# Online Appendix for "Estimating the U.S. Output Gap with State-Level Data" 

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## Board of Governors of the Federal Reserve System

[^0]
## Online Appendix

## A Contacts with the Literature

This work is related to the literature on trend-cycle decompositions using UC models. It also relates to the literature on common components estimation with dynamic factor models.

Introducing multidimensional variability into the UC model's setup can be done in three ways. One is to incorporate many variables belonging to one unit of interest to extract common trends and/or cycles. Another possibility is to incorporate many units of interest that use the same variable to extract common trends, cycles, or both. The third possibility considers a combination of many units of interest and many variables to extract common trends, cycles, or both. In this paper, I adopt the third possibility.

## A. 1 Many Variables and a Single Economic Unit

Crone and Clayton-Matthews (2005) employ the UC model discussed in Stock and Watson $(1988,1989)$ to describe how to use mixed frequency data from the U.S. states to estimate statelevel monthly indexes of economic activity for each state. Observable variables for each state are the first difference of the following: (the $\log$ of) nonagricultural employment, the unemployment rate, (the $\log$ of) average hours worked in manufacturing, and (the log of) real wage and salary disbursements. A scalar latent stationary series is common to the state-level observable variables and is interpreted as the state's cycle. The observable series load on the state's cycle with leads and/or lags. In this analysis, only the cross-sectional structure of the many variables is exploited, not the variability along the state-level dimension.

Basistha and Startz (2008) and Fleischman and Roberts (2011) also use several variables, but at the aggregate level, to estimate the U.S. NAIRU and the U.S. potential output and its associated business cycle, respectively. The authors of both papers emphasize the advantages of using a multivariate approach in the estimation of UC models, including better estimate precision and the ability to coherently assess the trade-offs of competing signals.

## A. 2 A Single Variable and Many Economic Units

Kouparitsas $(2002,2001)$ specifies a UC model to decompose U.S. regional per capita income fluctuations into their trend and cycle components. Using data from the eight BEA regions, the model assumes that real income in each region is the sum of region-specific trend and cyclical components. The trend is assumed to follow a unit root with drift, whereas the cyclical component is made up of a common cycle across regions as well as a regional cycle. Each region has a different sensitivity with respect to the common cycle. The estimation of this UC model provides an estimated U.S. business cycle that accords well with National Bureau of Economic Research (NBER) recession periods and region-specific counterparts. Overall, the results suggest that a large share of the regional business cycle variations is explained by the common component and that spillovers from one region to another are not a significant source of variations. Even though Kouparitsas assumes a common cycle, he does not assume the existence of a common trend. This assumption makes his model less general than the model I propose, which also incorporates the unemployment rate in addition to data on real GDP to inform the estimation of the cycle and allows one to interpret the estimated cycle as a measure of the output gap in contrast with the cycle obtained from real income data.

Del Negro and Otrok (2008) extend a factor model to incorporate time-varying factor loadings and stochastic volatility to extract the international business cycle using a panel of 19 countries. The model is estimated with Bayesian methods and allows one to obtain the common and countryspecific cycles for the GDP growth rates, but it does not consider the common GDP trends.

Mitra and Sinclair (2012) also use many units and one variable. They propose a multivariate UC model to simultaneously decompose real GDP for each of the G-7 countries into its respective trend and cycle components. The setup considers real GDP as the only observable variable and assumes that each country's GDP is driven by specific trend and cycle components. The setup allows for possible correlations between any of the contemporaneous shocks to the unobserved (trend and cycle) components.

Stock and Watson (2016) propose a multivariate dynamic factor model with time-varying coefficients and stochastic volatility estimated with Bayesian methods to calculate the U.S. trend inflation. They use 17 components of the personal consumption expenditure price inflation to con-
struct an index akin to core inflation. This work is the most similar to the present paper, although Stock and Watson's paper only considers one variable in the analysis - the inflation rate - while this paper considers two - real GDP and the unemployment rate, which are linked in a structural way assuming Okun's law to provide more information for estimating the cyclical component of aggregate output. Moreover, the framework of Stock and Watson is not suitable to analyze business cycle fluctuations because the common and idiosyncratic shocks are all assumed to not have serial correlation.

## A. 3 Many Variables and Many Economic Units

Gregory, Head and Raynauld (1997) use a dynamic factor model estimated with classical methods to decompose aggregate output, consumption, and investment for the G-7 countries into factors that are (i) common across all countries and aggregates, (ii) common across aggregates within a country, and (iii) specific to each individual aggregate. The authors have to detrend the data to use the dynamic factor models approach because the underlying factors are assumed to be stationary. Similarly, Kose, Otrok and Whiteman (2003) estimate a dynamic factor model, but with Bayesian methods, to extract common components from macroeconomic aggregates (output, consumption, and investment) in a 60 -country sample covering seven regions of the world. They allow factors common to the world, the regions, and the countries. Here, too, data are de-trended.

For the U.S., Owyang, Rapach and Wall (2009) use state-level income and payroll employment data to estimate a dynamic factor model of the 48 contiguous states and the District of Columbia in order to extract business cycle factors. The estimation of the model identifies three common factors underlying the fluctuations in state-level income and employment growth, with the first common factor resembling aggregate fluctuations in real activity at the national level. The factors explain a large proportion of the total variability in state-level variables, although there is still a substantial amount of cross-state heterogeneity.

## B State-space Model in Matrix Form

The model in equation (1)-equation (11) can be written in matrix form as the following:

$$
\begin{align*}
& \mathbf{z}_{i t}=\mathrm{C}\left(\boldsymbol{\Theta}_{\mathrm{mi}}\right)+\mathrm{H}\left(\boldsymbol{\Theta}_{\mathrm{mi}}\right) \mathbf{x}_{i t}+\mathbf{w}_{i t}, \quad \mathbf{w}_{i t} \mid \mathfrak{F}_{t-1} \sim \operatorname{iid} \mathbb{N}\left(0, \mathrm{R}\left(\boldsymbol{\Theta}_{\mathrm{mi}}\right)\right)  \tag{B.1}\\
& \mathbf{x}_{i t}=\mathrm{F}\left(\boldsymbol{\Theta}_{\mathrm{si}}\right) \mathbf{x}_{i, t-1}+\mathbf{G v}_{i t}, \quad \mathbf{v}_{i t} \mid \mathfrak{F}_{t-1} \sim \operatorname{iid} \mathbb{N}\left(0, \mathrm{Q}\left(\boldsymbol{\Theta}_{\mathrm{si}}\right)\right), \tag{B.2}
\end{align*}
$$

where

$$
\begin{aligned}
& \mathbf{z}_{i t}=\left[\begin{array}{c}
\Delta y_{i t} \\
\Delta u_{i t}
\end{array}\right], \quad \mathbf{x}_{i t}=\left[\begin{array}{c}
c_{t} \\
c_{t-1} \\
c_{t-2} \\
\eta_{t}^{y} \\
\eta_{t}^{u} \\
v_{i t} \\
v_{i, t-1} \\
v_{i, t-2}
\end{array}\right], \quad \mathbf{w}_{i t}=\left[\begin{array}{c}
\eta_{i t}^{y} \\
\eta_{i t}^{u}
\end{array}\right], \quad \mathbf{v}_{i t}=\left[\begin{array}{c}
\varepsilon_{t} \\
\eta_{t}^{y} \\
\eta_{t}^{u} \\
\zeta_{i t}
\end{array}\right], \\
& \mathrm{C}\left(\boldsymbol{\Theta}_{\mathrm{mi}}\right)=\left[\begin{array}{c}
\mu_{i}+\delta_{i}^{y} \mu \\
0
\end{array}\right], \\
& \mathrm{H}\left(\boldsymbol{\Theta}_{\mathrm{mi}}\right)=\left[\begin{array}{cccccccc}
\alpha_{i} & -\alpha_{i} & 0 & \delta_{i}^{y} & 0 & 1 & -1 & 0 \\
\alpha_{i} \theta_{1 i} & \alpha_{i}\left(\theta_{2 i}-\theta_{1 i}\right) & -\alpha_{i} \theta_{2 i} & 0 & \delta_{i}^{u} & \theta_{1 i} & \theta_{2 i}-\theta_{1 i} & -\theta_{2 i}
\end{array}\right], \\
& \left.\mathrm{R}\left(\boldsymbol{\Theta}_{\mathrm{vi}}\right)\right)=\left[\begin{array}{cc}
\sigma_{\eta_{i}^{y}}^{2} & 0 \\
0 & \sigma_{\eta_{i}^{u}}^{2}
\end{array}\right],
\end{aligned}
$$

$$
\begin{aligned}
\mathrm{F}\left(\boldsymbol{\Theta}_{\mathrm{si}}\right) & =\left[\begin{array}{cccccccc}
\phi_{1} & \phi_{2} & 0 & 0 & 0 & 0 & 0 & 0 \\
1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & \rho_{i} & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 1 & 0
\end{array}\right], \\
\mathbf{G} & =\left[\begin{array}{llll}
1 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 \\
0 & 1 & 0 & 0 \\
0 & 0 & 1 & 0 \\
0 & 0 & 0 & 1 \\
0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0
\end{array}\right], \\
\mathrm{Q}\left(\Theta_{\mathrm{vi}}\right) & =\left[\begin{array}{lll}
\sigma_{\varepsilon}^{2} & 0 & 0 \\
\sigma_{\varepsilon}^{2} & 0 \\
0 & \sigma_{\eta^{y}}^{2} & 0 \\
0 & 0 & \sigma_{\eta^{u}}^{2} \\
0 & 0 & 0 \\
0 \\
\sigma_{\zeta_{i}}^{2}
\end{array}\right]
\end{aligned}
$$

and

$$
\begin{aligned}
\boldsymbol{\Theta}_{\mathrm{mi}} & =\left\{\mu, \mu_{i}, \alpha_{i}, \theta_{1 i}, \theta_{2 i}, \delta_{i}^{y}, \delta_{i}^{u}, \sigma_{\eta_{i}^{y}}^{2}, \sigma_{\eta_{i}^{u}}^{2}\right\} \\
\boldsymbol{\Theta}_{\mathrm{si}} & =\left\{\phi_{1}, \phi_{2}, \rho_{i}, \sigma_{e}^{2}, \sigma_{\eta^{y}}^{2}, \sigma_{\eta^{u}}^{2}, \sigma_{\zeta_{i}}^{2}\right\},
\end{aligned}
$$

for $i=1, \ldots, n$.

## C Details on the Gibbs Sampler

Let $\mathbf{z}_{i t}, \mathbf{x}_{i t}, \boldsymbol{\Theta}_{\mathrm{mi}}$, and $\boldsymbol{\Theta}_{\mathrm{si}}$ for $i=1,2, \ldots, n$, be defined as in Appendix B. Let $\mathbf{Z}_{T}=\left\{\tilde{\mathbf{z}}_{1}, \tilde{\mathbf{z}}_{2}, \ldots, \tilde{\mathbf{z}}_{T}\right\}$ denote the observed data and let $\mathbf{X}_{T}=\left\{\tilde{\mathbf{x}}_{1}, \tilde{\mathbf{x}}_{2}, \ldots, \tilde{\mathbf{x}}_{T}\right\}$. Here, $\tilde{\mathbf{z}}_{t}=\left\{\mathbf{z}_{1 t}, \mathbf{z}_{2 t}, \ldots, \mathbf{z}_{n t}\right\}$ and $\tilde{\mathbf{x}}_{t}=\left\{\mathbf{x}_{1 t}, \mathbf{x}_{2 t}, \ldots, \mathbf{x}_{n t}\right\}$. Denote $\boldsymbol{\Theta}_{\mathrm{m}}=\bigcup_{\mathrm{i}=1}^{n} \boldsymbol{\Theta}_{\mathrm{mi}}$ and $\boldsymbol{\Theta}_{\mathrm{s}}=\bigcup_{\mathrm{i}=1}^{n} \boldsymbol{\Theta}_{\mathrm{si}}$.

Partition $\boldsymbol{\Theta}_{\mathrm{si}}=\boldsymbol{\Theta}_{\mathrm{si}}^{1} \cup \Theta_{\mathrm{si}}^{2} \cup \Theta_{\mathrm{si}}^{3}$, where

$$
\begin{aligned}
& \boldsymbol{\Theta}_{\mathrm{si}}^{1}=\left\{\phi_{1}, \phi_{2}\right\}, \\
& \boldsymbol{\Theta}_{\mathrm{si}}^{2}=\left\{\rho_{i}, \sigma_{\zeta_{i}}^{2}\right\}, \\
& \boldsymbol{\Theta}_{\mathrm{si}}^{3}=\left\{\sigma_{e}^{2}, \sigma_{\eta^{y}}^{2}, \sigma_{\eta^{u}}^{2}\right\} .
\end{aligned}
$$

Notice that the identification conditions imply that $\Theta_{\mathrm{si}}^{3}$ is not random.
Also, partition $\Theta_{\mathrm{mi}}=\Theta_{\mathrm{mi}}^{1} \cup \Theta_{\mathrm{mi}}^{2}$, where

$$
\begin{aligned}
& \boldsymbol{\Theta}_{\mathrm{mi}}^{1}=\left\{\mu_{i}, \alpha_{i}, \theta_{1 i}, \theta_{2 i}, \delta_{i}^{y}, \delta_{i}^{u}\right\}, \\
& \boldsymbol{\Theta}_{\mathrm{mi}}^{2}=\left\{\sigma_{\eta_{i}^{y}}^{2}, \sigma_{\eta_{i}^{u}}^{2}\right\},
\end{aligned}
$$

where $\mu$ has been excluded because it is fixed under the identification conditions.
The Gibbs sampler procedure is as follows:

