

ECON 452* -- NOTE 8

A General Regression Model with Dummy Variable Interactions**Starting Point: Model 5.4 in Standard Notation**

- The **population regression equation for Model 5.4** can be written in standard notation as:

$$Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 IN2_i + \beta_4 IN3_i + \beta_5 IN4_i + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i IN2_i + \delta_4 F_i IN3_i + \delta_5 F_i IN4_i + u_i \quad (5.4)$$

- The **population regression function for Model 5.4** is:

$$E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i) = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 IN2_i + \beta_4 IN3_i + \beta_5 IN4_i + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i IN2_i + \delta_4 F_i IN3_i + \delta_5 F_i IN4_i \quad (5.4')$$

- The **female population regression function for Model 5.4** is obtained by setting the female indicator $F_i = 1$ in (5.4'):

$$\begin{aligned} E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) &= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 IN2_i + \beta_4 IN3_i + \beta_5 IN4_i + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 IN2_i + \delta_4 IN3_i + \delta_5 IN4_i \\ &= (\beta_0 + \delta_0) + (\beta_1 + \delta_1) X_{i1} + (\beta_2 + \delta_2) X_{i2} + (\beta_3 + \delta_3) IN2_i + (\beta_4 + \delta_4) IN3_i + (\beta_5 + \delta_5) IN4_i \end{aligned} \quad (5.4f)$$

$$\begin{aligned}
& E(Y_i | X_{i1}, X_{i2}, X_{i3}, F_i, IN2_i, IN3_i, IN4_i) \\
&= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i3} + \beta_4 IN2_i + \beta_5 IN3_i + \beta_6 IN4_i + \delta_0 F_i \\
&\quad + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i3} + \delta_4 F_i IN2_i + \delta_5 F_i IN3_i + \delta_6 F_i IN4_i
\end{aligned} \tag{5.4'}$$

- The *male population regression function for Model 5.4* is obtained by setting the female indicator $F_i = 0$ in (5.4'):

$$E(Y_i | F_i = 0, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 IN2_i + \beta_4 IN3_i + \beta_5 IN4_i \tag{5.4m}$$

- The *female-male difference in conditional mean Y for Model 5.4* is obtained by subtracting the male regression function (5.4m) from the female regression function (5.4f):

$$\begin{aligned}
& E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\
&= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 IN2_i + \beta_4 IN3_i + \beta_5 IN4_i + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 IN2_i + \delta_4 IN3_i + \delta_5 IN4_i \\
&\quad - \beta_0 - \beta_1 X_{i1} - \beta_2 X_{i2} - \beta_3 IN2_i - \beta_4 IN3_i - \beta_5 IN4_i \\
&= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 IN2_i + \delta_4 IN3_i + \delta_5 IN4_i
\end{aligned}$$

Model 5.5: Quadratic terms in the continuous explanatory variables X_1 and X_2

Expand Model 5.4 to include quadratic terms in the two continuous explanatory variables X_1 and X_2 .

The **population regression equation** for Model 5.5 is:

$$Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i + u_i \quad (5.5)$$

Estimate by OLS the population regression equation for Model 5.5 using the following *Stata* command:

```
regress y x1 x2 x1sq x2sq x1x2 in2 in3 in4 f fx1 fx2 fx1sq fx2sq fx1x2 fin2
fin3 fin4
```

The **population regression function** for Model 5.5 is:

$$E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i) \\ = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i \quad (5.5')$$

$$E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i \quad (5.5')$$

- The **female population regression function for Model 5.5** is obtained by setting the female indicator $F_i = 1$ in (5.5'):

$$E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\ = (\beta_0 + \delta_0) + (\beta_1 + \delta_1) X_{i1} + (\beta_2 + \delta_2) X_{i2} + (\beta_3 + \delta_3) X_{i1}^2 + (\beta_4 + \delta_4) X_{i2}^2 + (\beta_5 + \delta_5) X_{i1} X_{i2} \\ + (\beta_6 + \delta_6) IN2_i + (\beta_7 + \delta_7) IN3_i + (\beta_8 + \delta_8) IN4_i \quad (5.5f)$$

- The **male population regression function for Model 5.5** is obtained by setting the female indicator $F_i = 0$ in (5.5'):

$$E(Y_i | F_i = 0, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i$$

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned}
& E(Y_i | F_i = 1, \mathbf{x}_i^T) - E(Y_i | F_i = 0, \mathbf{x}_i^T) \\
&= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
&\quad + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
&\quad - (\beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i) \\
&= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
&\quad + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
&\quad - \beta_0 - \beta_1 X_{i1} - \beta_2 X_{i2} - \beta_3 X_{i1}^2 - \beta_4 X_{i2}^2 - \beta_5 X_{i1} X_{i2} - \beta_6 IN2_i - \beta_7 IN3_i - \beta_8 IN4_i \\
&= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i
\end{aligned}$$

- **Table formats** for reporting coefficient estimates of **Model 5.5**:

Table 2: OLS Estimates of Model 5.5 on Pooled Sample of Females and Males

Regressor Name	Females		Males		Female-Male Differences	
	Coef. Estimate	t-ratio	Coef. Estimate	t-ratio	Coef. Estimate	t-ratio
Intercept						
X ₁						
X ₂						
X ₁ -sq						
X ₂ -sq						
X ₁ X ₂						
IN2						
IN3						
IN4						
No. of obs = RSS = R-squared = ANOVA F = p-value of F =						

Table 2: OLS Estimates of Model 5.5 on Pooled Sample of Females and Males

Regressor	Females		Males		Female-Male Differences	
	$\hat{\beta}_j + \hat{\delta}_j$	t-ratio	$\hat{\beta}_j$	t-ratio	$\hat{\delta}_j$	t-ratio
Intercept						
X ₁						
X ₂						
X ₁ -sq						
X ₂ -sq						
X ₁ X ₂						
IN2						
IN3						
IN4						
No. of obs = RSS = R-squared = ANOVA F = p-value of F =						

Table 5: Hypothesis Test Results for Model 5.5

#	Null Hypothesis H_0	Interpretation of H_0	q ^{1/}	p-value ^{2/}
1	$\beta_1 + \delta_1 = 0$ & $\beta_3 + \delta_3 = 0$ & $\beta_5 + \delta_5 = 0$	ME of X_1 is zero for females	3	0.0000
2	$\beta_3 + \delta_3 = 0$ & $\beta_5 + \delta_5 = 0$	ME of X_1 is constant for females	2	0.0274
3	$\beta_1 = 0$ & $\beta_3 = 0$ & $\beta_5 = 0$	ME of X_1 is zero for males	3	0.0014
4	$\beta_3 = 0$ & $\beta_5 = 0$	ME of X_1 is constant for males	2	0.0083
5	$\delta_1 = 0$ & $\delta_3 = 0$ & $\delta_5 = 0$	ME of X_1 equal for females & males	3	0.0038
6	$\delta_3 = 0$ & $\delta_5 = 0$	F-M difference in ME of X_1 a constant	2	0.1494
7	$\beta_2 + \delta_2 = 0$ & $\beta_4 + \delta_4$ & $\beta_5 + \delta_5 = 0$	ME of X_2 is zero for females	3	0.0000
8	$\beta_4 + \delta_4$ & $\beta_5 + \delta_5 = 0$	ME of X_2 is constant for females	2	0.0000
9	$\beta_2 = 0$ & $\beta_4 = 0$ & $\beta_5 = 0$	ME of X_2 is zero for males	3	0.0741
10	$\beta_4 = 0$ & $\beta_5 = 0$	ME of X_2 is constant for males	2	0.3185
11	$\delta_2 = 0$ & $\delta_4 = 0$ & $\delta_5 = 0$	ME of X_2 equal for females & males	3	0.0063
12	$\delta_4 = 0$ & $\delta_5 = 0$	F-M difference in ME of X_2 a constant	2	0.03119
13	$\beta_6 = 0$ & $\beta_7 = 0$ & $\beta_8 = 0$	No industry effects for males	3	0.0003
14	$\beta_6 + \delta_6 = 0$ & $\beta_7 + \delta_7$ & $\beta_8 + \delta_8 = 0$	No industry effects for females	3	0.0000
15	$\delta_6 = 0$ & $\delta_7 = 0$ & $\delta_8 = 0$	Industry effects equal, females & males	3	0.0083
16	$\delta_j = 0$ for all $j = 0, 1, \dots, 8$	F-M mean Y difference = 0	10	0.0000
17	$\delta_j = 0$ for all $j = 1, \dots, 8$	F-M mean Y difference is constant	9	0.0000

Notes: 1/. q denotes the number of coefficient restrictions specified by the null hypothesis H_0 . 2/. The p-values are two-tail p-values for the calculated sample value of the test statistic.

