X}_{-1} \boldsymbol{\beta}_0 + \mathbf{u}_{-1})^\top \mathbf{M}_X (\mathbf{X}_{-1} \boldsymbol{\beta}_0 + \mathbf{u}_{-1}) \right)^{-1} \\ &\quad \times \text{plim}_{n \rightarrow \infty} \left( \frac{1}{n} (\mathbf{X}_{-1} \boldsymbol{\beta}_0 + \mathbf{u}_{-1})^\top \mathbf{M}_X \mathbf{u}_{-1} \right). \end{aligned}$$

This simplifies to

$$\begin{aligned} &\frac{\rho_0^2}{\sigma_0^2} \sigma_0^2 \left( \sigma_0^2 + \text{plim}_{n \rightarrow \infty} \frac{1}{n} \|\mathbf{M}_X \mathbf{X}_{-1} \boldsymbol{\beta}_0\|^2 \right)^{-1} \sigma_0^2 \\ &= \rho_0^2 \left( 1 + \sigma_0^{-2} \text{plim}_{n \rightarrow \infty} \frac{1}{n} \|\mathbf{M}_X \mathbf{X}_{-1} \boldsymbol{\beta}_0\|^2 \right)^{-1}. \end{aligned}$$

Evidently,  $\cos^2 \phi$  for the test of  $H_0$  against  $H_2$  is the right-hand expression here divided by  $\rho_0^2$ , which is

$$\left( 1 + \frac{\text{plim}_{n \rightarrow \infty} n^{-1} \|\mathbf{M}_X \mathbf{X}_{-1} \boldsymbol{\beta}_0\|^2}{\sigma_0^2} \right)^{-1}. \quad (12.34)$$

This last result is worth comment. We have found that  $\cos^2 \phi$  for the test against  $H_2$  when the data were generated by  $H_1$ , expression (12.34), is identical to  $\cos^2 \phi$  for the test against  $H_1$  when the data were generated by  $H_2$ , expression (12.33). This result is true not just for this example, but for every case in which both alternatives involve one-degree-of-freedom tests. Geometrically, this equivalence simply reflects the fact that when  $\mathbf{z}$  is a vector, the angle between  $\alpha n^{-1/2} \mathbf{M}_X \mathbf{a}$  and the projection of  $\alpha n^{-1/2} \mathbf{M}_X \mathbf{a}$  onto  $\mathcal{S}(\mathbf{X}, \mathbf{z})$ , which is

$$\alpha n^{-1/2} \mathbf{M}_X \mathbf{z} (\mathbf{z}^\top \mathbf{M}_X \mathbf{z})^{-1} \mathbf{z}^\top \mathbf{M}_X \mathbf{a},$$

is the same as the angle between  $\alpha n^{-1/2} \mathbf{M}_X \mathbf{a}$  and  $\alpha n^{-1/2} \mathbf{M}_X \mathbf{z}$ . The reason for this is that  $(\mathbf{z}^\top \mathbf{M}_X \mathbf{z})^{-1} \mathbf{z}^\top \mathbf{M}_X \mathbf{a}$  is a scalar when  $\mathbf{z}$  is a vector. Hence, if we reverse the roles of  $\mathbf{a}$  and  $\mathbf{z}$ , the angle is unchanged. This geometrical fact also results in two numerical facts. First, in the regressions

$$\mathbf{y} = \mathbf{X} \boldsymbol{\alpha} + \gamma \mathbf{z} + \text{residuals} \quad \text{and}$$

$$\mathbf{z} = \mathbf{X} \boldsymbol{\beta} + \delta \mathbf{y} + \text{residuals},$$

the  $t$  statistic on  $\mathbf{z}$  in the first is equal to that on  $\mathbf{y}$  in the second. Second, in the regressions

$$\mathbf{M}_X \mathbf{y} = \gamma \mathbf{M}_X \mathbf{z} + \text{residuals} \quad \text{and}$$

$$\mathbf{M}_X \mathbf{z} = \delta \mathbf{M}_X \mathbf{y} + \text{residuals},$$

the  $t$  statistics on  $\gamma$  and  $\delta$  are numerically identical and so are the uncentered  $R^2$ 's.