1. Start with initial values for the model's parameters, $\boldsymbol{\Theta}=\boldsymbol{\Theta}_{\mathrm{m}} \cup \boldsymbol{\Theta}_{\mathrm{s}}$.
2. Draw $\mathbf{X}_{T}$ from $p\left(\mathbf{X}_{T} \mid \mathbf{Z}_{T}, \boldsymbol{\Theta}_{\mathrm{m}}, \mathbf{\Theta}_{\mathrm{s}}\right)$ using the Durbin and Koopman (2002) simulation smoother.
3. Draw $\boldsymbol{\Theta}_{\mathrm{s}}^{1}$ from $p\left(\mathbf{\Theta}_{\mathrm{s}}^{1} \mid \mathbf{Z}_{T}, \mathbf{X}_{T}, \boldsymbol{\Theta}_{\mathrm{m}}, \boldsymbol{\Theta}_{\mathrm{si}}^{2}, \boldsymbol{\Theta}_{\mathrm{si}}^{3}\right)$ using the conditional distributions implied by the independent normal-inverse-gamma prior.
4. For $i=1,2, \ldots, n$, sample as follows:
(a) Draw $\boldsymbol{\Theta}_{\mathrm{si}}^{2}$ from $p\left(\boldsymbol{\Theta}_{\mathrm{si}}^{2} \mid \mathbf{Z}_{T}, \mathbf{X}_{T}, \boldsymbol{\Theta}_{\mathrm{m}}, \boldsymbol{\Theta}_{\mathrm{s}}^{1}, \mathbf{\Theta}_{\mathrm{si}}^{2}\right)$ using the conditional distributions implied by the independent normal-inverse-gamma prior.
(b) Draw $\boldsymbol{\Theta}_{\mathrm{mi}}^{1}$ from $p\left(\boldsymbol{\Theta}_{\mathrm{mi}}^{1} \mid \mathbf{Z}_{T}, \mathbf{X}_{T}, \boldsymbol{\Theta}_{\mathrm{mi}}^{2}, \boldsymbol{\Theta}_{\mathrm{mi}}^{3}, \boldsymbol{\Theta}_{\mathrm{m}(/ \mathrm{i})}, \boldsymbol{\Theta}_{\mathrm{s}}\right)$ using the conditional distributions implied by the independent normal-inverse-gamma prior. Repeat similarly for $\boldsymbol{\Theta}_{\mathrm{mi}}^{2}$ and sample from

$$
\text { - } p\left(\mathbf{\Theta}_{\mathrm{mi}}^{2} \mid \mathbf{Z}_{T}, \mathbf{X}_{T}, \mathbf{\Theta}_{\mathrm{mi}}^{1}, \mathbf{\Theta}_{\mathrm{mi}}^{3}, \boldsymbol{\Theta}_{\mathrm{m}(/ \mathrm{i} \mathbf{i}}, \boldsymbol{\Theta}_{\mathrm{s}}\right)
$$

5. Return to step 2.

## D States Most Strongly Affected by the Great Recession and Growth Rates by States after the Great Recession

The states that were most strongly affected by the Great Recession in terms of real GDP growth rates in the 2007-09 period appear in figure M.1. Arizona, Florida, Michigan, and Nevada experienced the lowest average annualized growth rates of real GDP, with rates lower than negative 3 percent. By contrast, states such as Alaska, South Dakota, and North Dakota experienced average growth rates greater than 3 percent in that period (not shown). All told, 26 of the 50 states plus the District of Columbia experienced negative average growth rates. Figure M. 2 shows the average output growth rates that the states experienced during and after the recovery between 2010 and 2016. North Dakota and Texas have grown the fastest, with average rates above 3 percent, while Wyoming and Alaska have grown the slowest, with average rates close to negative 1 percent.
[Figure 1 about here.]

Several different factors can help to rationalize the heterogeneity in the growth rates of the U.S. states during and after recessions. Industry composition, demographics, credit demand and supply, and state fiscal policies, among others have been mentioned in the literature (see Owyang, Piger and Wall, 2005; Carlino and Defina, 1998; Mian and Sufi, 2009, 2010; Owyang and Zubairy, 2013, for example). I exploit the heterogeneity in the data at the state level to obtain a common estimate of the cycle - the aggregate output gap - while at the same time I relate the aforementioned factors to characterize the output gaps at the state level in the forthcoming sections.

## E Prior Distributions

The prior distributions of the parameters of the UC model with state-level data appear in table M.1. The equations for each parameter are in the last column.
[Table 1 about here.]

The prior means of the parameters of the common cycle, $\phi_{1}$ and $\phi_{2}$, are similar to those found in the literature on trend-cycle decompositions of output. For example, Morley, Nelson and Zivot (2003) find that the estimation of a UC model with aggregate data on real GDP in absence of correlation between trend and cycle innovations yields estimated coefficients $\hat{\phi}_{1}=1.53$ and $\hat{\phi}_{2}=-0.61$. Similarly, Gonzalez-Astudillo and Roberts (2016) also use aggregate data but include the unemployment rate along with real GDP and find estimates around 1.6 for $\phi_{1}$ and -0.65 for $\phi_{2}$ both with or without correlation between output innovations. In the present paper, the joint prior distribution of these two parameters is truncated to satisfy the weak stationarity feature of the common cycle. The mean growth rate of each state's GDP has a truncated normal prior distribution with mean 0.8 , which yields an annualized growth rate of real GDP around 3 percent, roughly the historical average. The distributions of the parameters that load on the common cycle, $\alpha_{i}$, are normal with means and standard deviations equal to one. ${ }^{1}$ The coefficients that load on the common trends, $\delta_{i}^{y}$, and $\delta_{i}^{u}$, have truncated normal distributions with means and standard deviations equal to one. The Okun's law coefficients, $\theta_{1 i}$ and $\theta_{2 i}$, are normally distributed with prior means equal to -0.25 each, such that the long-run Okun's law coefficient for each state is centered at -0.5 under the prior distribution. This is the usual Okun's law coefficient used in the literature (see Abel, Bernanke and Croushore, 2013). The prior distribution of the parameter of the idiosyncratic cycle, $\rho_{i}$, is a truncated standard normal. Finally, given the lack of previous estimates of these coefficients in the literature, I assume that $\sigma_{\eta_{i}^{y}}^{2}, \sigma_{\eta_{i}^{u}}^{2}$, and $\sigma_{\zeta_{i}}^{2}$ are distributed as inverse-gamma centered at one with undefined variance. In general, the standard deviations of the parameters imply that the distributions are neither too tight nor too narrow. ${ }^{2}$

## F Parameter Results and Convergence Diagnostics

This appendix lays out the results from the Bayesian estimation showing the estimates of the posterior mean and standard deviation of the parameters of the model, as well as the first and fiftieth order autocorrelation coefficient, the relative numerical efficiency (RNE) using a 4 percent taper, and the p-value of the Geweke (1991) convergence diagnostics using a 4 percent tapper as

[^1]well in which the null hypothesis considers equality of the means of the first 20 percent of draws with that of the last 50 percent. Given the large number of parameters, the results appear in tables M.2-M.11. In the tables, the parameters $\mu_{i}, \alpha_{i}, \delta_{i}^{y}, \delta_{i}^{u}, \theta_{1 i}, \theta_{2 i}, \rho_{i}, \sigma_{\zeta_{i}}^{2}, \sigma_{\eta_{i}^{y}}^{2}$, and $\sigma_{\eta_{i}^{u}}^{2}$ are numbered from $i=1, \ldots, 51$ according to the states and the District of Columbia in the following order: Alabama, Alaska, Arizona, Arkansas, California, Colorado, Connecticut, Delaware, District of Columbia, Florida, Georgia, Hawaii, Idaho, Illinois, Indiana, Iowa, Kansas, Kentucky, Louisiana, Maine, Maryland, Massachusetts, Michigan, Minnesota, Mississippi, Missouri, Montana, Nebraska, Nevada, New Hampshire, New Jersey, New Mexico, New York, North Carolina, North Dakota, Ohio, Oklahoma, Oregon, Pennsylvania, Rhode Island, South Carolina, South Dakota, Tennessee, Texas, Utah, Vermont, Virginia, Washington, West Virginia, Wisconsin, and Wyoming. $\phi_{1}$ and $\phi_{2}$ are the parameters of the $\mathrm{AR}(2)$ specification of the common component of the cycle.

The draws from the posterior distribution used to produce the results of the model are based on 300,000 draws after burning in the first 100,000 and thinning every 100th draw, which left me with 2,000 draws from the posterior distribution. The results of the diagnostics tests show that these 2,000 draws do not evidence significant autocorrelation of first order and almost no autocorrelation of order fifty. Apart from 3 out of 513 parameters, the p-values of the test of equality of means between the fist 20 percent of the draws and the last 50 percent are all above 1 percent, which indicate that the null hypothesis is not rejected for any of the parameters and the sampler has converged.
[Table 2 about here.]
[Table 3 about here.]
[Table 4 about here.]
[Table 5 about here.]
[Table 6 about here.]
[Table 7 about here.]
[Table 8 about here.]
[Table 9 about here.]
[Table 10 about here.]
[Table 11 about here.]

## G Obtaining the Aggregate Trend and Cycle

The objective of the estimation is to obtain the trend-cycle decomposition of aggregate GDP by exploiting the cross-sectional variability of state-level data. Because of the nonlinearity implicit in the aggregation of the variables and the fact that state-level GDP appears in logs in the specification equation (1)-equation (11), an approximation is needed. The quarterly growth rate of aggregate real GDP is given by the following:

$$
\begin{aligned}
\Delta \% Y_{t} & =\sum_{i=1}^{n} w_{i t} \Delta \% Y_{i t} \\
& \approx \sum_{i=1}^{n} w_{i t} \Delta y_{i t} \\
& =\sum_{i=1}^{n} w_{i t}\left(\Delta \tau_{i t}^{y}+\Delta c_{i t}\right) \\
& =\sum_{i=1}^{n} w_{i t}\left(\delta_{i}^{y}\left(\mu+\eta_{t}^{y}\right)+\mu_{i}+\eta_{i t}^{y}+\alpha_{i} \Delta c_{t}+\Delta v_{i t}\right) \\
& \approx \underbrace{\overline{\delta^{y}}\left(\mu+\eta_{t}^{y}\right)+\bar{\mu}+\bar{\eta}_{t}^{y}}_{\Delta \text { GDP Trend }}+\underbrace{\bar{\alpha} \Delta c_{t}+\bar{v}_{t}^{y}}_{\Delta \text { GDP Cycle }},
\end{aligned}
$$

where $Y_{i t}$ is real GDP of state $i, Y_{t}$ is aggregate real GDP in period $t$, and the contribution of of state's $i$ GDP to aggregate GDP is denoted by $w_{i t}$.

Hence, I can express the trend and cycle components of the aggregate GDP as

$$
\begin{aligned}
& \text { GDP Trend } \approx \overline{\delta^{y}} \tau_{t}^{y}+\bar{\xi}_{t}^{y}, \\
& \text { GDP Cycle } \approx \bar{\alpha} c_{t}+\bar{v}_{t}^{y},
\end{aligned}
$$

where

$$
\begin{aligned}
\bar{\mu}_{t} & =\sum_{i=1}^{n} w_{i} \mu_{i}, \\
\bar{\delta}^{y} & =\sum_{i=1}^{n} w_{i} \delta_{i}^{y}, \\
\bar{\eta}_{t}^{y} & =\sum_{i=1}^{n} w_{i} \eta_{i t}^{y}, \\
\bar{\xi}_{t}^{y} & =\sum_{i=1}^{n} w_{i} \xi_{i t}^{y}, \\
\bar{\alpha} & =\sum_{i=1}^{n} w_{i} \alpha_{i}, \\
\bar{v}_{t}^{y} & =\sum_{i=1}^{n} w_{i} v_{i t},
\end{aligned}
$$

where $w_{i}=\sum_{t=1}^{T} w_{i t} / T$ is the sample average of each state's contributions to aggregate GDP. I use the smoothed estimates $c_{t \mid T}$ and $v_{i t \mid T}$ to obtain the estimate of the cycle and, by residual, the trend.

## H Variance Decomposition of States' Cycles and Trends

In this section, I describe how each state's output cycle and trend variability are explained by the variability in the common cycle and trend during the period of analysis. To that end, I perform variance decompositions of the state's cycle and trend of output. Recall that the specification for the cycle of each state $i=1,2, \ldots, n$ is $c_{i t}=\alpha_{i} c_{t}+v_{i t}$. Hence, one can obtain the fraction of the variance of the cycle that is due to the common cycle and the fraction that corresponds to the idiosyncratic component. Similarly, because the output trend of each state is given by $\tau_{i t}^{y}=\delta_{i}^{y} \tau_{t}^{y}+\xi_{i t}^{y}$, one can obtain the proportion of the variability of the output trend of each state that is due to the common trend and that due to the idiosyncratic component. Figures M. 3 and M. 4 show the percent of the variance decomposition of the state's cycles and output trends, respectively.
[Figure 2 about here.]

Economic policies designed at the federal level can have different effects on the state economies, depending, among several factors, on industry composition, size of firms, the ability of banks to alter their balance sheets, demographics, and government spending composition as described by Carlino and DeFina (1999) and Owyang and Zubairy (2013). In that regard, the propagation of federal economic policies at the state level can be different depending, for instance, on how strongly a particular state's cycle or trend is linked to its common counterpart, which I assumed is the target of economic policies at the federal level. On the one hand, Nevada, California, Georgia, Arizona, and Florida are among the states with cycle variability that is explained the most by the variance of the common cycle, whereas West Virginia, Wyoming, Louisiana, and North Dakota have the smallest variation of their cycle attributed to the common cycle. On the other hand, the states whose trends are most strongly connected with the common trend are Kentucky, Alabama, and Indiana, whereas those with lowest associations are New York, Delaware, and Alaska.

## I Variance Decomposition of States' GDP Growth

The model also allows one to establish a measure of how cyclical the state economies are. In this section, I decompose the variance of real GDP growth at the state level in the proportion that is explained by variations of the cycle and the proportion that is explained by variations of the output trend during the period of analysis. This indicator is useful to characterize the sources of variations in GDP growth for policymaking decisions both at the state and the federal level, for example. Figure M. 5 illustrates the proportions for each state.
[Figure 3 about here.]

California, Utah, Florida, Arizona, Georgia, and Oregon have the highest proportion of their real GDP growth's variance explained by the variance of their respective output cycles. In contrast, states such as North Dakota, Wyoming, Vermont, and West Virginia have the lowest proportions. One can also characterize the features that make a state "more cyclical" than others. Table M. 13 shows that, on average, and at the 10 percent level of significance, states with higher participation of the sectors related with wholesale trade, retail trade, and transportation are more cyclical, likewise states that have a higher share of population between the ages of 18 and 44. In addition, states with
higher leverage tend to experience fluctuations in output that are explained more by fluctuations in their cyclical component than in their trend output.