The Marginal Effect of X_1 in Model 5.5

$$E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i \quad (5.5')$$

- The marginal effect of X_1 in Model 5.5 is:

$$\frac{\partial Y_i}{\partial X_{i1}} = \frac{\partial E(Y_i | F_i, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 F_i + 2\delta_3 F_i X_{i1} + \delta_5 F_i X_{i2}$$

- The marginal effect of X_1 for *females* in Model 5.5 is obtained by setting $F_i = 1$:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=1} = \frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \\ = (\beta_1 + \delta_1) + 2(\beta_3 + \delta_3) X_{i1} + (\beta_5 + \delta_5) X_{i2}$$

- The marginal effect of X_1 for *males* in Model 5.5 is obtained by setting $F_i = 0$:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=0} = \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2}$$

- The *female-male difference in the marginal effect of X_1 in Model 5.5* is:

$$\begin{aligned} & \frac{\partial E(Y_i | F_i = 1, \mathbf{x}_i^T)}{\partial X_{i1}} - \frac{\partial E(Y_i | F_i = 0, \mathbf{x}_i^T)}{\partial X_{i1}} \\ &= \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} - \beta_1 - 2\beta_3 X_{i1} - \beta_5 X_{i2} \\ &= \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \end{aligned}$$

Hypothesis Tests Respecting the Marginal Effect of X_1 for *Males* in Model 5.5

- The **marginal effect of X_1 for *males* in Model 5.5** is:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=0} = \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2}$$

- ◆ **Test 1m**: Test the hypothesis that the **marginal effect of X_1 on Y for *males* is zero** for all values of X_1 and X_2 .
- Sufficient conditions for $\partial Y_i / \partial X_{i1} = 0$ for all i for males are $\beta_1 = 0$ and $\beta_3 = 0$ and $\beta_5 = 0$.
- The **null and alternative hypotheses** are:
 - $H_0: \beta_1 = 0$ and $\beta_3 = 0$ and $\beta_5 = 0$
 - $H_1: \beta_1 \neq 0$ and/or $\beta_3 \neq 0$ and/or $\beta_5 \neq 0$
- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test x1 x1sq x1x2
```

- The **marginal effect of X_1 for males in Model 5.5** is:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=0} = \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2}$$

- ♦ **Test 2m:** Test the hypothesis that the *marginal effect of X_1 on Y for males is constant* – i.e., is unrelated to the values of X_1 and X_2 .
- Sufficient conditions for $\partial Y_i / \partial X_{i1} = \beta_1$ (a constant) for all males are $\beta_3 = 0$ and $\beta_5 = 0$.

- The **null and alternative hypotheses** are:

$$H_0: \beta_3 = 0 \text{ and } \beta_5 = 0$$

$$H_1: \beta_3 \neq 0 \text{ and/or } \beta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test x1sq x1x2
```

- ◆ ***Test 3m:*** Test the hypothesis that the **marginal effect of X_1 on Y for *males* is unrelated to, or does not depend upon, X_1 .**

- The **marginal effect of X_1 for *males* in Model 5.5** is:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=0} = \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2}$$

- A sufficient condition for $\partial Y_i / \partial X_{i1}$ to be unrelated to X_{i1} for all males is $\beta_3 = 0$.

- The ***null and alternative hypotheses*** are:

$$H_0: \beta_3 = 0$$

$$H_1: \beta_3 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

test x1sq *or* **test x1sq = 0**

- Equivalently, compute a **two-tail t-test** of H_0 against H_1 using the following *Stata lincom* command:

lincom _b[x1sq]

- ◆ ***Test 4m:*** Test the hypothesis that the **marginal effect of X_1 on Y for *males* is unrelated to, or does not depend upon, X_2 .**

- The **marginal effect of X_1 for *males* in Model 5.5** is:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=0} = \frac{\partial E(Y_i | F_i = 0, \mathbf{x}_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2}$$

- A sufficient condition for $\partial Y_i / \partial X_{i1}$ to be unrelated to X_{i2} for all males is $\beta_5 = 0$.

- The ***null and alternative hypotheses*** for this proposition are:

$$H_0: \beta_5 = 0$$

$$H_1: \beta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

test x1x2 *or* **test x1x2 = 0**

- Equivalently, compute a **two-tail t-test** of H_0 against H_1 using the following *Stata lincom* command:

lincom _b[x1x2]

Hypothesis Tests Respecting the Marginal Effect of X_1 for *Females* in Model 5.5

- The **marginal effect of X_1 for females in Model 5.5** is:

$$\begin{aligned} \left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=1} &= \frac{\partial E(Y_i | F_i = 1, X_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \\ &= (\beta_1 + \delta_1) + 2(\beta_3 + \delta_3)X_{i1} + (\beta_5 + \delta_5)X_{i2} \end{aligned}$$

- ♦ ***Test 1f:*** Test the hypothesis that the *marginal effect of X_1 on Y for females is zero* for all values of X_1 and X_2 .

- Sufficient conditions for $\partial Y_i / \partial X_{i1} = 0$ for all females are $\beta_1 + \delta_1 = 0$ and $\beta_3 + \delta_3 = 0$ and $\beta_5 + \delta_5 = 0$.

- The ***null and alternative hypotheses*** are:

$$H_0: \beta_1 + \delta_1 = 0 \text{ and } \beta_3 + \delta_3 = 0 \text{ and } \beta_5 + \delta_5 = 0$$

$$H_1: \beta_1 + \delta_1 \neq 0 \text{ and/or } \beta_3 + \delta_3 \neq 0 \text{ and/or } \beta_5 + \delta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* commands:

```
test x1 + fx1 = 0, notest
test x1sq + fx1sq = 0, accumulate notest
test x1x2 + fx1x2 = 0, accumulate
```

- ◆ **Test 2f:** Test the hypothesis that the *marginal effect of X_1 on Y for females is constant* – i.e., is unrelated to the values of X_1 and X_2 .
- The **marginal effect of X_1 for females in Model 5.5** is:

$$\begin{aligned} \left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=1} &= \frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \\ &= (\beta_1 + \delta_1) + 2(\beta_3 + \delta_3)X_{i1} + (\beta_5 + \delta_5)X_{i2} \end{aligned}$$

- Sufficient conditions for $\partial Y_i / \partial X_{i1} = \beta_1 + \delta_1$ (a constant) for all females are $\beta_3 + \delta_3 = 0$ and $\beta_5 + \delta_5 = 0$.
- The **null and alternative hypotheses** are:

$$H_0: \beta_3 + \delta_3 = 0 \text{ and } \beta_5 + \delta_5 = 0$$

$$H_1: \beta_3 + \delta_3 \neq 0 \text{ and/or } \beta_5 + \delta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test x1sq + fx1sq = 0, notest
test x1x2 + fx1x2 = 0, accumulate
```

- ◆ ***Test 3f***: Test the hypothesis that the **marginal effect of X_1 on Y for females is unrelated to, or does not depend upon, X_1** .

- The **marginal effect of X_1 for females in Model 5.5** is:

$$\begin{aligned}\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=1} &= \frac{\partial E(Y_i | F_i = 1, X_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \\ &= (\beta_1 + \delta_1) + 2(\beta_3 + \delta_3)X_{i1} + (\beta_5 + \delta_5)X_{i2}\end{aligned}$$

- A sufficient condition for $\partial Y_i / \partial X_{i1}$ to be unrelated to X_{i1} for all females is $\beta_3 + \delta_3 = 0$.

- The ***null and alternative hypotheses*** are:

$$H_0: \beta_3 + \delta_3 = 0$$

$$H_1: \beta_3 + \delta_3 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test x1sq + fx1sq = 0
```