## J Okun's Law Coefficients at the State Level

As a final point, the results shed light on the cyclical features of the state-level labor markets, in particular the sensitivity of the unemployment gap to the output gap in each state. Policies designed to affect the cyclical position of the economy at the federal level can propagate differently across states' labor markets via two channels. First, these policies can affect the state output gaps differently. Second, the state labor markets can have unique reactions to their respective output gaps. I describe the first channel in online Appendix H. I measure the second channel by the sum of the Okun's law coefficients, $\theta_{1 i}+\theta_{1 i}$. The higher the absolute value of this sum, the more responsive the unemployment rate is to the cyclical fluctuations of output at the state level. Figure M. 6 presents the results grouped by values of the posterior mean estimates.

## [Figure 4 about here.]

According to the estimates of the model, the states with more-cyclical labor markets are Louisiana, Rhode Island, and Mississippi whereas the states with less-cyclical labor markets are Nebraska, Kansas, and Alaska. Guisinger et al. (2018) find that Lousiana has the highest comparable Okun's law coefficient, just as in this paper. Rhode Island also appears in the top three states classified by their Okun's law coefficients. On the other extreme, Guisinger et al. find South Dakota and North Dakota, among other states, with very low coefficients, similar to the findings in this paper.

The average of the posterior mean of the sum of the Okun's law coefficients across states is about -0.3 , which is smaller in absolute value than the usual coefficient of -0.5 . Likewise, Owyang and Sekhposyan (2012) and Grant (2018) report a statistically significant decline in the reaction of the unemployment gap to the output gap during the Great Recession. Similarly, Ball, Leigh and Loungani (2017) also find that the Okun's law coefficient has declined recently, but the difference with respect to the past does not appear to be statistically significant.

Table M. 14 shows that those states with higher participation of the construction industry tend to be characterized by more-cyclical labor markets in this period, whereas those states that have higher contributions from agriculture and mining tend to have less-cyclical labor markets, as is expected from these sectors. Guisinger et al. (2018) emphasize that institutional differences can help explain the variation in the size of Okun's law coefficients across states. One of those institutional factors is the percentage of the workforce who are union members. I add said variable to the set of regressors previously used, but the results show that this variable is not statistically significant at conventional levels.

As a final exercise with respect to labor market cyclicality, the model allows one to estimate the effect of an increase in the aggregate output gap on each state's unemployment rate. This exercise is particularly useful to understand how the unemployment rate reacts at the state level with respect to policies designed at the aggregate level. Figure M. 7 shows the response of the statelevel unemployment rate to such a shock over 20 quarters for the three most sensitive and the three least sensitive states. ${ }^{3}$ On the one hand, a 0.8 percentage point increase in the aggregate output gap causes the unemployment rates of Nevada, Arizona, and California to decline by about 1, 0.8 , and 0.75 percentage point, respectively, within 3 quarters of the shock. On the other hand, the responses of Wyoming, North Dakota, and West Virginia imply a muted effect on the unemployment rates of those states.

## K Variance Decomposition of Aggregate GDP Growth

As described in Section 2, the common trend and cycle components imply co-movement among the state economies that, in turn, influence the variability of the aggregate GDP. One can compute the variance decomposition of aggregate GDP growth, which provides an indicator of how cyclical is the U.S. economy, as follows: ${ }^{4}$

```
var( }\Delta%\mathrm{ Aggregate GDP ) = var( }\Delta\mathrm{ Aggregate GDP trend ) + var( }\Delta\mathrm{ Output Gap).
```

[^2]Therefore, the percent of the variance of GDP growth that is due to the variance of the change in the aggregate cycle is given by the following expression:

$$
\frac{\operatorname{var}(\Delta \text { Output Gap })}{\operatorname{var}(\Delta \text { Aggregate GDP trend })+\operatorname{var}(\Delta \text { Output Gap })}=62 \% .
$$

The contribution of the variability of the cycle to the variability of GDP growth of 62 percent is close to estimates obtained with longer samples. For example, Gonzalez-Astudillo and Roberts (2016) find that the contribution is around 60 percent or 65 percent, depending on the assumption about the correlation between trend and cycle components.

## L Derivation of the Variance Decomposition of Aggregate GDP

Given that (the log of) aggregate GDP can be specified as the sum of its aggregate trend and cycle, one can write the following:

$$
\begin{aligned}
\operatorname{var}\left(\Delta \% Y_{t}\right) & =\operatorname{var}(\Delta \text { Aggregate GDP trend })+\operatorname{var}(\Delta \text { Output Gap }) \\
& \approx \operatorname{var}\left(\overline{\delta^{y}} \Delta \tau_{t}+\Delta \bar{\xi}_{t}^{y}\right)+\operatorname{var}\left(\bar{\alpha} \Delta c_{t}+\Delta \bar{v}_{t}^{y}\right) \\
& =\bar{\delta}^{2} \operatorname{var}\left(\eta_{t}^{y}+\bar{\eta}^{y} t\right)+\bar{\alpha}^{2} \operatorname{var}\left(\Delta c_{t}+\Delta \bar{v}_{t}^{y}\right) \\
& \approx \bar{\delta}^{2}\left(1+\sum_{i=1}^{n} w_{i}^{2} \sigma_{\eta_{i}^{y}}^{2}\right)+\bar{\alpha}^{2}\left(2\left(\gamma_{0}-\gamma_{1}\right)+\sum_{i=1}^{n} w_{i}^{2} \sigma_{\Delta v_{i t}}^{2}\right),
\end{aligned}
$$

where

$$
\begin{aligned}
\gamma_{0} & =\frac{\frac{1-\phi_{2}}{1+\phi_{2}}}{\left(1-\phi_{2}^{2}\right)-\phi_{1}^{2}}, \\
\gamma_{1} & =\frac{\phi_{1}}{1-\phi_{2}} \gamma_{0} .
\end{aligned}
$$

## M Prior and Posterior Distribution Results from the Aggregate UC Model

The prior distribution choices and the posterior mean and standard deviation results appear in table M.12.
[Table 12 about here.]

The mean annual growth rate of potential GDP $(4 \mu)$ is estimated to be close to 1.6 percent, 1 percentage point lower than the estimate with state-level data obtained from weight averaging the estimated state-level mean growth rates $\left(4 \sum_{i=1}^{51} w_{i} \mu_{i}\right)$. With respect to the Okun's law coefficients, the long-run sensitivity of the unemployment gap to the output gap is about -0.59 , more than twice the simple average obtained from the state-level data. Owyang, Vermann and Sekhposyan (2013) also investigate the Okun's law coefficients across states, although they use the growth rates version of the law, and find that there can be discrepancies between the estimates with state-level data and the estimate obtained from aggregate data.

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Table M.1: Prior Distributions of the Parameters using State-Level Data

| Parameter | Distribution | Mean | Standard Deviation | Equation |
| ---: | :---: | :---: | :---: | :--- |
| $\phi_{1}$ | Truncated Normal | 1.5 | 1 | $c_{t}=\phi_{1} c_{t-1}+\phi_{2} c_{t-2}+\varepsilon_{t}$ |
| $\phi_{2}$ | Truncated Normal | -0.6 | 1 | $c_{t}=\phi_{1} c_{t-1}+\phi_{2} c_{t-2}+\varepsilon_{t}$ |
| $\mu_{i}$ | Truncated Normal | 0.8 | 0.8 | $\xi_{i t}^{y}=\mu_{i}+\xi_{i, t-1}^{y}+\eta_{i t}^{y}$ |
| $\alpha_{i}$ | Normal | 1 | 1 | $c_{i t}=\alpha_{i} c_{t}+v_{i t}$ |
| $\delta_{i}^{y}$ | Truncated Normal | 1 | 1 | $\tau_{i t}^{y}=\delta_{i}^{y} \tau_{t}^{y}+\xi_{i t}^{y}$ |
| $\delta_{i}^{u}$ | Truncated Normal | 1 | 1 | $\tau_{i t}^{u}=\delta_{i}^{u} \tau_{t}^{u}+\xi_{i t}^{u}$ |
| $\theta_{1 i}$ | Normal | -0.25 | 0.25 | $u_{i t}=\tau_{i t}^{u}+\theta_{1 i} c_{i t}+\theta_{2 i} c_{i, t-1}$ |
| $\theta_{2 i}$ | Normal | -0.25 | 0.25 | $u_{i t}=\tau_{i t}^{u}+\theta_{1 i} c_{i t}+\theta_{2 i} c_{i, t-1}$ |
| $\rho_{i}$ | Truncated Normal | 0 | 1 | $v_{i t}=\rho_{i} v_{i, t-1}+\zeta_{i t}$ |
| $\sigma_{\eta_{i}^{y}}^{2}$ | Inverse Gamma | 1 | $\xi_{i t}=\mu_{i}+\xi_{i, t-1}^{y}+\eta_{i t}^{y}, \quad \operatorname{var}\left(\eta_{i t}^{y}\right)=\sigma_{\eta_{i}^{y}}^{2}$ |  |
| $\sigma_{\eta_{i}^{u}}^{2}$ | Inverse Gamma | 1 | $\operatorname{Inf}$ | $\xi_{i t}^{u}=\xi_{i, t-1}^{u}+\eta_{i t}^{u}, \quad \operatorname{var}\left(\eta_{i t}^{u}\right)=\sigma_{\eta_{i}^{u}}^{2}$ |
| $\sigma_{\zeta_{i}}^{2}$ | Inverse Gamma | 1 | Inf | $v_{i t}=\rho_{i} v_{i, t-1}+\zeta_{i t}, \quad \operatorname{var}\left(\zeta_{i t}\right)=\sigma_{\zeta_{i}}^{2}$ |

Table M.2: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\phi_{1}$ | 1.518 | 0.159 | 0.026 | 0.014 | 0.840 | 0.923 |
| $\phi_{2}$ | -0.555 | 0.157 | 0.025 | 0.014 | 0.813 | 0.783 |
| $c_{T}$ | 0.920 | 10.458 | 0.005 | 0.004 | 0.868 | 0.587 |
| $\mu_{1}$ | 0.553 | 0.135 | 0.182 | 0.006 | 0.401 | 0.946 |
| $\mu_{2}$ | 0.524 | 0.272 | 0.165 | -0.025 | 0.309 | 0.548 |
| $\mu_{3}$ | 0.676 | 0.148 | 0.101 | 0.015 | 0.597 | 0.973 |
| $\mu_{4}$ | 0.676 | 0.190 | 0.075 | 0.011 | 0.680 | 0.567 |
| $\mu_{5}$ | 0.744 | 0.123 | 0.049 | -0.015 | 0.542 | 0.830 |
| $\mu_{6}$ | 0.757 | 0.119 | 0.077 | 0.022 | 0.833 | 0.787 |
| $\mu_{7}$ | 0.271 | 0.157 | 0.010 | -0.023 | 0.898 | 0.235 |
| $\mu_{8}$ | 0.472 | 0.275 | -0.010 | -0.012 | 1.043 | 0.844 |
| $\mu_{9}$ | 0.564 | 0.116 | 0.019 | -0.003 | 0.718 | 0.058 |
| $\mu_{10}$ | 0.570 | 0.133 | 0.153 | 0.018 | 0.497 | 0.101 |
| $\mu_{11}$ | 0.570 | 0.121 | 0.130 | -0.009 | 0.353 | 0.744 |
| $\mu_{12}$ | 0.606 | 0.115 | 0.159 | -0.000 | 0.470 | 0.688 |
| $\mu_{13}$ | 0.749 | 0.149 | 0.048 | 0.034 | 0.777 | 0.245 |
| $\mu_{14}$ | 0.485 | 0.114 | 0.122 | -0.013 | 0.612 | 0.615 |
| $\mu_{15}$ | 0.827 | 0.193 | 0.253 | 0.019 | 0.202 | 0.425 |
| $\mu_{16}$ | 0.699 | 0.181 | 0.082 | 0.014 | 0.495 | 0.867 |
| $\mu_{17}$ | 0.843 | 0.223 | 0.049 | -0.018 | 0.911 | 0.057 |
| $\mu_{18}$ | 0.749 | 0.176 | 0.196 | 0.004 | 0.303 | 0.370 |
| $\mu_{19}$ | 0.235 | 0.148 | 0.025 | -0.016 | 1.441 | 0.187 |
| $\mu_{20}$ | 0.464 | 0.139 | 0.164 | -0.015 | 0.343 | 0.935 |
| $\mu_{21}$ | 0.639 | 0.114 | 0.163 | 0.034 | 0.331 | 0.886 |
| $\mu_{22}$ | 0.799 | 0.124 | 0.124 | -0.012 | 0.502 | 0.528 |
| $\mu_{23}$ | 0.521 | 0.197 | 0.188 | -0.006 | 0.320 | 0.444 |
| $\mu_{24}$ | 0.634 | 0.153 | 0.173 | -0.005 | 0.228 | 0.532 |
| $\mu_{25}$ | 0.564 | 0.185 | 0.112 | -0.013 | 0.457 | 0.616 |
| $\mu_{26}$ | 0.453 | 0.132 | 0.054 | -0.014 | 0.886 | 0.342 |
| $\mu_{27}$ | 0.549 | 0.130 | 0.031 | -0.003 | 0.813 | 0.431 |
| $\mu_{28}$ | 0.762 | 0.162 | 0.124 | 0.001 | 0.340 | 0.206 |
| $\mu_{29}$ | 0.541 | 0.166 | 0.106 | -0.005 | 0.808 | 0.078 |
| $\mu_{30}$ | 0.584 | 0.145 | 0.018 | 0.025 | 0.661 | 0.660 |
| $\mu_{31}$ | 0.524 | 0.133 | 0.132 | 0.014 | 0.507 | 0.780 |
| $\mu_{32}$ | 0.372 | 0.179 | 0.152 | -0.011 | 0.259 | 0.293 |
| $\mu_{33}$ | 0.507 | 0.188 | -0.001 | -0.027 | 0.837 | 0.123 |
| $\mu_{34}$ | 0.675 | 0.160 | 0.071 | 0.020 | 0.406 | 0.309 |
| $\mu_{35}$ | 1.309 | 0.307 | 0.008 | 0.026 | 0.881 | 0.283 |
| $\mu_{36}$ | 0.542 | 0.151 | 0.184 | 0.023 | 0.279 | 0.746 |
| $\mu_{37}$ | 0.839 | 0.276 | 0.057 | -0.023 | 0.338 | 0.778 |
| $\mu_{38}$ | 0.820 | 0.152 | 0.029 | -0.010 | 0.958 | 0.702 |
| $\mu_{39}$ | 0.741 | 0.137 | 0.208 | 0.033 | 0.226 | 0.838 |
| $\mu_{40}$ | 0.383 | 0.161 | 0.071 | 0.022 | 0.683 | 0.072 |
| $\mu_{41}$ | 0.720 | 0.124 | 0.128 | 0.011 | 0.358 | 0.743 |
| $\mu_{42}$ | 0.689 | 0.237 | 0.016 | -0.014 | 0.756 | 0.631 |
| $\mu_{43}$ | 0.682 | 0.129 | 0.155 | 0.017 | 0.546 | 0.231 |
| $\mu_{44}$ | 0.853 | 0.136 | 0.088 | 0.017 | 0.449 | 0.505 |
| $\mu_{45}$ | 0.877 | 0.126 | 0.045 | 0.018 | 1.111 | 0.072 |
| $\mu_{46}$ | 0.594 | 0.151 | 0.100 | -0.013 | 0.575 | 0.691 |
| $\mu_{47}$ | 0.527 | 0.104 | 0.117 | 0.055 | 0.545 | 0.326 |
| $\mu_{48}$ | 1.053 | 0.140 | 0.084 | 0.003 | 0.482 | 0.752 |
| $\mu_{49}$ | 0.521 | 0.159 | 0.110 | -0.010 | 0.874 | 0.850 |
| $\mu_{50}$ | 0.553 | 0.118 | 0.145 | 0.004 | 0.375 | 0.968 |
| $\mu_{51}$ | 0.614 | 0.306 | 0.114 | -0.001 | 0.256 | 0.765 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; p-value: p -value of the null hypothesis that the draws; RNE: relative numerical efficiency using a 4 percent taper; p -value: p -value of the null hypothesis that the
mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper mean of the first 20 per for standard error.

Table M.3: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\alpha_{1}$ | 0.712 | 0.252 | 0.231 | -0.009 | 0.195 | 0.476 |
| $\alpha_{2}$ | -0.420 | 0.570 | 0.085 | -0.018 | 0.407 | 0.811 |
| $\alpha_{3}$ | 1.728 | 0.303 | 0.035 | 0.016 | 1.015 | 0.950 |
| $\alpha_{4}$ | 0.766 | 0.338 | 0.161 | 0.012 | 0.231 | 0.637 |
| $\alpha_{5}$ | 1.601 | 0.313 | 0.103 | -0.013 | 0.277 | 0.811 |
| $\alpha_{6}$ | 1.036 | 0.254 | 0.045 | 0.035 | 0.540 | 0.424 |
| $\alpha_{7}$ | 1.005 | 0.318 | 0.027 | 0.012 | 0.783 | 0.784 |
| $\alpha_{8}$ | 1.073 | 0.432 | -0.012 | 0.028 | 0.795 | 0.904 |
| $\alpha_{9}$ | 0.235 | 0.210 | 0.107 | -0.027 | 0.393 | 0.212 |
| $\alpha_{10}$ | 1.642 | 0.285 | 0.038 | 0.012 | 0.517 | 0.703 |
| $\alpha_{11}$ | 1.316 | 0.235 | 0.021 | 0.009 | 0.792 | 0.894 |
| $\alpha_{12}$ | 0.813 | 0.200 | 0.079 | -0.027 | 0.477 | 0.379 |
| $\alpha_{13}$ | 1.160 | 0.276 | 0.020 | 0.002 | 0.964 | 0.025 |
| $\alpha_{14}$ | 0.891 | 0.215 | 0.075 | 0.002 | 0.635 | 0.389 |
| $\alpha_{15}$ | 1.162 | 0.349 | 0.243 | -0.016 | 0.180 | 0.865 |
| $\alpha_{16}$ | 0.855 | 0.336 | 0.089 | -0.005 | 0.398 | 0.221 |
| $\alpha_{17}$ | 0.957 | 0.405 | 0.063 | 0.008 | 0.471 | 0.735 |
| $\alpha_{18}$ | 0.797 | 0.327 | 0.314 | -0.026 | 0.143 | 0.903 |
| $\alpha_{19}$ | 0.060 | 0.326 | 0.028 | -0.002 | 0.571 | 0.851 |
| $\alpha_{20}$ | 0.517 | 0.262 | 0.224 | -0.030 | 0.228 | 0.829 |
| $\alpha_{21}$ | 0.503 | 0.207 | 0.104 | -0.007 | 0.257 | 0.677 |
| $\alpha_{22}$ | 0.666 | 0.212 | 0.137 | 0.025 | 0.364 | 0.987 |
| $\alpha_{23}$ | 1.074 | 0.456 | 0.152 | -0.023 | 0.238 | 0.900 |
| $\alpha_{24}$ | 0.814 | 0.295 | 0.107 | -0.010 | 0.344 | 0.954 |
| $\alpha_{25}$ | 0.547 | 0.304 | 0.113 | -0.028 | 0.251 | 0.695 |
| $\alpha_{26}$ | 0.219 | 0.247 | 0.214 | -0.026 | 0.166 | 0.969 |
| $\alpha_{27}$ | 0.722 | 0.247 | -0.004 | 0.009 | 0.894 | 0.443 |
| $\alpha_{28}$ | 0.478 | 0.298 | 0.063 | -0.013 | 0.442 | 0.718 |
| $\alpha_{29}$ | 1.836 | 0.331 | 0.044 | 0.011 | 0.756 | 0.655 |
| $\alpha_{30}$ | 0.674 | 0.271 | 0.099 | -0.033 | 0.393 | 0.350 |
| $\alpha_{31}$ | 0.834 | 0.231 | 0.118 | 0.004 | 0.362 | 0.641 |
| $\alpha_{32}$ | 0.314 | 0.298 | 0.024 | 0.044 | 0.790 | 0.700 |
| $\alpha_{33}$ | 0.631 | 0.332 | -0.020 | -0.039 | 1.496 | 0.711 |
| $\alpha_{34}$ | 1.090 | 0.295 | 0.020 | -0.021 | 0.556 | 0.084 |
| $\alpha_{35}$ | -0.083 | 0.469 | 0.084 | -0.011 | 0.318 | 0.830 |
| $\alpha_{36}$ | 0.917 | 0.277 | 0.129 | 0.019 | 0.327 | 0.055 |
| $\alpha_{37}$ | 0.474 | 0.483 | 0.006 | 0.025 | 1.071 | 0.848 |
| $\alpha_{38}$ | 1.307 | 0.292 | 0.010 | 0.022 | 0.854 | 0.742 |
| $\alpha_{39}$ | 0.667 | 0.244 | 0.191 | -0.014 | 0.211 | 0.790 |
| $\alpha_{40}$ | 0.570 | 0.292 | 0.046 | -0.039 | 0.682 | 0.865 |
| $\alpha_{41}$ | 1.246 | 0.238 | 0.056 | -0.028 | 0.807 | 0.822 |
| $\alpha_{42}$ | 0.318 | 0.438 | 0.064 | 0.004 | 0.541 | 0.173 |
| $\alpha_{43}$ | 0.958 | 0.234 | 0.026 | 0.014 | 0.694 | 0.213 |
| $\alpha_{44}$ | 0.737 | 0.300 | 0.057 | 0.006 | 0.409 | 0.516 |
| $\alpha_{45}$ | 1.301 | 0.311 | 0.058 | 0.073 | 0.502 | 0.960 |
| $\alpha_{46}$ | 0.156 | 0.312 | 0.313 | -0.001 | 0.128 | 0.637 |
| $\alpha_{47}$ | 0.454 | 0.191 | 0.113 | 0.005 | 0.266 | 0.124 |
| $\alpha_{48}$ | 1.185 | 0.263 | -0.010 | 0.012 | 1.784 | 0.823 |
| $\alpha_{49}$ | -0.022 | 0.295 | 0.203 | 0.011 | 0.221 | 0.870 |
| $\alpha_{50}$ | 0.813 | 0.219 | 0.094 | -0.054 | 0.399 | 0.304 |
| $\alpha_{51}$ | 0.056 | 0.501 | 0.056 | 0.025 | 1.122 | 0.344 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; p -value: p -value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper for the standard error.

Table M.4: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\delta_{1}^{y}$ | 1.391 | 0.242 | 0.252 | -0.053 | 0.190 | 0.868 |
| $\delta_{2}^{y}$ | 0.563 | 0.558 | 0.527 | -0.031 | 0.078 | 0.854 |
| $\delta_{3}^{y}$ | 1.254 | 0.279 | 0.050 | 0.012 | 0.520 | 0.604 |
| $\delta_{4}^{y}$ | 1.637 | 0.347 | 0.012 | 0.030 | 0.574 | 0.960 |
| $\delta_{5}^{4}$ | 0.637 | 0.264 | 0.030 | -0.017 | 1.312 | 0.533 |
| $\delta_{6}^{y}$ | 0.792 | 0.240 | 0.046 | 0.006 | 0.773 | 0.990 |
| $\delta_{7}^{y}$ | 0.841 | 0.330 | 0.015 | -0.032 | 0.906 | 0.632 |
| $\delta_{8}^{y}$ | 0.647 | 0.434 | -0.004 | 0.034 | 0.951 | 0.013 |
| $\delta_{9}^{y}$ | 0.529 | 0.238 | 0.052 | -0.015 | 0.454 | 0.845 |
| $\delta_{10}^{y}$ | 1.134 | 0.249 | 0.095 | 0.040 | 0.501 | 0.405 |
| $\delta_{11}^{y}$ | 0.944 | 0.242 | 0.161 | -0.031 | 0.238 | 0.881 |
| $\delta_{12}^{y}$ | 1.056 | 0.210 | 0.147 | -0.057 | 0.289 | 0.530 |
| $\delta_{13}^{y}$ | 1.081 | 0.282 | -0.009 | -0.002 | 1.377 | 0.994 |
| $\delta_{14}^{y}$ | 1.038 | 0.214 | 0.100 | -0.030 | 0.453 | 0.871 |
| $\delta_{15}^{y}$ | 1.992 | 0.366 | 0.363 | -0.013 | 0.113 | 0.730 |
| $\delta_{16}^{y}$ | 1.367 | 0.341 | 0.144 | -0.042 | 0.284 | 0.539 |
| $\delta_{17}^{y}$ | 1.600 | 0.418 | 0.022 | 0.017 | 0.847 | 0.064 |
| $\delta_{18}^{y}$ | 1.932 | 0.306 | 0.256 | -0.044 | 0.167 | 0.265 |
| $\delta_{19}^{y}$ | 1.163 | 0.329 | -0.024 | -0.013 | 0.819 | 0.332 |
| $\delta_{20}^{y}$ | 1.294 | 0.257 | 0.217 | -0.034 | 0.169 | 0.804 |
| $\delta_{21}^{y}$ | 0.911 | 0.234 | 0.217 | -0.045 | 0.188 | 0.949 |
| $\delta_{22}^{y}$ | 1.098 | 0.233 | 0.158 | -0.049 | 0.334 | 0.491 |
| $\delta_{23}^{y}$ | 1.570 | 0.420 | 0.233 | -0.016 | 0.171 | 0.822 |
| $\delta_{24}^{y}$ | 1.228 | 0.310 | 0.314 | -0.015 | 0.125 | 0.478 |
| $\delta_{25}^{y}$ | 1.516 | 0.349 | 0.162 | 0.008 | 0.240 | 0.418 |
| $\delta_{26}^{9}$ | 1.103 | 0.247 | 0.018 | -0.034 | 1.078 | 0.178 |
| $\delta_{27}^{y}$ | 0.513 | 0.254 | -0.007 | 0.012 | 0.883 | 0.408 |
| $\delta_{28}^{y}$ | 1.021 | 0.357 | 0.237 | -0.058 | 0.164 | 0.215 |
| $\delta_{29}^{y}$ | 1.476 | 0.299 | 0.024 | -0.008 | 1.421 | 0.224 |
| $\delta_{30}^{y}$ | 0.868 | 0.303 | 0.049 | -0.029 | 0.650 | 0.533 |
| $\delta_{31}^{y}$ | 1.294 | 0.240 | 0.186 | -0.016 | 0.227 | 0.546 |
| $\delta_{32}^{y}$ | 0.667 | 0.427 | 0.403 | -0.025 | 0.101 | 0.596 |
| $\delta_{33}^{y}$ | 0.251 | 0.210 | -0.019 | -0.017 | 0.988 | 0.260 |
| $\delta_{34}^{y}$ | 1.039 | 0.313 | 0.116 | -0.017 | 0.292 | 0.632 |
| $\delta_{35}^{y}$ | 0.911 | 0.496 | -0.017 | -0.018 | 0.845 | 0.687 |
| $\delta_{36}^{y}$ | 1.247 | 0.302 | 0.303 | -0.015 | 0.133 | 0.362 |
| $\delta_{37}^{y}$ | 0.873 | 0.541 | 0.290 | -0.003 | 0.140 | 0.790 |
| $\delta_{38}^{y}$ | 0.814 | 0.296 | 0.060 | -0.019 | 0.745 | 0.242 |
| $\delta_{39}^{y}$ | 1.304 | 0.265 | 0.298 | -0.029 | 0.145 | 0.746 |
| $\delta_{40}^{y}$ | 1.066 | 0.307 | 0.055 | 0.026 | 0.538 | 0.445 |
| $\delta_{41}^{y}$ | 1.214 | 0.236 | 0.181 | -0.014 | 0.263 | 0.931 |
| $\delta_{42}^{y}$ | 0.693 | 0.405 | 0.000 | 0.005 | 0.618 | 0.505 |
| $\delta_{43}^{y}$ | 1.062 | 0.240 | 0.166 | 0.012 | 0.290 | 0.935 |
| $\delta_{44}^{y}$ | 0.396 | 0.281 | 0.303 | -0.018 | 0.124 | 0.627 |
| $\delta_{45}^{y}$ | 0.567 | 0.249 | 0.053 | 0.003 | 0.768 | 0.055 |
| $\delta_{46}^{y}$ | 1.415 | 0.285 | 0.108 | -0.032 | 0.311 | 0.591 |
| $\delta_{47}^{y}$ | 0.925 | 0.202 | 0.155 | 0.004 | 0.305 | 0.482 |
| $\delta_{48}^{y}$ | 1.132 | 0.262 | 0.061 | 0.019 | 0.465 | 0.564 |
| $\delta_{49}^{y}$ | 1.234 | 0.324 | 0.112 | -0.025 | 0.361 | 0.889 |
| $\delta_{50}^{y}$ | 1.072 | 0.227 | 0.140 | 0.016 | 0.272 | 0.804 |
| $\delta_{51}^{y}$ | 0.774 | 0.703 | 0.543 | -0.014 | 0.073 | 0.914 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; p-value: p-value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper for the standard error.