- Equivalently, compute a **two-tail t-test** of H_0 against H_1 using the following *Stata lincom* command:

```
lincom _b[x1sq] + _b[fx1sq]
```

- ◆ **Test 4f:** Test the hypothesis that the **marginal effect of X_1 on Y for females is unrelated to, or does not depend upon, X_2 .**
- The **marginal effect of X_1 for females in Model 5.5** is:

$$\begin{aligned}\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=1} &= \frac{\partial E(Y_i | F_i = 1, X_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \\ &= (\beta_1 + \delta_1) + 2(\beta_3 + \delta_3)X_{i1} + (\beta_5 + \delta_5)X_{i2}\end{aligned}$$

- A sufficient condition for $\partial Y_i / \partial X_{i1}$ to be unrelated to X_{i2} for all females is $\beta_5 + \delta_5 = 0$.
- The **null and alternative hypotheses** for this proposition are:

$$H_0: \beta_5 + \delta_5$$

$$H_1: \beta_5 + \delta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test x1x2 + fx1x2 = 0
```

- Equivalently, compute a **two-tail t-test** of H_0 against H_1 using the following *Stata lincom* command:

```
lincom _b[x1x2] + _b[fx1x2]
```

Hypothesis Tests for *Female-Male Differences* in the Marginal Effect of X_1 in Model 5.5

- The *female-male difference* in the marginal effect of X_1 in Model 5.5 is:

$$\frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} - \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2}$$

- ◆ **Test 5:** Test the hypothesis that the *marginal effect of X_1 on Y for females equals the marginal effect of X_1 on Y for males* for any values of X_1 and X_2 – i.e., the *female-male difference* in the *marginal effect of X_1 on Y is zero* for any values of X_1 and X_2 .
- Sufficient conditions for the **female-male difference** in the marginal effect of X_1 on Y to equal zero for all values of X_1 and X_2 are $\delta_1 = 0$ and $\delta_3 = 0$ and $\delta_5 = 0$.
- The **null and alternative hypotheses** are:
 - $H_0: \delta_1 = 0$ and $\delta_3 = 0$ and $\delta_5 = 0$
 - $H_1: \delta_1 \neq 0$ and/or $\delta_3 \neq 0$ and/or $\delta_5 \neq 0$
- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test fx1 fx1sq fx1x2
```

- **Remark:** The null hypothesis H_0 implies that the female-male difference in conditional mean Y is unrelated to the explanatory variable X_1 .

The *female-male difference in conditional mean Y* for Model 5.5 is:

$$\begin{aligned}
 & E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\
 &= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
 &= \delta_0 + \delta_2 X_{i2} + \delta_4 X_{i2}^2 + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \quad \text{under } H_0: \delta_1 = 0 \text{ and } \delta_3 = 0 \text{ and } \delta_5 = 0
 \end{aligned}$$

- ◆ **Test 6:** Test the hypothesis that the *female-male difference* in the *marginal effect of X_1 on Y* is a *constant* for any values of X_1 and X_2 .

- The *female-male difference* in the marginal effect of X_1 in Model 5.5 is:

$$\frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} - \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2}$$

- Sufficient conditions for the *female-male difference* in the marginal effect of X_1 on Y to equal the constant δ_1 for all values of X_1 and X_2 are $\delta_3 = 0$ and $\delta_5 = 0$.

- The *null and alternative hypotheses* are:

$$H_0: \delta_3 = 0 \text{ and } \delta_5 = 0$$

$$H_1: \delta_3 \neq 0 \text{ and/or } \delta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test fx1sq fx1x2
```

- **Remark:** The null hypothesis H_0 implies that the female-male difference in conditional mean Y is unrelated to the regressors X_{i1}^2 and $X_{i1}X_{i2}$.

The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned}
 & E(Y_i | F_i = 1, \mathbf{x}_i^T) - E(Y_i | F_i = 0, \mathbf{x}_i^T) \\
 &= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
 &= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_4 X_{i2}^2 + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \quad \text{under } H_0: \delta_3 = 0 \text{ and } \delta_5 = 0
 \end{aligned}$$

- ◆ **Test 7:** Test the hypothesis that the *female-male difference* in the *marginal effect* of X_1 on Y is **unrelated to, or does not depend upon, X_1** .

- The *female-male difference* in the marginal effect of X_1 in Model 5.5 is:

$$\frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} - \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2}$$

- A sufficient condition for the *female-male difference* in the marginal effect of X_1 on Y to be **unrelated to X_1** is $\delta_3 = 0$.

- The *null and alternative hypotheses* are:

$$H_0: \delta_3 = 0$$

$$H_1: \delta_3 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test fx1sq          or          test fx1sq = 0
```

- Equivalently, compute a **two-tail t-test** of H_0 against H_1 using the following *Stata lincom* command:

```
lincom _b[fx1sq]
```

- **Remark:** The null hypothesis H_0 implies that the female-male difference in conditional mean Y is unrelated to the regressor X_{i1}^2 .

The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned}
 & E(Y_i | F_i = 1, \mathbf{x}_i^T) - E(Y_i | F_i = 0, \mathbf{x}_i^T) \\
 &= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
 &= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \quad \text{under } H_0: \delta_3 = 0
 \end{aligned}$$

- ◆ **Test 8:** Test the hypothesis that the *female-male difference* in the *marginal effect* of X_1 on Y is **unrelated to, or does not depend upon, X_2** .

- The *female-male difference* in the marginal effect of X_1 in Model 5.5 is:

$$\frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} - \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2}$$

- A sufficient condition for the *female-male difference* in the marginal effect of X_1 on Y to be **unrelated to X_2** is $\delta_5 = 0$.

- The *null and alternative hypotheses* are:

$$H_0: \delta_5 = 0$$

$$H_1: \delta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

test fx1x2 *or* **test fx1x2 = 0**

- Equivalently, compute a **two-tail t-test** of H_0 against H_1 using the following *Stata lincom* command:

lincom _b[fx1x2]

- **Remark:** The null hypothesis H_0 implies that the female-male difference in conditional mean Y is unrelated to the regressor $X_{i1}X_{i2}$.

The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned}
 & E(Y_i | F_i = 1, \mathbf{x}_i^T) - E(Y_i | F_i = 0, \mathbf{x}_i^T) \\
 &= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
 &= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \quad \text{under } H_0: \delta_5 = 0
 \end{aligned}$$

Hypothesis Tests Respecting the Effects of Industry in Model 5.5

- ◆ ***Test 1-Industry:*** Test the hypothesis of **no industry effects for males**. This is equivalent to the hypothesis that conditional mean Y for males is unrelated to industry, i.e., that there **are no inter-industry differences in conditional mean Y for males**.

- The **male population regression function for Model 5.5** is:

$$E(Y_i | F_i = 0, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\ = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i$$

- Sufficient conditions for the conditional mean value of Y for males to be unrelated to industry are $\beta_6 = 0$ and $\beta_7 = 0$ and $\beta_8 = 0$.

- The **null and alternative hypotheses** are:

$$H_0: \beta_6 = 0 \text{ and } \beta_7 = 0 \text{ and } \beta_8 = 0$$

$$H_1: \beta_6 \neq 0 \text{ and/or } \beta_7 \neq 0 \text{ and/or } \beta_8 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test in2 in3 in4
```

- ♦ ***Test 2-Industry:*** Test the hypothesis of **no industry effects for females**. This is equivalent to the hypothesis that conditional mean Y for females is unrelated to industry, i.e., that there **are no inter-industry differences in conditional mean Y for females**.
- The **female population regression function for Model 5.5** is:

$$\begin{aligned}
 E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) & \\
 &= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
 &\quad + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
 &= (\beta_0 + \delta_0) + (\beta_1 + \delta_1) X_{i1} + (\beta_2 + \delta_2) X_{i2} + (\beta_3 + \delta_3) X_{i1}^2 + (\beta_4 + \delta_4) X_{i2}^2 + (\beta_5 + \delta_5) X_{i1} X_{i2} \\
 &\quad + (\beta_6 + \delta_6) IN2_i + (\beta_7 + \delta_7) IN3_i + (\beta_8 + \delta_8) IN4_i \tag{5.5f}
 \end{aligned}$$

- Sufficient conditions for the conditional mean value of Y for females to be unrelated to industry are $\beta_6 + \delta_6 = 0$ and $\beta_7 + \delta_7 = 0$ and $\beta_8 + \delta_8 = 0$.
- The **null and alternative hypotheses** are:
 - $H_0: \beta_6 + \delta_6 = 0$ and $\beta_7 + \delta_7 = 0$ and $\beta_8 + \delta_8 = 0$
 - $H_1: \beta_6 + \delta_6 \neq 0$ and/or $\beta_7 + \delta_7 \neq 0$ and/or $\beta_8 + \delta_8 \neq 0$
- Compute an **F-test** of H_0 against H_1 using the following *Stata test* commands:

```

test in2 + fin2 = 0, notest
test in3 + fin3 = 0, accumulate notest
test in4 + fin4 = 0, accumulate

```

- ◆ ***Test 3-Industry***: Test the hypothesis of **no female-male differences in industry effects** – i.e., that the **female-male difference** in conditional mean Y is **unrelated to industry**.

This is equivalent to the hypothesis that **industry effects are equal for females and males**, i.e., that **inter-industry differences in conditional mean Y for females equal inter-industry differences in conditional mean Y for males**.

- The **female-male difference in conditional mean Y for Model 5.5** is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

- Sufficient conditions for the female-male difference in conditional mean Y to be unrelated to industry (for equal industry effects for males and females) are $\delta_6 = 0$ and $\delta_7 = 0$ and $\delta_8 = 0$.

- The **null and alternative hypotheses** are:

$$H_0: \delta_6 = 0 \text{ and } \delta_7 = 0 \text{ and } \delta_8 = 0$$

$$H_1: \delta_6 \neq 0 \text{ and/or } \delta_7 \neq 0 \text{ and/or } \delta_8 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test fin2 fin3 fin4
```

Hypothesis Tests Respecting *Female-Male Differences* in Conditional Mean Y in Model 5.5

- The *female-male difference* in conditional mean Y for Model 5.5 is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

- ◆ **Test 1:** The *female-male difference* in conditional mean Y equals zero for all observations, i.e., for any given values of the explanatory variables X_1 , X_2 , X_3 , and industry.

- The *null and alternative hypotheses* are:

$$H_0: \delta_0 = 0 \text{ and } \delta_1 = 0 \text{ and } \delta_2 = 0 \text{ and } \delta_3 = 0 \text{ and } \delta_4 = 0 \text{ and } \delta_5 = 0 \text{ and } \delta_6 = 0 \text{ and } \delta_7 = 0 \text{ and } \delta_8 = 0$$

or

$$\delta_j = 0 \quad \text{for all } j = 0, 1, \dots, 8$$

$$H_1: \delta_0 \neq 0 \text{ and/or } \delta_1 \neq 0 \text{ and/or } \delta_2 \neq 0 \text{ and/or } \delta_3 \neq 0 \text{ and/or } \delta_4 \neq 0 \text{ and/or } \delta_5 \neq 0 \\ \text{and/or } \delta_6 \neq 0 \text{ and/or } \delta_7 \neq 0 \text{ and/or } \delta_8 \neq 0$$

or

$$\delta_j \neq 0 \quad j = 0, 1, \dots, 8$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test f fx1 fx2 fx1sq fx2sq fx1x2 fin2 fin3 fin4
```

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

- ♦ **Test 2:** The *female-male difference in conditional mean Y equals a constant*, i.e., it does not depend on the values of the explanatory variables X_1 , X_2 , X_3 , and industry.
- The *null and alternative hypotheses* are:

$$H_0: \delta_1 = 0 \text{ and } \delta_2 = 0 \text{ and } \delta_3 = 0 \text{ and } \delta_4 = 0 \text{ and } \delta_5 = 0 \text{ and } \delta_6 = 0 \text{ and } \delta_7 = 0 \text{ and } \delta_8 = 0$$

or

$$\delta_j = 0 \quad \text{for all } j = 1, 2, \dots, 8$$

$$H_1: \delta_1 \neq 0 \text{ and/or } \delta_2 \neq 0 \text{ and/or } \delta_3 \neq 0 \text{ and/or } \delta_4 \neq 0 \text{ and/or } \delta_5 \neq 0 \text{ and/or } \delta_6 \neq 0 \text{ and/or } \delta_7 \neq 0 \text{ and/or } \delta_8 \neq 0$$

or

$$\delta_j \neq 0 \quad j = 1, 2, \dots, 8$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test fx1 fx2 fx1sq fx2sq fx1x2 fin2 fin3 fin4
```

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

- ♦ ***Test 3:*** The *female-male difference* in conditional mean Y **does not depend on X_1** – i.e., the **marginal effect of X_1 is equal** for *males and females*.

- The *null and alternative hypotheses* are:

$$H_0: \delta_1 = 0 \text{ and } \delta_3 = 0 \text{ and } \delta_5 = 0$$

$$H_1: \delta_1 \neq 0 \text{ and/or } \delta_3 \neq 0 \text{ and/or } \delta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test fx1 fx1sq fx1x2
```

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

- ♦ ***Test 4:*** The *female-male difference* in conditional mean Y **does not depend on X_2** – i.e., the **marginal effect of X_2 is equal** for *males and females*.

- The *null and alternative hypotheses* are:

$$H_0: \delta_2 = 0 \text{ and } \delta_4 = 0 \text{ and } \delta_5 = 0$$

$$H_1: \delta_2 \neq 0 \text{ and/or } \delta_4 \neq 0 \text{ and/or } \delta_5 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test fx2 fx2sq fx1x2
```

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

- ◆ ***Test 5:*** The *female-male difference* in conditional mean Y **does not depend on industry** – i.e., **industry effects are equal for males and females.**

- The *null and alternative hypotheses* are:

$$H_0: \delta_6 = 0 \text{ and } \delta_7 = 0 \text{ and } \delta_8 = 0$$

$$H_1: \delta_6 \neq 0 \text{ and/or } \delta_7 \neq 0 \text{ and/or } \delta_8 \neq 0$$

- Compute an **F-test** of H_0 against H_1 using the following *Stata test* command:

```
test fin2 fin3 fin4
```

Evaluating the Marginal Effects of *Continuous* Explanatory Variables in Model 5.5

Review of Model 5.5: Quadratic terms in the continuous explanatory variables X_1 and X_2

The **population regression equation** for Model 5.5 is:

$$Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i + u_i \quad (5.5)$$

The **population regression function** for Model 5.5 is:

$$E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i) \\ = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i \quad (5.5')$$

Stata command for computing OLS estimates of the pooled, full-interaction regression equation (5.5):

```
regress y x1 x2 x1sq x2sq x1x2 in2 in3 in4 f fx1 fx2 fx1sq fx2sq fx1x2 fin2
fin3 fin4
```

The **population regression function** for Model 5.5 is:

$$\begin{aligned}
 E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i) \\
 = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
 + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i
 \end{aligned} \tag{5.5'}$$

- The **female population regression function** for Model 5.5 is obtained by setting the female indicator $F_i = 1$ in (5.5'):

$$\begin{aligned}
 E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\
 = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
 + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
 = (\beta_0 + \delta_0) + (\beta_1 + \delta_1) X_{i1} + (\beta_2 + \delta_2) X_{i2} + (\beta_3 + \delta_3) X_{i1}^2 + (\beta_4 + \delta_4) X_{i2}^2 + (\beta_5 + \delta_5) X_{i1} X_{i2} \\
 + (\beta_6 + \delta_6) IN2_i + (\beta_7 + \delta_7) IN3_i + (\beta_8 + \delta_8) IN4_i
 \end{aligned} \tag{5.5f}$$

- The **male population regression function** for Model 5.5 is obtained by setting the female indicator $F_i = 0$ in (5.5'):

$$\begin{aligned}
 E(Y_i | F_i = 0, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\
 = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i
 \end{aligned}$$

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} & E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ &= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

The Marginal Effect of X_1 in Model 5.5

$$E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i \quad (5.5')$$

- The **marginal effect of X_1 in Model 5.5** is:

$$\frac{\partial Y_i}{\partial X_{i1}} = \frac{\partial E(Y_i | F_i, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 F_i + 2\delta_3 F_i X_{i1} + \delta_5 F_i X_{i2}$$