Table M.5: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\delta_{1}^{u}$ | 1.601 | 0.272 | 0.079 | -0.031 | 0.957 | 0.413 |
| $\delta_{2}^{u}$ | 0.406 | 0.148 | 0.008 | -0.043 | 1.292 | 0.946 |
| $\delta_{3}^{u}$ | 0.732 | 0.227 | 0.184 | -0.033 | 0.242 | 0.932 |
| $\delta_{4}^{u}$ | 0.747 | 0.186 | 0.109 | 0.013 | 0.425 | 0.807 |
| $\delta_{5}^{u}$ | 0.975 | 0.212 | 0.101 | -0.017 | 0.366 | 0.629 |
| $\delta_{6}^{u}$ | 0.619 | 0.199 | 0.121 | -0.044 | 0.347 | 0.972 |
| $\delta_{7}^{u}$ | 0.511 | 0.191 | 0.127 | 0.011 | 0.243 | 0.718 |
| $\delta_{8}^{u}$ | 0.844 | 0.214 | 0.036 | -0.020 | 1.181 | 0.199 |
| $\delta_{9}^{u}$ | 0.812 | 0.215 | 0.045 | 0.012 | 0.826 | 0.884 |
| $\delta_{10}^{u}$ | 1.083 | 0.219 | 0.036 | 0.018 | 0.641 | 0.495 |
| $\delta_{11}^{u}$ | 0.941 | 0.202 | 0.103 | -0.014 | 0.426 | 0.664 |
| $\delta_{12}^{u}$ | 0.680 | 0.192 | 0.039 | -0.029 | 0.888 | 0.338 |
| $\delta_{13}^{u}$ | 0.854 | 0.260 | 0.097 | 0.001 | 0.367 | 0.661 |
| $\delta_{14}^{u}$ | 1.187 | 0.242 | 0.043 | -0.005 | 0.752 | 0.426 |
| $\delta_{15}^{u}$ | 1.229 | 0.248 | 0.078 | -0.046 | 0.763 | 0.654 |
| $\delta_{16}^{u}$ | 0.649 | 0.170 | 0.068 | -0.008 | 0.534 | 0.694 |
| $\delta_{17}^{u}$ | 0.709 | 0.163 | 0.068 | -0.031 | 0.669 | 0.993 |
| $\delta_{18}^{u}$ | 1.184 | 0.236 | 0.078 | -0.019 | 0.764 | 0.386 |
| $\delta_{19}^{u}$ | 0.784 | 0.400 | 0.040 | 0.032 | 0.488 | 0.298 |
| $\delta_{20}^{u}$ | 0.805 | 0.190 | 0.095 | -0.023 | 0.608 | 0.693 |
| $\delta_{21}^{u}$ | 0.765 | 0.181 | 0.029 | -0.005 | 0.999 | 0.576 |
| $\delta_{22}^{u}$ | 0.679 | 0.173 | 0.033 | -0.016 | 0.854 | 0.901 |
| $\delta_{23}^{u}$ | 1.679 | 0.339 | 0.140 | -0.012 | 0.326 | 0.621 |
| $\delta_{24}^{u}$ | 0.785 | 0.181 | 0.066 | -0.032 | 0.725 | 0.544 |
| $\delta_{25}^{u}$ | 0.831 | 0.277 | 0.135 | 0.007 | 0.378 | 0.779 |
| $\delta_{26}^{u}$ | 1.117 | 0.210 | 0.098 | -0.019 | 0.570 | 0.687 |
| $\delta_{27}^{u}$ | 0.499 | 0.176 | 0.015 | -0.014 | 0.569 | 0.116 |
| $\delta_{28}^{u}$ | 0.460 | 0.131 | 0.029 | -0.047 | 0.885 | 0.488 |
| $\delta_{29}^{u}$ | 1.003 | 0.270 | 0.211 | -0.040 | 0.198 | 0.758 |
| $\delta_{30}^{u}$ | 0.705 | 0.182 | 0.108 | -0.005 | 0.510 | 0.508 |
| $\delta_{31}^{u}$ | 0.904 | 0.216 | 0.026 | 0.012 | 0.751 | 0.953 |
| $\delta_{32}^{u}$ | 0.908 | 0.215 | 0.075 | 0.005 | 0.836 | 0.257 |
| $\delta_{33}^{u}$ | 0.803 | 0.201 | 0.056 | -0.021 | 0.716 | 0.520 |
| $\delta_{34}^{u}$ | 1.257 | 0.243 | 0.089 | -0.018 | 0.701 | 0.613 |
| $\delta_{35}^{u}$ | 0.391 | 0.151 | 0.060 | -0.038 | 0.432 | 0.550 |
| $\delta_{36}^{u}$ | 1.223 | 0.232 | 0.073 | -0.016 | 0.537 | 0.745 |
| $\delta_{37}^{u}$ | 0.819 | 0.197 | 0.080 | 0.008 | 0.477 | 0.784 |
| $\delta_{38}^{u}$ | 1.307 | 0.289 | 0.218 | 0.018 | 0.242 | 0.971 |
| $\delta_{39}^{u}$ | 0.831 | 0.183 | 0.042 | -0.019 | 0.790 | 0.286 |
| $\delta_{40}^{u}$ | 0.974 | 0.245 | 0.042 | 0.026 | 0.938 | 0.515 |
| $\delta_{41}^{u}$ | 1.198 | 0.241 | 0.070 | 0.031 | 0.845 | 0.253 |
| $\delta_{42}^{u}$ | 0.582 | 0.162 | 0.041 | -0.052 | 0.788 | 0.788 |
| $\delta_{43}^{u}$ | 1.279 | 0.256 | 0.025 | 0.018 | 0.764 | 0.897 |
| $\delta_{44}^{u}$ | 0.766 | 0.191 | 0.078 | -0.023 | 0.619 | 0.151 |
| $\delta_{45}^{u}$ | 1.006 | 0.216 | 0.077 | -0.021 | 0.979 | 0.393 |
| $\delta_{46}^{u}$ | 0.647 | 0.171 | 0.061 | -0.028 | 0.423 | 0.870 |
| $\delta_{47}^{u}$ | 0.863 | 0.171 | 0.051 | -0.066 | 0.625 | 0.719 |
| $\delta_{48}^{u}$ | 0.946 | 0.223 | 0.120 | -0.025 | 0.420 | 0.683 |
| $\delta_{49}^{u}$ | 1.144 | 0.242 | 0.096 | -0.044 | 0.623 | 0.543 |
| $\delta_{50}^{u}$ | 1.115 | 0.216 | 0.080 | -0.009 | 0.569 | 0.549 |
| $\delta_{51}^{u}$ | 1.078 | 0.252 | 0.057 | -0.046 | 0.957 | 0.191 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; p-value: p-value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper for the standard error.

Table M.6: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\theta_{1,1}$ | -0.299 | 0.078 | -0.002 | -0.020 | 0.906 | 0.884 |
| $\theta_{1,2}$ | -0.092 | 0.064 | 0.094 | 0.007 | 0.577 | 0.494 |
| $\theta_{1,3}$ | -0.186 | 0.046 | 0.018 | -0.038 | 0.795 | 0.134 |
| $\theta_{1,4}$ | -0.094 | 0.050 | -0.015 | -0.004 | 1.109 | 0.176 |
| $\theta_{1,5}$ | -0.162 | 0.042 | -0.043 | -0.022 | 1.397 | 0.339 |
| $\theta_{1,6}$ | -0.212 | 0.059 | 0.013 | -0.032 | 0.620 | 0.629 |
| $\theta_{1,7}$ | -0.167 | 0.062 | 0.029 | -0.029 | 0.714 | 0.645 |
| $\theta_{1,8}$ | -0.165 | 0.066 | 0.051 | 0.005 | 0.650 | 0.252 |
| $\theta_{1,9}$ | -0.250 | 0.076 | -0.055 | -0.026 | 0.861 | 0.282 |
| $\theta_{1,10}$ | -0.187 | 0.047 | 0.072 | 0.003 | 0.773 | 0.422 |
| $\theta_{1,11}$ | -0.211 | 0.050 | -0.012 | 0.033 | 1.168 | 0.449 |
| $\theta_{1,12}$ | -0.251 | 0.066 | 0.014 | -0.021 | 0.720 | 0.191 |
| $\theta_{1,13}$ | -0.209 | 0.073 | 0.034 | 0.047 | 0.439 | 0.127 |
| $\theta_{1,14}$ | -0.247 | 0.069 | 0.028 | 0.001 | 0.961 | 0.974 |
| $\theta_{1,15}$ | -0.222 | 0.062 | 0.003 | -0.018 | 0.506 | 0.630 |
| $\theta_{1,16}$ | -0.111 | 0.059 | -0.001 | 0.013 | 1.252 | 0.512 |
| $\theta_{1,17}$ | -0.078 | 0.048 | 0.002 | -0.001 | 0.876 | 0.777 |
| $\theta_{1,18}$ | -0.199 | 0.073 | 0.022 | -0.022 | 0.767 | 0.787 |
| $\theta_{1,19}$ | -0.382 | 0.101 | 0.035 | 0.029 | 0.534 | 0.945 |
| $\theta_{1,20}$ | -0.235 | 0.070 | -0.016 | 0.006 | 1.052 | 0.806 |
| $\theta_{1,21}$ | -0.229 | 0.068 | -0.014 | 0.021 | 0.879 | 0.918 |
| $\theta_{1,22}$ | -0.196 | 0.060 | 0.042 | 0.019 | 0.552 | 0.611 |
| $\theta_{1,23}$ | -0.176 | 0.072 | 0.022 | 0.038 | 0.813 | 0.938 |
| $\theta_{1,24}$ | -0.114 | 0.056 | 0.007 | 0.030 | 0.872 | 0.462 |
| $\theta_{1,25}$ | -0.238 | 0.093 | 0.129 | -0.002 | 0.322 | 0.905 |
| $\theta_{1,26}$ | -0.193 | 0.070 | 0.022 | 0.055 | 0.755 | 0.403 |
| $\theta_{1,27}$ | -0.196 | 0.063 | -0.022 | 0.012 | 1.115 | 0.456 |
| $\theta_{1,28}$ | -0.057 | 0.036 | -0.031 | -0.011 | 0.879 | 0.866 |
| $\theta_{1,29}$ | -0.207 | 0.054 | -0.012 | -0.029 | 0.715 | 0.866 |
| $\theta_{1,30}$ | -0.093 | 0.050 | 0.024 | -0.024 | 0.572 | 0.701 |
| $\theta_{1,31}$ | -0.237 | 0.069 | -0.005 | 0.045 | 0.446 | 0.898 |
| $\theta_{1,32}$ | -0.200 | 0.081 | 0.043 | -0.006 | 0.465 | 0.391 |
| $\theta_{1,33}$ | -0.180 | 0.072 | -0.011 | -0.004 | 0.694 | 0.353 |
| $\theta_{1,34}$ | -0.231 | 0.063 | -0.005 | -0.026 | 0.630 | 0.382 |
| $\theta_{1,35}$ | -0.107 | 0.064 | 0.007 | 0.044 | 1.415 | 0.312 |
| $\theta_{1,36}$ | -0.165 | 0.067 | -0.021 | 0.019 | 0.921 | 0.107 |
| $\theta_{1,37}$ | -0.107 | 0.069 | 0.006 | -0.013 | 0.781 | 0.264 |
| $\theta_{1,38}$ | -0.209 | 0.067 | 0.097 | 0.045 | 0.322 | 0.467 |
| $\theta_{1,39}$ | -0.198 | 0.065 | 0.028 | 0.012 | 0.841 | 0.914 |
| $\theta_{1,40}$ | -0.267 | 0.074 | 0.002 | -0.006 | 1.039 | 0.625 |
| $\theta_{1,41}$ | -0.228 | 0.059 | -0.006 | -0.010 | 1.144 | 0.417 |
| $\theta_{1,42}$ | -0.082 | 0.060 | 0.072 | -0.005 | 0.865 | 0.863 |
| $\theta_{1,43}$ | -0.205 | 0.073 | -0.018 | 0.034 | 1.385 | 0.123 |
| $\theta_{1,44}$ | -0.170 | 0.065 | -0.011 | 0.018 | 0.891 | 0.152 |
| $\theta_{1,45}$ | -0.143 | 0.055 | 0.017 | -0.055 | 0.777 | 0.524 |
| $\theta_{1,46}$ | -0.160 | 0.068 | -0.015 | -0.020 | 1.011 | 0.378 |
| $\theta_{1,47}$ | -0.222 | 0.066 | -0.011 | 0.015 | 1.270 | 0.854 |
| $\theta_{1,48}$ | -0.172 | 0.060 | 0.043 | -0.040 | 0.794 | 0.212 |
| $\theta_{1,49}$ | -0.237 | 0.080 | 0.032 | -0.005 | 1.093 | 0.227 |
| $\theta_{1,50}$ | -0.168 | 0.065 | -0.006 | -0.013 | 0.732 | 0.469 |
| $\theta_{1,51}$ | -0.189 | 0.081 | 0.019 | 0.036 | 1.126 | 0.557 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; p -value: p -value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper for the standard error.

Table M.7: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\theta_{2,1}$ | -0.080 | 0.074 | 0.020 | 0.032 | 1.039 | 0.817 |
| $\theta_{2,2}$ | -0.035 | 0.049 | 0.069 | 0.027 | 0.605 | 0.615 |
| $\theta_{2,3}$ | -0.087 | 0.048 | 0.113 | 0.017 | 0.304 | 0.765 |
| $\theta_{2,4}$ | -0.069 | 0.045 | 0.012 | -0.021 | 1.387 | 0.716 |
| $\theta_{2,5}$ | -0.100 | 0.043 | 0.064 | -0.043 | 0.632 | 0.248 |
| $\theta_{2,6}$ | -0.123 | 0.058 | 0.075 | 0.012 | 0.387 | 0.539 |
| $\theta_{2,7}$ | -0.110 | 0.056 | 0.084 | 0.020 | 0.282 | 0.254 |
| $\theta_{2,8}$ | -0.095 | 0.051 | 0.006 | -0.009 | 0.607 | 0.381 |
| $\theta_{2,9}$ | -0.146 | 0.074 | -0.011 | 0.004 | 1.139 | 0.314 |
| $\theta_{2,10}$ | -0.057 | 0.046 | 0.007 | 0.019 | 0.911 | 0.307 |
| $\theta_{2,11}$ | -0.083 | 0.051 | 0.033 | -0.059 | 0.928 | 0.544 |
| $\theta_{2,12}$ | -0.103 | 0.064 | 0.018 | -0.006 | 0.637 | 0.361 |
| $\theta_{2,13}$ | -0.145 | 0.070 | 0.021 | -0.002 | 0.607 | 0.895 |
| $\theta_{2,14}$ | -0.084 | 0.067 | -0.029 | 0.032 | 1.294 | 0.826 |
| $\theta_{2,15}$ | -0.072 | 0.062 | 0.018 | 0.038 | 1.021 | 0.303 |
| $\theta_{2,16}$ | -0.041 | 0.050 | 0.012 | -0.027 | 0.656 | 0.525 |
| $\theta_{2,17}$ | -0.043 | 0.039 | 0.024 | 0.041 | 1.395 | 0.580 |
| $\theta_{2,18}$ | -0.093 | 0.068 | 0.041 | -0.005 | 1.293 | 0.718 |
| $\theta_{2,19}$ | -0.065 | 0.098 | -0.008 | 0.025 | 1.488 | 0.950 |
| $\theta_{2,20}$ | -0.091 | 0.066 | -0.046 | 0.019 | 0.818 | 0.270 |
| $\theta_{2,21}$ | -0.113 | 0.067 | 0.045 | -0.001 | 0.698 | 0.676 |
| $\theta_{2,22}$ | -0.122 | 0.056 | 0.010 | 0.014 | 0.643 | 0.488 |
| $\theta_{2,23}$ | -0.098 | 0.069 | 0.077 | 0.001 | 0.439 | 0.156 |
| $\theta_{2,24}$ | -0.056 | 0.053 | 0.004 | -0.015 | 1.129 | 0.466 |
| $\theta_{2,25}$ | -0.186 | 0.085 | 0.084 | 0.008 | 0.361 | 0.824 |
| $\theta_{2,26}$ | -0.113 | 0.065 | -0.022 | 0.001 | 1.574 | 0.639 |
| $\theta_{2,27}$ | -0.079 | 0.060 | 0.009 | -0.003 | 0.885 | 0.444 |
| $\theta_{2,28}$ | -0.033 | 0.035 | -0.001 | 0.025 | 0.716 | 0.654 |
| $\theta_{2,29}$ | -0.086 | 0.056 | 0.170 | -0.023 | 0.277 | 0.911 |
| $\theta_{2,30}$ | -0.073 | 0.048 | 0.037 | -0.000 | 0.564 | 0.647 |
| $\theta_{2,31}$ | -0.121 | 0.066 | 0.064 | 0.001 | 0.838 | 0.336 |
| $\theta_{2,32}$ | -0.087 | 0.070 | 0.070 | 0.051 | 0.851 | 0.386 |
| $\theta_{2,33}$ | -0.088 | 0.061 | -0.020 | 0.004 | 1.060 | 0.007 |
| $\theta_{2,34}$ | -0.042 | 0.059 | 0.003 | 0.022 | 0.783 | 0.059 |
| $\theta_{2,35}$ | -0.040 | 0.054 | 0.000 | 0.009 | 0.911 | 0.337 |
| $\theta_{2,36}$ | -0.071 | 0.062 | 0.037 | 0.008 | 0.751 | 0.428 |
| $\theta_{2,37}$ | -0.055 | 0.053 | -0.000 | 0.031 | 0.510 | 0.024 |
| $\theta_{2,38}$ | -0.075 | 0.065 | 0.159 | 0.003 | 0.274 | 0.713 |
| $\theta_{2,39}$ | -0.075 | 0.060 | -0.052 | 0.027 | 1.149 | 0.369 |
| $\theta_{2,40}$ | -0.165 | 0.063 | 0.035 | 0.022 | 0.942 | 0.644 |
| $\theta_{2,41}$ | -0.077 | 0.058 | 0.044 | 0.021 | 0.798 | 0.067 |
| $\theta_{2,42}$ | -0.048 | 0.046 | 0.032 | -0.003 | 1.052 | 0.980 |
| $\theta_{2,43}$ | -0.046 | 0.070 | 0.007 | 0.035 | 0.858 | 0.214 |
| $\theta_{2,44}$ | -0.083 | 0.056 | 0.020 | -0.033 | 0.952 | 0.575 |
| $\theta_{2,45}$ | -0.097 | 0.052 | 0.066 | 0.007 | 0.585 | 0.919 |
| $\theta_{2,46}$ | -0.095 | 0.064 | -0.007 | 0.031 | 1.171 | 0.967 |
| $\theta_{2,47}$ | -0.098 | 0.063 | 0.025 | -0.017 | 1.104 | 0.542 |
| $\theta_{2,48}$ | -0.068 | 0.054 | -0.030 | 0.020 | 0.926 | 0.684 |
| $\theta_{2,49}$ | -0.062 | 0.077 | 0.073 | 0.016 | 0.497 | 0.350 |
| $\theta_{2,50}$ | -0.061 | 0.063 | 0.045 | -0.000 | 0.525 | 0.035 |
| $\theta_{2,51}$ | -0.078 | 0.068 | 0.073 | -0.031 | 0.702 | 0.255 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; p -value: p -value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper for the standard error.

Table M.8: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\rho_{1}$ | 0.749 | 0.210 | -0.006 | 0.035 | 0.987 | 0.409 |
| $\rho_{2}$ | 0.616 | 0.299 | -0.011 | -0.019 | 1.043 | 0.428 |
| $\rho_{3}$ | 0.702 | 0.308 | 0.052 | -0.017 | 0.713 | 0.755 |
| $\rho_{4}$ | -0.108 | 0.321 | 0.016 | -0.027 | 1.026 | 0.613 |
| $\rho_{5}$ | 0.477 | 0.352 | -0.025 | 0.020 | 1.024 | 0.323 |
| $\rho_{6}$ | 0.706 | 0.304 | 0.005 | 0.052 | 1.021 | 0.400 |
| $\rho_{7}$ | 0.253 | 0.382 | -0.009 | 0.011 | 1.389 | 0.013 |
| $\rho_{8}$ | 0.253 | 0.437 | 0.016 | -0.027 | 0.956 | 0.663 |
| $\rho_{9}$ | 0.141 | 0.399 | 0.025 | -0.009 | 0.870 | 0.052 |
| $\rho_{10}$ | 0.818 | 0.203 | 0.013 | -0.008 | 0.880 | 0.607 |
| $\rho_{11}$ | 0.578 | 0.326 | 0.014 | -0.026 | 1.044 | 0.803 |
| $\rho_{12}$ | 0.741 | 0.201 | 0.015 | -0.010 | 1.225 | 0.978 |
| $\rho_{13}$ | 0.344 | 0.338 | 0.058 | 0.004 | 0.978 | 0.670 |
| $\rho_{14}$ | 0.704 | 0.241 | 0.013 | 0.024 | 0.667 | 0.741 |
| $\rho_{15}$ | 0.790 | 0.171 | 0.018 | 0.009 | 1.295 | 0.122 |
| $\rho_{16}$ | 0.729 | 0.217 | -0.007 | -0.048 | 1.246 | 0.936 |
| $\rho_{17}$ | 0.230 | 0.358 | -0.006 | 0.032 | 1.259 | 0.086 |
| $\rho_{18}$ | 0.599 | 0.286 | -0.075 | -0.039 | 1.398 | 0.551 |
| $\rho_{19}$ | 0.732 | 0.230 | 0.028 | 0.011 | 0.656 | 0.095 |
| $\rho_{20}$ | 0.717 | 0.288 | -0.019 | 0.004 | 1.368 | 0.559 |
| $\rho_{21}$ | 0.673 | 0.282 | 0.031 | 0.007 | 1.832 | 0.899 |
| $\rho_{22}$ | 0.412 | 0.339 | 0.008 | -0.014 | 1.099 | 0.975 |
| $\rho_{23}$ | 0.702 | 0.332 | -0.012 | -0.002 | 1.346 | 0.816 |
| $\rho_{24}$ | 0.421 | 0.375 | 0.004 | -0.014 | 1.100 | 0.218 |
| $\rho_{25}$ | 0.498 | 0.377 | -0.017 | 0.022 | 1.022 | 0.708 |
| $\rho_{26}$ | 0.081 | 0.400 | 0.014 | 0.011 | 0.608 | 0.618 |
| $\rho_{27}$ | 0.583 | 0.327 | -0.013 | -0.023 | 1.015 | 0.596 |
| $\rho_{28}$ | -0.130 | 0.266 | 0.042 | 0.007 | 1.243 | 0.687 |
| $\rho_{29}$ | 0.522 | 0.325 | 0.048 | -0.039 | 0.551 | 0.378 |
| $\rho_{30}$ | -0.390 | 0.252 | 0.007 | 0.009 | 0.903 | 0.897 |
| $\rho_{31}$ | 0.581 | 0.325 | 0.042 | 0.012 | 0.931 | 0.199 |
| $\rho_{32}$ | 0.507 | 0.322 | 0.005 | 0.011 | 1.088 | 0.404 |
| $\rho_{33}$ | 0.466 | 0.367 | -0.030 | -0.033 | 0.896 | 0.279 |
| $\rho_{34}$ | 0.548 | 0.263 | -0.003 | 0.009 | 1.217 | 0.509 |
| $\rho_{35}$ | 0.595 | 0.369 | 0.049 | 0.019 | 0.978 | 0.740 |
| $\rho_{36}$ | 0.750 | 0.240 | -0.007 | 0.016 | 1.187 | 0.506 |
| $\rho_{37}$ | 0.381 | 0.420 | -0.030 | -0.061 | 0.932 | 0.136 |
| $\rho_{38}$ | 0.485 | 0.351 | -0.001 | 0.008 | 1.394 | 0.642 |
| $\rho_{39}$ | 0.629 | 0.273 | -0.008 | -0.018 | 1.065 | 0.477 |
| $\rho_{40}$ | 0.765 | 0.250 | 0.058 | -0.009 | 0.808 | 0.353 |
| $\rho_{41}$ | 0.760 | 0.192 | 0.012 | 0.001 | 0.826 | 0.268 |
| $\rho_{42}$ | 0.192 | 0.387 | -0.011 | -0.032 | 1.404 | 0.249 |
| $\rho_{43}$ | 0.715 | 0.235 | 0.043 | -0.023 | 1.049 | 0.836 |
| $\rho_{44}$ | 0.775 | 0.212 | 0.020 | -0.019 | 1.171 | 0.430 |
| $\rho_{45}$ | 0.762 | 0.278 | 0.005 | 0.041 | 1.111 | 0.221 |
| $\rho_{46}$ | 0.276 | 0.436 | 0.051 | -0.008 | 0.476 | 0.396 |
| $\rho_{47}$ | 0.750 | 0.214 | -0.004 | -0.023 | 0.961 | 0.622 |
| $\rho_{48}$ | 0.631 | 0.311 | 0.008 | 0.019 | 1.079 | 0.656 |
| $\rho_{49}$ | 0.700 | 0.202 | -0.003 | 0.040 | 1.008 | 0.756 |
| $\rho_{50}$ | 0.544 | 0.334 | 0.001 | 0.008 | 1.712 | 0.671 |
| $\rho_{51}$ | 0.677 | 0.288 | -0.014 | -0.013 | 0.856 | 0.242 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; p-value: p-value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper for the standard error.