- The **marginal effect of X_1 for females in Model 5.5** is obtained by setting $F_i = 1$:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=1} = \frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \\ = (\beta_1 + \delta_1) + 2(\beta_3 + \delta_3)X_{i1} + (\beta_5 + \delta_5)X_{i2}$$

- The **marginal effect of X_1 for males in Model 5.5** is obtained by setting $F_i = 0$:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=0} = \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2}$$

- The *female-male difference in the marginal effect of X_1 in Model 5.5* is:

$$\begin{aligned} \frac{\partial E(Y_i | F_i = 1, \mathbf{x}_i^T)}{\partial X_{i1}} - \frac{\partial E(Y_i | F_i = 0, \mathbf{x}_i^T)}{\partial X_{i1}} \\ &= \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} - \beta_1 - 2\beta_3 X_{i1} - \beta_5 X_{i2} \\ &= \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \end{aligned}$$

Evaluate the Marginal Effect of X_1 in Model 5.5 for Males and Females

- First, select specific values of X_1 and X_2 at which to evaluate the marginal effect of X_1 for *males* and *females* in Model 5.5. Common choices for typical values of X_1 and X_2 are:

(1) *sample mean* values of X_1 and X_2 , denoted as \bar{X}_1 and \bar{X}_2 ;

(2) *sample median* values of X_1 and X_2 , denoted as $X_{1,50p}$ and $X_{2,50p}$.

Stata commands for defining as scalars the *sample median* values of X_1 and X_2 :

```
summarize x1, detail
return list
scalar x1med = r(p50)

summarize x2, detail
return list
scalar x2med = r(p50)

scalar list x1med x2med
```

- The **marginal effect of X_1 for males in Model 5.5** is:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=0} = \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2}$$

Evaluated at the **sample median values of X_1 and X_2** , the **marginal effect of X_1 for males** in Model 5.5 is:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=0} = \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{1,50p} + \beta_5 X_{2,50p}$$

Stata command:

```
lincom _b[x1] + 2*_b[x1sq]*x1med + _b[x1x2]*x2med
```

- The **marginal effect of X_1 for females in Model 5.5** is:

$$\begin{aligned} \left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=1} &= \frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} = \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \\ &= (\beta_1 + \delta_1) + 2(\beta_3 + \delta_3)X_{i1} + (\beta_5 + \delta_5)X_{i2} \end{aligned}$$

Evaluated at the **sample median values of X_1 and X_2** , the **marginal effect of X_1 for females** in Model 5.5 is:

$$\left. \frac{\partial Y_i}{\partial X_{i1}} \right|_{F_i=1} = \frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} = (\beta_1 + \delta_1) + 2(\beta_3 + \delta_3)X_{1,50p} + (\beta_5 + \delta_5)X_{2,50p}$$

Stata command:

```
lincom _b[x1] + _b[fx1] + 2*(_b[x1sq] + _b[fx1sq])*x1med + (_b[x1x2] +
_b[fx1x2])*x2med
```

- The *female-male difference in the marginal effect of X_1 in Model 5.5* is:

$$\begin{aligned} \frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} - \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} \\ &= \beta_1 + 2\beta_3 X_{i1} + \beta_5 X_{i2} + \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} - \beta_1 - 2\beta_3 X_{i1} - \beta_5 X_{i2} \\ &= \delta_1 + 2\delta_3 X_{i1} + \delta_5 X_{i2} \end{aligned}$$

Evaluated at the *sample median values of X_1 and X_2* , the *female-male difference in the marginal effect of X_1 in Model 5.5* is:

$$\frac{\partial E(Y_i | F_i = 1, x_i^T)}{\partial X_{i1}} - \frac{\partial E(Y_i | F_i = 0, x_i^T)}{\partial X_{i1}} = \delta_1 + 2\delta_3 X_{1,50p} + \delta_5 X_{2,50p}$$

Stata command:

```
lincom _b[fx1] + 2*_b[fx1sq])*x1med + _b[fx1x2])*x2med
```

Evaluating the *Marginal Effects* of the *Categorical Explanatory Variable* in Model 5.5

General Nature: The marginal effects of a *categorical* explanatory variable such as industry consist of the differences in conditional mean values of Y between *pairs* of industry categories – e.g., the conditional mean Y difference between *males* in industries 4 and 2, and the conditional mean Y difference between *females* in industries 4 and 2.

Recall that the **population regression function** for Model 5.5 is:

$$\begin{aligned}
 E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i) \\
 = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
 + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i
 \end{aligned} \tag{5.5'}$$

- The **female population regression function** for Model 5.5 is obtained by setting the female indicator $F_i = 1$ in (5.5'):

$$\begin{aligned}
 E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\
 = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
 + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
 = (\beta_0 + \delta_0) + (\beta_1 + \delta_1) X_{i1} + (\beta_2 + \delta_2) X_{i2} + (\beta_3 + \delta_3) X_{i1}^2 + (\beta_4 + \delta_4) X_{i2}^2 + (\beta_5 + \delta_5) X_{i1} X_{i2} \\
 + (\beta_6 + \delta_6) IN2_i + (\beta_7 + \delta_7) IN3_i + (\beta_8 + \delta_8) IN4_i
 \end{aligned} \tag{5.5f}$$

$$E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i \quad (5.5')$$

- The **male population regression function for Model 5.5** is obtained by setting the female indicator $F_i = 0$ in (5.5'):

$$E(Y_i | F_i = 0, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i$$

- The **female-male difference in conditional mean Y for Model 5.5** is:

$$E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T)$$

$$= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i$$

Marginal Effects of Industry for Males in Model 5.5

- The *male* population regression function for Model 5.5 is obtained by setting the female indicator $F_i = 0$ in (5.5')

$$E(Y_i | F_i = 0, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\ = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i$$

The *Stata* command for computing OLS estimates of Model 5.5 is:

```
regress y x1 x2 x1sq x2sq x1x2 in2 in3 in4 f fx1 fx2 fx1sq fx2sq fx1x2 fin2
fin3 fin4
```

1. The **industry 2-industry 1 difference in conditional mean Y for males** equals β_6 in Model 5.5. To display an estimate of β_6 in Model 5.5, use the following *Stata* **limcom** command:

```
lincom _b[in2]
```

2. The **industry 3-industry 1 difference in conditional mean Y for males** equals β_7 in Model 5.5. To display an estimate of β_7 in Model 5.5, use the following *Stata* **limcom** command:

```
lincom _b[in3]
```

- The *male* population regression function for Model 5.5 is:

$$E(Y_i | F_i = 0, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\ = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i$$

3. The **industry 4-industry 1 difference in conditional mean Y for males** equals β_8 in Model 5.5. To display an estimate of β_8 in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[in4]
```

4. The **industry 3-industry 2 difference in conditional mean Y for males** equals $\beta_7 - \beta_6$ in Model 5.5. To compute an estimate of $\beta_7 - \beta_6$ in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[in3] - _b[in2]
```

5. The **industry 4-industry 2 difference in conditional mean Y for males** equals $\beta_8 - \beta_6$ in Model 5.5. To compute an estimate of $\beta_8 - \beta_6$ in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[in4] - _b[in2]
```

6. The **industry 4-industry 3 difference in conditional mean Y for males** equals $\beta_8 - \beta_7$ in Model 5.5. To compute an estimate of $\beta_8 - \beta_7$ in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[in4] - _b[in3]
```

Marginal Effects of Industry for Females in Model 5.5

- The *female* population regression function for Model 5.5 is:

$$\begin{aligned}
 E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\
 = (\beta_0 + \delta_0) + (\beta_1 + \delta_1)X_{i1} + (\beta_2 + \delta_2)X_{i2} + (\beta_3 + \delta_3)X_{i1}^2 + (\beta_4 + \delta_4)X_{i2}^2 + (\beta_5 + \delta_5)X_{i1}X_{i2} \\
 + (\beta_6 + \delta_6)IN2_i + (\beta_7 + \delta_7)IN3_i + (\beta_8 + \delta_8)IN4_i
 \end{aligned} \tag{5.5f}$$

Again, the *Stata* command for computing OLS estimates of regression equation (5.5) is:

```
regress y x1 x2 x1sq x2sq x1x2 in2 in3 in4 f fx1 fx2 fx1sq fx2sq fx1x2 fin2
fin3 fin4
```

- The **industry 2-industry 1 difference in conditional mean Y for females** equals $\beta_6 + \delta_6$ in Model 5.5. To compute an estimate of $(\beta_6 + \delta_6)$ in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[in2] + _b[fin2]
```

- The **industry 3-industry 1 difference in conditional mean Y for females** equals $\beta_7 + \delta_7$ in Model 5.5. To compute an estimate of $(\beta_7 + \delta_7)$ in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[in3] + _b[fin3]
```

- The *female* population regression function for Model 5.5 is:

$$\begin{aligned}
 E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\
 = (\beta_0 + \delta_0) + (\beta_1 + \delta_1)X_{i1} + (\beta_2 + \delta_2)X_{i2} + (\beta_3 + \delta_3)X_{i1}^2 + (\beta_4 + \delta_4)X_{i2}^2 + (\beta_5 + \delta_5)X_{i1}X_{i2} \\
 + (\beta_6 + \delta_6)IN2_i + (\beta_7 + \delta_7)IN3_i + (\beta_8 + \delta_8)IN4_i
 \end{aligned} \tag{5.5f}$$

3. The **industry 4-industry 1 difference in conditional mean Y for females** equals $\beta_8 + \delta_8$ in Model 5.5. To compute an estimate of $(\beta_8 + \delta_8)$ in Model 5.5, use the following *Stata* **limcom** command:

```
lincom _b[in4] + _b[fin4]
```

4. The **industry 3-industry 2 difference in conditional mean Y for females** equals $(\beta_7 + \delta_7) - (\beta_6 + \delta_6)$ in Model 5.5. To compute an estimate of $(\beta_7 + \delta_7) - (\beta_6 + \delta_6)$ in Model 5.5, use the following *Stata* **limcom** command:

```
lincom _b[in3] + _b[fin3] - (_b[in2] + _b[fin2])
```

5. The **industry 4-industry 2 difference in conditional mean Y for females** equals $(\beta_8 + \delta_8) - (\beta_6 + \delta_6)$ in Model 5.5. To compute an estimate of $(\beta_8 + \delta_8) - (\beta_6 + \delta_6)$ in Model 5.5, use the following *Stata* **limcom** command:

```
lincom _b[in4] + _b[fin4] - (_b[in2] + _b[fin2])
```

- The *female* population regression function for Model 5.5 is:

$$\begin{aligned}
 E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\
 = (\beta_0 + \delta_0) + (\beta_1 + \delta_1)X_{i1} + (\beta_2 + \delta_2)X_{i2} + (\beta_3 + \delta_3)X_{i1}^2 + (\beta_4 + \delta_4)X_{i2}^2 + (\beta_5 + \delta_5)X_{i1}X_{i2} \\
 + (\beta_6 + \delta_6)IN2_i + (\beta_7 + \delta_7)IN3_i + (\beta_8 + \delta_8)IN4_i
 \end{aligned} \tag{5.5f}$$

6. The **industry 4-industry 3 difference in conditional mean Y for females** equals $(\beta_8 + \delta_8) - (\beta_7 + \delta_7)$ in Model 5.5. To compute an estimate of $(\beta_8 + \delta_8) - (\beta_7 + \delta_7)$ in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[in4] + _b[fin4] - (_b[in3] + _b[fin3])
```

Female-Male Differences in the *Marginal* Effects of Industry in Model 5.5

- The *female-male difference* in conditional mean Y for Model 5.5 is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

Again, the *Stata command* for computing OLS estimates of regression equation (5.5) is:

```
regress y x1 x2 x1sq x2sq x1x2 in2 in3 in4 f fx1 fx2 fx1sq fx2sq fx1x2 fin2
fin3 fin4
```

- The **industry 2-industry 1 difference** in conditional mean Y for *females* equals $\beta_6 + \delta_6$ in Model 5.5. The **industry 2-industry 1 difference** in conditional mean Y for *males* equals β_6 .

The **female-male difference** in the **industry 2-industry 1 difference** in conditional mean Y

$$= \beta_6 + \delta_6 - \beta_6 = \delta_6$$

To display an estimate of δ_6 in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[fin2]
```

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

2. The **industry 3-industry 1 difference in conditional mean Y for females** equals $\beta_7 + \delta_7$ in Model 5.5. The **industry 3-industry 1 difference in conditional mean Y for males** equals β_7 .

The **female-male difference in the industry 3-industry 1 difference in conditional mean Y**

$$= \beta_7 + \delta_7 - \beta_7 = \delta_7$$

To display an estimate of δ_7 in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[fin3]
```

3. The **industry 4-industry 1 difference in conditional mean Y for females** equals $\beta_8 + \delta_8$ in Model 5.5. The **industry 4-industry 1 difference in conditional mean Y for males** equals β_8 .

The **female-male difference in the industry 4-industry 1 difference in conditional mean Y**

$$= \beta_8 + \delta_8 - \beta_8 = \delta_8$$

To display an estimate of δ_8 in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[fin4]
```

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

4. The **industry 3-industry 2 difference in conditional mean Y for females** equals $(\beta_7 + \delta_7) - (\beta_6 + \delta_6)$ in Model 5.5. The **industry 3-industry 2 difference in conditional mean Y for males** equals $\beta_7 - \beta_6$.

The **female-male difference in the industry 3-industry 2 difference in conditional mean Y**

$$\begin{aligned} &= (\beta_7 + \delta_7) - (\beta_6 + \delta_6) - (\beta_7 - \beta_6) \\ &= \beta_7 + \delta_7 - \beta_6 - \delta_6 - (\beta_7 - \beta_6) \\ &= (\beta_7 - \beta_6) + (\delta_7 - \delta_6) - (\beta_7 - \beta_6) \\ &= (\delta_7 - \delta_6) \end{aligned}$$

To compute an estimate of $(\delta_7 - \delta_6)$ in Model 5.5, use the following *Stata* **limcom** command:

```
lincom _b[fin3] - _b[fin2]
```

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

5. The **industry 4-industry 2 difference in conditional mean Y for females** equals $(\beta_8 + \delta_8) - (\beta_6 + \delta_6)$ in Model 5.5. The **industry 4-industry 2 difference in conditional mean Y for males** equals $\beta_8 - \beta_6$.

The **female-male difference in the industry 4-industry 2 difference in conditional mean Y**

$$\begin{aligned} &= (\beta_8 + \delta_8) - (\beta_6 + \delta_6) - (\beta_8 - \beta_6) \\ &= \beta_8 + \delta_8 - \beta_6 - \delta_6 - (\beta_8 - \beta_6) \\ &= (\beta_8 - \beta_6) + (\delta_8 - \delta_6) - (\beta_8 - \beta_6) \\ &= (\delta_8 - \delta_6) \end{aligned}$$

To compute an estimate of $(\delta_8 - \delta_6)$ in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[fin4] - _b[fin2]
```

- The *female-male difference in conditional mean Y for Model 5.5* is:

$$\begin{aligned} E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T) \\ = \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \end{aligned}$$

6. The **industry 4-industry 3 difference in conditional mean Y for females** equals $(\beta_8 + \delta_8) - (\beta_7 + \delta_7)$ in Model 5.5. The **industry 4-industry 3 difference in conditional mean Y for males** equals $\beta_8 - \beta_7$.

The **female-male difference in the industry 4-industry 2 difference in conditional mean Y**

$$\begin{aligned} &= (\beta_8 + \delta_8) - (\beta_7 + \delta_7) - (\beta_8 - \beta_7) \\ &= \beta_8 + \delta_8 - \beta_7 - \delta_7 - (\beta_8 - \beta_7) \\ &= (\beta_8 - \beta_7) + (\delta_8 - \delta_7) - (\beta_8 - \beta_7) \\ &= (\delta_8 - \delta_7) \end{aligned}$$

To compute an estimate of $(\delta_8 - \delta_7)$ in Model 5.5, use the following *Stata* **lincom** command:

```
lincom _b[fin4] - _b[fin3]
```

Evaluating the *Conditional* Effects of the *Categorical* Explanatory Variable in Model 5.5

General Nature: The conditional effects of a *categorical* explanatory variable such as industry consist of the **conditional mean value of Y for each of the four industry categories** – industry 1, industry 2, industry 3, and industry 4 – for pre-selected values of the other explanatory variables in the model, such as the continuous explanatory variables X_1 and X_2 in Model 5.5.

Recall that the **population regression function for Model 5.5** is:

$$\begin{aligned}
 E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i) \\
 = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
 + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i
 \end{aligned} \tag{5.5'}$$

- The **female population regression function for Model 5.5** is obtained by setting the female indicator $F_i = 1$ in (5.5'):

$$\begin{aligned}
 E(Y_i | F_i = 1, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i) \\
 = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\
 + \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i \\
 = (\beta_0 + \delta_0) + (\beta_1 + \delta_1) X_{i1} + (\beta_2 + \delta_2) X_{i2} + (\beta_3 + \delta_3) X_{i1}^2 + (\beta_4 + \delta_4) X_{i2}^2 + (\beta_5 + \delta_5) X_{i1} X_{i2} \\
 + (\beta_6 + \delta_6) IN2_i + (\beta_7 + \delta_7) IN3_i + (\beta_8 + \delta_8) IN4_i
 \end{aligned} \tag{5.5f}$$

$$E(Y_i | X_{i1}, X_{i2}, F_i, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i \quad (5.5')$$

- The **male population regression function for Model 5.5** is obtained by setting the female indicator $F_i = 0$ in (5.5'):

$$E(Y_i | F_i = 0, X_{i1}, X_{i2}, IN2_i, IN3_i, IN4_i)$$

$$= \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i$$

- The **female-male difference in conditional mean Y for Model 5.5** is:

$$E(Y_i | F_i = 1, x_i^T) - E(Y_i | F_i = 0, x_i^T)$$

$$= \delta_0 + \delta_1 X_{i1} + \delta_2 X_{i2} + \delta_3 X_{i1}^2 + \delta_4 X_{i2}^2 + \delta_5 X_{i1} X_{i2} + \delta_6 IN2_i + \delta_7 IN3_i + \delta_8 IN4_i$$

Objective: To compute estimates of the *conditional effect of industry* for **females and males**, and of the **female-male difference** in the conditional effect of industry **for pre-selected values of all the other explanatory variables** in the model.

- First, **select the values of X_1 and X_2** at which to evaluate the **conditional effect of industry for females and males in Model 5.5**.

Suppose we select the *pooled sample means* of the continuous explanatory variables X_1 and X_2 , which are denoted as \bar{X}_1 and \bar{X}_2 .

Stata commands:

```
summarize x1, detail
return list
scalar x1bar = r(mean)

summarize x2, detail
return list
scalar x2bar = r(mean)

scalar list x1bar x2bar
```

- Second, estimate by OLS the pooled, full-interaction regression equation (5.5):

$$Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \beta_3 X_{i1}^2 + \beta_4 X_{i2}^2 + \beta_5 X_{i1} X_{i2} + \beta_6 IN2_i + \beta_7 IN3_i + \beta_8 IN4_i \\ + \delta_0 F_i + \delta_1 F_i X_{i1} + \delta_2 F_i X_{i2} + \delta_3 F_i X_{i1}^2 + \delta_4 F_i X_{i2}^2 + \delta_5 F_i X_{i1} X_{i2} + \delta_6 F_i IN2_i + \delta_7 F_i IN3_i + \delta_8 F_i IN4_i + u_i \quad (5.5)$$

The *Stata* command for computing OLS estimates of regression equation (5.5) is:

```
regress y x1 x2 x1sq x2sq x1x2 in2 in3 in4 f fx1 fx2 fx1sq fx2sq fx1x2 fin2
fin3 fin4
```

- For *females*, evaluate the *female sample regression function* for each of the **four industry categories** at the pre-selected values of the other explanatory variables X_1 and X_2 , i.e., at the pooled sample means \bar{X}_1 and \bar{X}_2 .

$$\hat{E}(Y_i | F_i = 1, \bar{X}_1, \bar{X}_2, IN2_i, IN3_i, IN4_i)$$

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2 \\ + (\hat{\beta}_6 + \hat{\delta}_6)IN2_i + (\hat{\beta}_7 + \hat{\delta}_7)IN3_i + (\hat{\beta}_8 + \hat{\delta}_8)IN4_i$$

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2$$

for industry 1 ($IN2_i = 0, IN3_i = 0, IN4_i = 0$)

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2 + (\hat{\beta}_6 + \hat{\delta}_6)$$

for industry 2 ($IN2_i = 1, IN3_i = 0, IN4_i = 0$)

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2 + (\hat{\beta}_7 + \hat{\delta}_7)$$

for industry 3 ($IN2_i = 0, IN3_i = 1, IN4_i = 0$)

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2 + (\hat{\beta}_8 + \hat{\delta}_8)$$

for industry 4 ($IN2_i = 0, IN3_i = 0, IN4_i = 1$)

Stata commands for females:

To compute $\hat{E}(Y_i | F_i = 1, \bar{X}_1, \bar{X}_2, IN2_i = 0, IN3_i = 0, IN4_i = 0)$

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2$$

for industry 1 ($IN2_i = 0, IN3_i = 0, IN4_i = 0$)

```
lincom _b[_cons] + _b[f] + (_b[x1] + _b[fx1])*x1bar + (_b[x2]
+_b[fx2])*x2bar + (_b[x1sq] + _b[fx1sq])*x1bar*x1bar + (_b[x2sq] +
_b[fx2sq])*x2bar*x2bar + (_b[x1x2] + _b[fx1x2])*x1bar*x2bar
```

To compute $\hat{E}(Y_i | F_i = 1, \bar{X}_1, \bar{X}_2, IN2_i = 1, IN3_i = 0, IN4_i = 0)$

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2 + (\hat{\beta}_6 + \hat{\delta}_6)$$

for industry 2 ($IN2_i = 1, IN3_i = 0, IN4_i = 0$)

```
lincom _b[_cons] + _b[f] + (_b[x1] + _b[fx1])*x1bar + (_b[x2] +
_b[fx2])*x2bar + (_b[x1sq] + _b[fx1sq])*x1bar*x1bar + (_b[x2sq] +
_b[fx2sq])*x2bar*x2bar + (_b[x1x2] + _b[fx1x2])*x1bar*x2bar + (_b[in2] +
_b[fin2])
```

To compute $\hat{E}(Y_i | F_i = 1, \bar{X}_1, \bar{X}_2, IN2_i = 0, IN3_i = 1, IN4_i = 0)$

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2 + (\hat{\beta}_7 + \hat{\delta}_7)$$

for industry 3 ($IN2_i = 0, IN3_i = 1, IN4_i = 0$)

```
lincom _b[_cons] + _b[f] + (_b[x1] + _b[fx1])*x1bar + (_b[x2] +
_b[fx2])*x2bar + (_b[x1sq] + _b[fx1sq])*x1bar*x1bar + (_b[x2sq] +
_b[fx2sq])*x2bar*x2bar + (_b[x1x2] + _b[fx1x2])*x1bar*x2bar + (_b[in3] +
_b[fin3])
```

To compute $\hat{E}(Y_i | F_i = 1, \bar{X}_1, \bar{X}_2, IN2_i = 0, IN3_i = 0, IN4_i = 1)$

$$= (\hat{\beta}_0 + \hat{\delta}_0) + (\hat{\beta}_1 + \hat{\delta}_1)\bar{X}_1 + (\hat{\beta}_2 + \hat{\delta}_2)\bar{X}_2 + (\hat{\beta}_3 + \hat{\delta}_3)\bar{X}_1^2 + (\hat{\beta}_4 + \hat{\delta}_4)\bar{X}_2^2 + (\hat{\beta}_5 + \hat{\delta}_5)\bar{X}_1\bar{X}_2 + (\hat{\beta}_8 + \hat{\delta}_8)$$

for industry 4 ($IN2_i = 0, IN3_i = 0, IN4_i = 1$)