Table M.9: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\sigma_{\zeta_{1}}^{2}$ | 0.170 | 0.051 | 0.029 | 0.032 | 0.886 | 0.182 |
| $\sigma_{\zeta_{2}}^{2}$ | 0.667 | 0.758 | 0.134 | 0.009 | 1.104 | 0.403 |
| $\sigma^{2}$ | 0.249 | 0.094 | 0.085 | -0.016 | 0.450 | 0.840 |
| $\sigma_{\zeta_{4}}^{2}$ | 0.402 | 0.193 | 0.013 | 0.030 | 0.773 | 0.291 |
| $\sigma_{5}^{2}$ | 0.209 | 0.079 | 0.034 | 0.017 | 0.545 | 0.509 |
| $\sigma_{\zeta_{6}}^{2}$ | 0.220 | 0.082 | -0.004 | 0.031 | 0.909 | 0.964 |
| $\sigma_{\zeta_{7}}^{2}$ | 0.266 | 0.132 | 0.037 | 0.000 | 0.616 | 0.079 |
| $\sigma_{\zeta_{8}}^{2}$ | 0.413 | 0.585 | 0.285 | -0.013 | 0.586 | 0.137 |
| $\sigma_{\zeta_{9}}^{2}$ | 0.203 | 0.072 | 0.065 | 0.001 | 0.880 | 0.119 |
| $\sigma_{\zeta_{10}}^{2}$ | 0.170 | 0.054 | 0.043 | -0.045 | 0.662 | 0.988 |
| $\sigma_{\zeta_{1}}^{2}$ | 0.149 | 0.046 | -0.008 | -0.013 | 0.994 | 0.384 |
| $\sigma_{\zeta_{12}}^{2}$ | 0.154 | 0.044 | -0.020 | 0.041 | 1.422 | 0.556 |
| $\sigma_{\zeta_{13}}^{2}$ | 0.237 | 0.091 | 0.008 | -0.003 | 0.605 | 0.312 |
| $\sigma_{\zeta_{14}}^{2}$ | 0.156 | 0.045 | 0.010 | 0.033 | 1.223 | 0.809 |
| $\sigma_{\zeta_{15}}^{2}$ | 0.220 | 0.079 | -0.015 | 0.011 | 0.747 | 0.902 |
| $\sigma_{\zeta_{1}}^{2}$ | 0.284 | 0.144 | 0.011 | -0.006 | 0.824 | 0.280 |
| $\sigma_{\zeta_{17}}^{2}$ | 0.428 | 0.251 | -0.003 | -0.038 | 1.026 | 0.250 |
| $\sigma_{\zeta_{18}}^{2}$ | 0.163 | 0.056 | -0.005 | -0.035 | 0.667 | 0.563 |
| $\sigma_{\zeta_{19}}^{2}$ | 0.394 | 0.180 | 0.029 | 0.002 | 0.994 | 0.495 |
| $\sigma_{\zeta_{20}}^{2}$ | 0.179 | 0.061 | 0.023 | 0.029 | 1.063 | 0.640 |
| $\sigma_{\zeta_{21}}^{2}$ | 0.158 | 0.050 | 0.016 | -0.000 | 1.126 | 0.996 |
| $\sigma_{\zeta_{22}}^{2}$ | 0.206 | 0.069 | 0.039 | 0.025 | 0.852 | 0.735 |
| $\sigma_{\zeta_{23}}^{2}$ | 0.423 | 0.216 | -0.011 | 0.000 | 0.681 | 0.779 |
| $\sigma_{\zeta_{24}}^{2}$ | 0.244 | 0.092 | -0.015 | 0.040 | 0.885 | 0.192 |
| $\sigma_{\zeta_{25}}^{2}$ | 0.284 | 0.129 | -0.032 | -0.006 | 0.766 | 0.456 |
| $\sigma_{\zeta_{26}}^{2}$ | 0.193 | 0.069 | 0.008 | -0.007 | 0.965 | 0.030 |
| $\sigma_{\zeta_{27}}^{2}$ | 0.255 | 0.112 | -0.022 | 0.011 | 1.000 | 0.032 |
| $\sigma_{\zeta_{28}}^{2}$ | 0.475 | 0.193 | -0.025 | -0.036 | 1.189 | 0.370 |
| $\sigma_{\zeta_{29}}^{2}$ | 0.245 | 0.115 | 0.041 | -0.040 | 0.542 | 0.139 |
| $\sigma_{\zeta_{30}}^{2}$ | 0.305 | 0.128 | 0.023 | -0.032 | 0.692 | 0.754 |
| $\sigma_{\zeta_{31}}^{2}$ | 0.169 | 0.056 | 0.019 | 0.065 | 0.552 | 0.828 |
| $\sigma_{\zeta_{32}}^{2}$ | 0.290 | 0.164 | 0.044 | -0.033 | 0.893 | 0.676 |
| $\sigma_{\zeta_{33}}^{2}$ | 0.269 | 0.174 | 0.021 | 0.003 | 1.016 | 0.503 |
| $\sigma_{\zeta_{34}}^{2}$ | 0.276 | 0.123 | 0.035 | 0.011 | 0.915 | 0.407 |
| $\sigma_{\zeta_{35}}^{2}$ | 0.496 | 0.658 | 0.170 | 0.007 | 1.032 | 0.831 |
| $\sigma_{\zeta_{36}}^{2}$ | 0.210 | 0.076 | -0.024 | -0.002 | 0.750 | 0.001 |
| $\sigma_{\zeta_{37}}^{2}$ | 0.580 | 0.574 | 0.040 | -0.027 | 0.937 | 0.789 |
| $\sigma_{\chi_{38}}^{2}$ | 0.250 | 0.101 | 0.034 | 0.037 | 0.512 | 0.689 |
| $\sigma_{\zeta_{39}}^{2}$ | 0.169 | 0.053 | -0.027 | 0.003 | 0.855 | 0.261 |
| $\sigma_{\zeta_{40}}^{2}$ | 0.350 | 0.158 | 0.031 | 0.014 | 1.078 | 0.931 |
| $\sigma_{\zeta_{41}}^{L^{40}}$ | 0.154 | 0.045 | -0.006 | -0.008 | 1.103 | 0.984 |
| $\sigma_{\zeta_{42}}^{2}$ | 0.829 | 0.729 | 0.099 | 0.001 | 0.882 | 0.403 |
| $\sigma_{\zeta_{43}}^{2^{42}}$ | 0.174 | 0.058 | 0.021 | 0.003 | 1.041 | 0.727 |
| $\sigma_{\zeta_{44}}^{2}$ | 0.332 | 0.152 | 0.025 | 0.030 | 0.795 | 0.943 |
| $\sigma_{\zeta_{45}}^{2}$ | 0.236 | 0.093 | 0.003 | 0.013 | 1.086 | 0.833 |
| $\sigma_{\zeta_{46}}^{2}$ | 0.194 | 0.078 | 0.018 | -0.012 | 0.591 | 0.391 |
| $\sigma_{\zeta_{47}}^{2}$ | 0.147 | 0.044 | -0.006 | -0.036 | 0.988 | 0.455 |
| $\sigma_{\zeta_{48}}^{2}$ | 0.225 | 0.082 | 0.035 | -0.013 | 0.843 | 0.252 |
| $\sigma_{\zeta_{49}}^{2}$ | 0.333 | 0.149 | 0.010 | -0.000 | 0.780 | 0.525 |
| $\sigma_{\zeta_{50}}^{2}$ | 0.143 | 0.043 | -0.030 | 0.037 | 0.895 | 0.083 |
| $\sigma_{\zeta_{51}}^{2}$ | 0.600 | 0.807 | 0.242 | 0.012 | 0.536 | 0.732 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; $p$-value: $p$-value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper for the standard error.

Table M.10: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\sigma_{\eta_{1}^{\prime}}^{2}$ | 0.169 | 0.050 | 0.032 | -0.030 | 0.915 | 0.066 |
| $\sigma^{2}{ }^{2}$ | 3.804 | 1.398 | 0.255 | 0.003 | 0.188 | 0.665 |
| $\sigma^{2}{ }^{2}$ | 0.285 | 0.098 | 0.039 | 0.029 | 0.807 | 0.432 |
| $\sigma^{2}{ }^{2}$ | 0.689 | 0.292 | -0.005 | 0.015 | 0.967 | 0.281 |
| $\sigma^{\eta^{2}}$ | 0.316 | 0.117 | 0.085 | -0.007 | 0.403 | 0.367 |
| $\sigma^{2}{ }^{5}$ | 0.308 | 0.100 | -0.012 | 0.005 | 1.653 | 0.328 |
| $\sigma^{2}{ }^{2}{ }^{2}$ | 1.168 | 0.314 | -0.002 | 0.011 | 0.874 | 0.480 |
| $\sigma^{\eta_{7}^{2}}$ | 5.712 | 1.464 | 0.063 | -0.051 | 1.041 | 0.656 |
| $\sigma^{2}{ }_{y}$ | 0.382 | 0.123 | 0.032 | -0.018 | 1.332 | 0.004 |
| $\sigma^{2}{ }^{\eta_{9}^{y}}$ | 0.201 | 0.063 | 0.046 | -0.004 | 0.396 | 0.279 |
| $\sigma_{n^{y}}^{2^{10}}$ | 0.219 | 0.064 | 0.014 | -0.031 | 1.136 | 0.499 |
| $\sigma^{2}{ }^{11}$ | 0.188 | 0.054 | -0.002 | 0.041 | 1.143 | 0.607 |
| $\sigma_{\eta^{2}}^{\eta_{12}}$ | 0.529 | 0.162 | -0.008 | 0.006 | 1.144 | 0.561 |
| $\sigma^{2}{ }^{2}{ }^{13}$ | 0.199 | 0.058 | 0.058 | -0.004 | 1.046 | 0.907 |
| $\sigma_{\eta^{1}}^{\eta^{14}}$ | 0.378 | 0.122 | 0.021 | -0.028 | 0.752 | 0.812 |
| $\sigma_{\eta^{y}}^{\eta_{15}}$ | 0.712 | 0.225 | -0.028 | -0.035 | 0.920 | 0.992 |
| $\sigma^{\eta^{2}}{ }^{2}$ | 1.342 | 0.453 | -0.005 | 0.024 | 1.030 | 0.590 |
|  | 0.263 | 0.086 | 0.045 | 0.015 | 0.562 | 0.349 |
|  | 0.745 | 0.254 | 0.049 | -0.006 | 0.752 | 0.183 |
| $\sigma^{\eta^{2}}{ }^{19}$ | 0.265 | 0.081 | 0.018 | -0.030 | 0.663 | 0.704 |
| $\sigma^{\eta_{20}}{ }^{2}$ | 0.245 | 0.072 | 0.020 | 0.010 | 0.877 | 0.101 |
| ${ }^{\eta_{21}^{2}}$ | 0.249 | 0.083 | -0.006 | 0.033 | 1.052 | 0.083 |
| $\sigma^{\eta_{22}}$ | 0.765 | 0.270 | 0.015 | -0.017 | 0.556 | 0.789 |
| $\sigma^{2}{ }_{y}$ | 0.395 | 0.130 | -0.013 | -0.019 | 1.808 | 0.527 |
| ${ }^{\eta_{24}^{2}}$ | 0.688 | 0.204 | -0.017 | 0.029 | 1.405 | 0.796 |
| ${ }_{\square}^{\eta_{25}^{2}}$ | 0.320 | 0.104 | 0.021 | 0.005 | 0.835 | 0.259 |
|  | 0.521 | 0.161 | 0.015 | 0.008 | 0.995 | 0.270 |
| ${ }^{\eta_{27}^{y}}{ }^{2}{ }^{2}$ | 0.608 | 0.266 | -0.016 | -0.001 | 0.900 | 0.523 |
| ${ }^{\eta_{28}^{y}}$ | 0.608 | 0.266 | -0.016 | -0.000 | 0.900 | 0.523 0.639 |
| $\sigma_{\eta} \eta_{29}^{y}$ | . 485 | 0.154 | 0.026 | -0.006 | 0.862 | 0.639 |
| $\sigma^{2}{ }^{2}{ }_{30}$ | 0.547 | 0.203 | -0.005 | -0.006 | 1.076 | 0.767 |
| $\sigma^{2}{ }^{20}$ | 0.252 | 0.074 | -0.012 | -0.003 | 1.710 | 0.302 |
| $\sigma^{2}{ }_{y}$ | 0.980 | 0.318 | 0.133 | -0.047 | 0.270 | 0.327 |
| $\sigma^{2}{ }_{y}^{\eta_{32}}$ | 1.756 | 0.474 | 0.029 | -0.026 | 0.888 | 0.793 |
| ${ }^{\eta^{2}{ }_{y}^{3}}$ | 0.646 | 0.202 | 0.009 | -0.025 | 1.053 | 0.906 |
| ${ }^{\eta^{2}{ }_{y 4}}$ | 4.445 | 1.167 | 0.045 | 0.023 | 1.696 | 0.468 |
| $\sigma_{\eta 5}^{y}$ | 4.445 | 1.167 | 0.045 | 0.023 | 1.696 | 0.468 |
| $\sigma_{\eta_{36}}^{2}$ | 0.332 | 0.100 | 0.031 | 0.005 | 0.704 | 0.000 |
| $\sigma_{\eta^{2}}^{2}$ | 3.031 | 0.989 | 0.058 | -0.057 | 0.633 | 0.462 |
| $\sigma^{2}{ }^{27}$ | 0.547 | 0.172 | 0.027 | -0.013 | 0.622 | 0.200 |
| $\sigma^{2}{ }^{28}$ | 0.232 | 0.071 | -0.033 | -0.033 | 0.847 | 0.487 |
| $\sigma^{2}{ }_{y}{ }^{29}$ | 0.613 | 0.205 | 0.003 | 0.050 | 0.941 | 0.108 |
| $\sigma^{2}{ }^{2}{ }^{2}$ | 0.173 | 0.052 | 0.007 | 0.005 | 0.997 | 0.076 |
| ${ }^{\eta^{2}{ }^{2}}$ | 2.424 | 1.149 | 0.055 | 0.002 | 0.771 | 0.435 |
| $\sigma^{2}{ }^{2}{ }^{2}$ | 0.286 | 0.087 | -0.007 | -0.016 | 0.741 | 0.244 |
| $\sigma^{\eta^{2}}{ }^{43}$ | 0.532 | 0.189 | 0.020 | 0.008 | 0.930 | 0.922 |
| ${ }^{\eta_{44}^{2}}$ | 0.396 | 0.134 | 0.037 | -0.007 | 0.479 | 0.271 |
|  | 0.301 | 0.107 | 0.049 | -0.012 | 0.384 | 0.458 |
| $\eta_{46}^{y}$ $\sigma^{2}$ | 0.175 | 0.050 | -0.010 | 0.043 | 0.384 0.796 | 0.458 0.360 |
| $\Gamma_{-2}^{\eta_{47}^{y}}$ | 0.175 | 0.050 |  | 0.043 | 0.796 | 0.360 |
| $\sigma^{2}{ }^{4}$ | 0.364 | 0.113 | -0.024 | -0.000 | 0.984 | 0.860 |
| $\sigma^{2}{ }^{48}$ | 0.471 | 0.187 | 0.055 | -0.002 | 0.457 | 0.968 |
| $\sigma^{2}{ }^{29}$ | 0.183 | 0.053 | -0.008 | -0.000 | 1.321 | 0.189 |
| $\begin{gathered} \eta_{50}^{y} \\ \sigma_{\eta_{51}^{2}}^{2} \end{gathered}$ | 5.062 | 1.443 | 2.230 | -0.008 | 0.213 | 0.785 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; p-value: p-value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper

Table M.11: Posterior Distribution Estimates and Convergence Diagnostics

| Parameter | Mean | Sd. | auto1 | auto50 | RNE | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\sigma_{\eta_{1}^{u}}^{\sim}$ | 0.086 | 0.018 | 0.000 | -0.004 | 0.870 | 0.191 |
| $\sigma_{\eta_{2}^{u}}^{2}$ | 0.048 | 0.010 | -0.011 | -0.048 | 1.116 | 0.178 |
| $\sigma_{\eta_{3}}^{2}$ | 0.063 | 0.013 | 0.052 | -0.029 | 0.847 | 0.007 |
| $\sigma_{\eta_{4}^{2}}^{2}$ | 0.063 | 0.013 | -0.027 | -0.013 | 1.236 | 0.126 |
| $\sigma_{\eta_{5}^{u}}^{2}$ | 0.060 | 0.013 | 0.000 | -0.028 | 1.044 | 0.044 |
| $\sigma_{\eta_{6}{ }^{u}}$ | 0.070 | 0.015 | -0.007 | -0.012 | 1.975 | 0.778 |
| $\sigma_{\eta_{7}^{u}}{ }^{\text {u }}$ | 0.060 | 0.013 | 0.002 | -0.024 | 1.821 | 0.292 |
| $\sigma_{\eta_{8}^{u}}^{2}$ | 0.060 | 0.013 | 0.030 | -0.031 | 0.870 | 0.641 |
| $\sigma_{\eta_{g}^{u}}^{2}$ | 0.093 | 0.020 | 0.070 | 0.011 | 0.922 | 0.131 |
| $\sigma_{\eta_{10}^{u}}^{2}$ | 0.061 | 0.013 | 0.006 | 0.014 | 0.950 | 0.192 |
| $\sigma_{\eta_{11}^{u}}^{2}$ | 0.055 | 0.011 | -0.002 | 0.015 | 0.740 | 0.911 |
| $\sigma_{\eta_{12}^{\prime}}^{2}$ | 0.061 | 0.013 | 0.046 | -0.009 | 0.845 | 0.148 |
| $\sigma_{\eta_{13}^{\prime}}^{2}$ | 0.120 | 0.025 | -0.011 | 0.028 | 1.126 | 0.186 |
| $\sigma_{\eta_{14}^{\prime}}^{2}$ | 0.084 | 0.018 | 0.046 | -0.014 | 0.942 | 0.388 |
| $\sigma_{\eta_{15}^{u}}^{2}$ | 0.068 | 0.014 | 0.004 | 0.018 | 1.047 | 0.654 |
| $\sigma_{\eta_{16}}^{2}$ | 0.049 | 0.010 | 0.013 | -0.010 | 0.968 | 0.020 |
| $\sigma_{\eta_{17}^{\prime}}^{2}$ | 0.049 | 0.010 | 0.014 | -0.018 | 1.322 | 0.595 |
| $\sigma_{\eta_{18}^{u}}^{2}$ | 0.072 | 0.015 | 0.032 | 0.011 | 1.091 | 0.640 |
| $\sigma_{\eta_{19}^{u}}^{2}$ | 0.618 | 0.133 | 0.038 | -0.044 | 1.026 | 0.308 |
| $\sigma_{\eta_{20}^{u}}^{2}$ | 0.057 | 0.012 | 0.017 | 0.004 | 1.118 | 0.812 |
| $\sigma_{\eta_{21}^{u}}^{2}$ | 0.053 | 0.011 | 0.005 | -0.040 | 1.005 | 0.379 |
| $\sigma_{\eta_{22}}^{2}$ | 0.058 | 0.012 | 0.014 | -0.023 | 0.935 | 0.609 |
| $\sigma_{\eta_{23}^{u}}^{2}$ | 0.089 | 0.020 | 0.016 | 0.017 | 1.047 | 0.204 |
| $\sigma_{\eta_{24}^{u}}^{2}$ | 0.056 | 0.011 | 0.024 | 0.021 | 1.318 | 0.587 |
| $\sigma_{\eta_{25}^{u}}^{2 u}$ | 0.142 | 0.032 | 0.016 | -0.010 | 0.964 | 0.767 |
| $\sigma_{\eta_{26}^{\prime}}^{2}$ | 0.067 | 0.014 | -0.011 | -0.008 | 1.277 | 0.775 |
| $\sigma_{\eta_{27}^{u}}^{2}$ | 0.062 | 0.013 | -0.007 | 0.054 | 0.680 | 0.655 |
| $\sigma_{\eta_{28}^{u}}^{2}$ | 0.046 | 0.009 | -0.019 | -0.010 | 0.997 | 0.032 |
| $\sigma_{\eta_{29}^{u}}^{2}$ | 0.087 | 0.019 | 0.002 | -0.002 | 1.310 | 0.197 |
| $\sigma_{\eta_{30}}^{2}$ | 0.058 | 0.012 | -0.001 | 0.008 | 0.899 | 0.160 |
| $\sigma_{\eta_{31}^{\prime}}^{2}$ | 0.072 | 0.015 | -0.014 | -0.012 | 0.902 | 0.196 |
| $\sigma_{\eta_{32}}^{2}$ | 0.073 | 0.016 | -0.009 | -0.015 | 1.331 | 0.822 |
| $\sigma_{\eta_{33}^{\prime}}^{2}$ | 0.061 | 0.013 | 0.033 | 0.011 | 0.626 | 0.821 |
| $\sigma_{\eta_{34}}^{2}$ | 0.070 | 0.015 | -0.012 | -0.012 | 1.014 | 0.018 |
| $\sigma_{\eta_{35}^{u}}^{2^{34}}$ | 0.048 | 0.010 | -0.002 | 0.007 | 1.550 | 0.776 |
| $\sigma_{\eta_{36}^{u}}^{2}$ | 0.068 | 0.015 | -0.017 | 0.007 | 1.009 | 0.453 |
| $\sigma_{\eta_{37}^{\prime}}^{2}$ | 0.061 | 0.013 | 0.005 | 0.061 | 1.165 | 0.177 |
| $\sigma_{\eta_{38}^{\prime}}^{2^{37}}$ | 0.088 | 0.019 | 0.056 | -0.009 | 0.587 | 0.688 |
| $\sigma_{\eta_{39}}^{2}$ | 0.054 | 0.011 | 0.039 | 0.005 | 0.821 | 0.976 |
| $\sigma_{\eta_{40}}^{2}$ | 0.081 | 0.019 | -0.003 | 0.004 | 0.997 | 0.144 |
| $\sigma_{\eta_{41}^{4}}^{2}$ | 0.073 | 0.016 | 0.003 | -0.031 | 1.173 | 0.768 |
| $\sigma_{\eta_{42}}^{2}$ | 0.053 | 0.011 | 0.029 | 0.001 | 1.180 | 0.274 |
| $\sigma_{\eta_{43}^{u}}^{2^{42}}$ | 0.093 | 0.020 | -0.004 | -0.012 | 1.177 | 0.211 |
| $\sigma_{\eta_{44}}^{2}$ | 0.064 | 0.013 | -0.014 | 0.026 | 1.015 | 0.844 |
| $\sigma_{\eta_{45}^{u}}^{2}$ | 0.069 | 0.015 | 0.019 | 0.026 | 1.424 | 0.557 |
| $\sigma_{\eta_{46}^{u}}^{2}$ | 0.056 | 0.011 | 0.024 | -0.019 | 1.396 | 0.524 |
| $\sigma_{\eta_{47}^{u}}^{2}$ | 0.049 | 0.010 | 0.007 | -0.024 | 1.063 | 0.594 |
| $\sigma_{\eta_{48}^{2}}^{2}$ | 0.068 | 0.015 | 0.040 | 0.034 | 0.717 | 0.209 |
| $\sigma_{\eta_{49}^{4}}^{2}$ | 0.091 | 0.020 | 0.032 | 0.041 | 0.898 | 0.325 |
| $\sigma_{\eta_{50}}^{2}$ | 0.059 | 0.013 | 0.006 | 0.011 | 0.845 | 0.509 |
| $\sigma_{\eta_{51}^{2}}^{2^{u}}$ | 0.080 | 0.018 | -0.000 | -0.006 | 1.034 | 0.323 |

Note: Statistics are based on 2,000 draws from the Gibbs sampler. Mean: average across draws; Sd.: standard deviation across draws; auto1: first order autocorrelation across draws; auto50: fiftieth order autocorrelation across draws; RNE: relative numerical efficiency using a 4 percent taper; $p$-value: $p$-value of the null hypothesis that the mean of the first 20 percent of draws is equal to the mean of the last 50 percent of draws using a 4 percent taper for the standard error.

Table M.12: Prior and Posterior Distributions of the Parameters using Aggregate Data

| Parameter | Prior Distribution | Prior Mean | Prior Standard Deviation | Posterior Mean | Posterior Standard Deviation |
| ---: | :---: | :---: | :---: | :---: | :---: |
| $\mu$ | Normal | 0.8 | 0.8 | 0.40 | 0.07 |
| $\phi_{1}$ | Truncated Normal | 1.5 | 0.5 | 1.55 | 0.13 |
| $\phi_{2}$ | Truncated Normal | -0.6 | 0.2 | -0.59 | 0.13 |
| $\theta_{1}$ | Normal | -0.4 | 0.1 | -0.39 | 0.07 |
| $\theta_{2}$ | Normal | -0.1 | 0.1 | -0.20 | 0.07 |
| $\sigma_{\varepsilon}^{2}$ | Inverse Gamma | 1 | $\operatorname{Inf}$ | 0.21 | 0.06 |
| $\sigma_{\eta^{y}}^{2}$ | Inverse Gamma | 1 | $\operatorname{Inf}$ | 0.18 | 0.05 |
| $\sigma_{\eta^{u}}^{2}$ | Inverse Gamma | 1 | $\operatorname{Inf}$ | 0.07 | 0.02 |

Figure M.1: States with Negative Average Real GDP Annual Growth in the 2007-2009 Period


Note: States in light gray experienced positive average growth rates. Source: BEA Regional Economic Accounts.

Figure M.2: Average Real GDP Annual Growth by State in the 2010-2016 Period


Source: BEA Regional Economic Accounts.

Figure M.3: Variance Decomposition of the State's Cycles


Note: $\mathrm{x}=100 \times \frac{\alpha_{i}^{2} \operatorname{var}\left(c_{t}\right)}{\operatorname{var}\left(c_{i t}\right)}$ is the proportion of the variance of the state's cycle that is explained by the variance of the common cycle

Figure M.4: Variance Decomposition of the State's Output Trends


Note: $\mathrm{x}=100 \times \frac{\left(\delta_{i}^{y}\right)^{2}}{\left(\delta_{i}^{y}\right)^{2}+\sigma_{\eta}^{2}}$ is the proportion of the variance of the
state's output trend that is explained by the variance of the common output trend.

Figure M.5: Variance Decomposition of the State's GDP Growth


Table M.13: Variance Decomposition of the State's GDP Growth Explained by Industry Composition, Demographics, and State Policy

| Variable | Coefficient |
| :--- | :---: |
| Intercept | $-1.52^{* *}$ |
| Agriculture, forestry, fishing, and hunting | 0.50 |
| Mining, quarrying, and oil and gas extraction | -0.55 |
| Construction | -1.64 |
| Manufacturing | 0.16 |
| Wholesale trade; Retail trade; Transportation and warehousing | $2.30^{*}$ |
| Finance and insurance; Real estate and rental and leasing | -0.04 |
| Government and government enterprises | -1.23 |
| Share of 25+ population with a bachelor's degree | -0.35 |
| State's personal income tax as a share of personal income | -1.26 |
| Share of population between the ages of 18 and 44 | $3.34^{* * *}$ |
| Debt-to-income ratio | $0.85^{* * *}$ |

Note: The dependent variable is the proportion of the variance of the state's real GDP growth rate that is explained by the variance of it cycle. Explanatory variables are expressed in their $2005-2017$ averages.
$* * *, * *$, and $*$ denote statistical significance at the 1,5, and 10 percent levels, respectively. $\bar{R}^{2}=0.55$.

[^3]Table M.14: State Labor Market Cyclicality Explained by Industry Composition, Demographics, State Policy, and Union Membership

| Variable | Coefficient |
| :--- | ---: |
| Intercept | 0.36 |
| Agriculture, forestry, fishing, and hunting | $-1.82^{* * *}$ |
| Mining, quarrying, and oil and gas extraction | $-0.46^{*}$ |
| Construction | $3.68^{* * *}$ |
| Manufacturing | 0.05 |
| Wholesale trade; Retail trade; Transportation and warehousing | -1.07 |
| Finance and insurance, Real estate and rental and leasing | -0.23 |
| Government and government enterprises | 0.24 |
| Share of 25+ population with a bachelor's degree | -0.45 |
| State's personal income tax as a share of personal income | 0.70 |
| Share of population between the ages of 18 and 44 | 0.10 |
| Union membership | 0.00 |
| Note: The dependent variable is the absolute value of the sum $\theta_{1 i}+\theta_{2 i}$ |  |
| for each state. Explanatory variables are expressed in their 2005-2017 |  |
| averages. ***, **, and $*$ denote statistical significance at the 1, 5, and |  |
| 10 percent levels, respectively. $\bar{R}=0.40$. |  |

Note: $\theta$ is the sum $\theta_{1}+\theta_{2}$ for each state. Source: Author's calculations.

Figure M.7: Impulse-Response Analysis for the Cyclical Component of the Unemployment Rate
$\rightarrow$ Nevada - Arizona $*$ California $\square$ Wyoming $\rightarrow$ North Dakota $\rightarrow$ West Virginia


Note: Response to a 0.8 percentage point increase in the aggregate output gap.


[^0]:    ${ }^{1}$ The views expressed in this paper are solely the responsibility of the author and should not be interpreted as reflecting the views of the Board of Governors of the Federal Reserve System.

[^1]:    ${ }^{1}$ In the case of California, the prior distribution of $\alpha_{i}$ is a truncated normal with mean and standard deviation equal to one.
    ${ }^{2} \mathrm{I}$ also explored assuming flat priors for $\sigma_{\eta_{i}^{y}}^{2}, \sigma_{\eta_{i}^{u}}^{2}$, and $\sigma_{\zeta_{i}}^{2}$. The results do not change materially.

[^2]:    ${ }^{3}$ The state-level unemployment rate trend and the idiosyncratic component of the state's cycle remain constant in this exercise.
    ${ }^{4}$ The full derivation appears in online Appendix L.

[^3]:    Note: $\mathrm{x}=100 \times \frac{\alpha_{i}^{2} \operatorname{var}\left(\Delta c_{t}\right)}{\operatorname{var}\left(\Delta c_{i t}\right)}$ is the proportion of the variance of the state's real GDP growth rate that is explained by the variance of its cycle.