```
lincom _b[_cons] + _b[f] + (_b[x1] + _b[fx1])*x1bar + (_b[x2] +
_b[fx2])*x2bar + (_b[x1sq] + _b[fx1sq])*x1bar*x1bar + (_b[x2sq] +
_b[fx2sq])*x2bar*x2bar + (_b[x1x2] + _b[fx1x2])*x1bar*x2bar + (_b[in4] +
_b[fin4])
```

- For *males*, evaluate the *male sample regression function* for each of the **four industry categories** at the pre-selected values of the other explanatory variables X_1 and X_2 , i.e., at the pooled sample means \bar{X}_1 and \bar{X}_2 .

$$\hat{E}(Y_i | F_i = 0, \bar{X}_1, \bar{X}_2, IN2_i, IN3_i, IN4_i)$$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 + \hat{\beta}_6 IN2_i + \hat{\beta}_7 IN3_i + \hat{\beta}_8 IN4_i$$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 \quad \text{for industry 1 (IN2}_i = 0, \text{IN3}_i = 0, \text{IN4}_i = 0)$$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 + \hat{\beta}_6 \quad \text{for industry 2 (IN2}_i = 1, \text{IN3}_i = 0, \text{IN4}_i = 0)$$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 + \hat{\beta}_7 \quad \text{for industry 3 (IN2}_i = 0, \text{IN3}_i = 1, \text{IN4}_i = 0)$$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 + \hat{\beta}_8 \quad \text{for industry 4 (IN2}_i = 0, \text{IN3}_i = 0, \text{IN4}_i = 1)$$

Stata commands for males:

To compute $\hat{E}(Y_i | F_i = 0, \bar{X}_1, \bar{X}_2, IN2_i = 0, IN3_i = 0, IN4_i = 0)$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 \quad \text{for industry 1 (IN2}_i = 0, \text{IN3}_i = 0, \text{IN4}_i = 0)$$

```
lincom _b[_cons] + _b[x1]*x1bar + _b[x2]*x2bar + _b[x1sq]*x1bar*x1bar +
_b[x2sq]*x2bar*x2bar + _b[x1x2]*x1bar*x2bar
```

To compute $\hat{E}(Y_i | F_i = 0, \bar{X}_1, \bar{X}_2, IN2_i = 1, IN3_i = 0, IN4_i = 0)$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 + \hat{\beta}_6 \quad \text{for industry 2 (IN2}_i = 1, IN3_i = 0, IN4_i = 0)$$

```
lincom _b[_cons] + _b[x1]*x1bar + _b[x2]*x2bar + _b[x1sq]*x1bar*x1bar +
_b[x2sq]*x2bar*x2bar + _b[x1x2]*x1bar*x2bar + _b[in2]
```

To compute $\hat{E}(Y_i | F_i = 0, \bar{X}_1, \bar{X}_2, IN2_i = 0, IN3_i = 1, IN4_i = 0)$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 + \hat{\beta}_7 \quad \text{for industry 3 (IN2}_i = 0, IN3_i = 1, IN4_i = 0)$$

```
lincom _b[_cons] + _b[x1]*x1bar + _b[x2]*x2bar + _b[x1sq]*x1bar*x1bar +
_b[x2sq]*x2bar*x2bar + _b[x1x2]*x1bar*x2bar + _b[in3]
```

To compute $\hat{E}(Y_i | F_i = 0, \bar{X}_1, \bar{X}_2, IN2_i = 0, IN3_i = 0, IN4_i = 1)$

$$= \hat{\beta}_0 + \hat{\beta}_1 \bar{X}_1 + \hat{\beta}_2 \bar{X}_2 + \hat{\beta}_3 \bar{X}_1^2 + \hat{\beta}_4 \bar{X}_2^2 + \hat{\beta}_5 \bar{X}_1 \bar{X}_2 + \hat{\beta}_8 \quad \text{for industry 4 (IN2}_i = 0, IN3_i = 0, IN4_i = 1)$$

```
lincom _b[_cons] + _b[x1]*x1bar + _b[x2]*x2bar + _b[x1sq]*x1bar*x1bar +
_b[x2sq]*x2bar*x2bar + _b[x1x2]*x1bar*x2bar + _b[in4]
```

- **For the female-male differences in the conditional effect of industry**, evaluate the **female-male difference in sample regression functions** for each of the four industry categories at the pre-selected values of the other explanatory variables X_1 and X_2 , i.e., at the pooled sample means \bar{X}_1 and \bar{X}_2 .

$$\begin{aligned} & \hat{E}(Y_i | F_i = 1, \bar{x}^T, IN2_i, IN3_i, IN4_i) - \hat{E}(Y_i | F_i = 0, \bar{x}^T, IN2_i, IN3_i, IN4_i) \\ &= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 + \hat{\delta}_6 IN2_i + \hat{\delta}_7 IN3_i + \hat{\delta}_8 IN4_i \\ &= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 && \text{for industry 1 (IN2}_i = 0, \text{IN3}_i = 0, \text{IN4}_i = 0) \\ &= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 + \hat{\delta}_6 && \text{for industry 2 (IN2}_i = 1, \text{IN3}_i = 0, \text{IN4}_i = 0) \\ &= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 + \hat{\delta}_7 && \text{for industry 3 (IN2}_i = 0, \text{IN3}_i = 1, \text{IN4}_i = 0) \\ &= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 + \hat{\delta}_8 && \text{for industry 4 (IN2}_i = 0, \text{IN3}_i = 0, \text{IN4}_i = 1) \end{aligned}$$

Stata commands for female-male differences:

To compute $\hat{E}(Y_i | F_i = 1, \bar{x}^T, IN1_i = 1) - \hat{E}(Y_i | F_i = 0, \bar{x}^T, IN1_i = 1)$

$$= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 \quad \text{for industry 1 (IN2}_i = 0, \text{IN3}_i = 0, \text{IN4}_i = 0)$$

```
lincom _b[f] + _b[fx1]*x1bar + _b[fx2]*x2bar + _b[fx1sq]*x1bar*x1bar +
_b[fx2sq]*x2bar*x2bar + _b[fx1x2]*x1bar*x2bar
```

To compute $\hat{E}(Y_i | F_i = 1, \bar{x}^T, IN2_i = 1) - \hat{E}(Y_i | F_i = 0, \bar{x}^T, IN2_i = 1)$

$$= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 + \hat{\delta}_6 \quad \text{for industry 2 (IN2}_i = 1, \text{IN3}_i = 0, \text{IN4}_i = 0)$$

```
lincom _b[f] + _b[fx1]*x1bar + _b[fx2]*x2bar + _b[fx1sq]*x1bar*x1bar +
_b[fx2sq]*x2bar*x2bar + _b[fx1x2]*x1bar*x2bar + _b[fin2]
```

To compute $\hat{E}(Y_i | F_i = 1, \bar{x}^T, IN3_i = 1) - \hat{E}(Y_i | F_i = 0, \bar{x}^T, IN3_i = 1)$

$$= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 + \hat{\delta}_7 \quad \text{for industry 3 (IN2}_i = 0, \text{IN3}_i = 1, \text{IN4}_i = 0)$$

```
lincom _b[f] + _b[fx1]*x1bar + _b[fx2]*x2bar + _b[fx1sq]*x1bar*x1bar +
_b[fx2sq]*x2bar*x2bar + _b[fx1x2]*x1bar*x2bar + _b[fin3]
```

To compute $\hat{E}(Y_i | F_i = 1, \bar{x}^T, IN4_i = 1) - \hat{E}(Y_i | F_i = 0, \bar{x}^T, IN4_i = 1)$

$$= \hat{\delta}_0 + \hat{\delta}_1 \bar{X}_1 + \hat{\delta}_2 \bar{X}_2 + \hat{\delta}_3 \bar{X}_1^2 + \hat{\delta}_4 \bar{X}_2^2 + \hat{\delta}_5 \bar{X}_1 \bar{X}_2 + \hat{\delta}_8 \quad \text{for industry 4 (IN2}_i = 0, \text{IN3}_i = 0, \text{IN4}_i = 1)$$

```
lincom _b[f] + _b[fx1]*x1bar + _b[fx2]*x2bar + _b[fx1sq]*x1bar*x1bar +
_b[fx2sq]*x2bar*x2bar + _b[fx1x2]*x1bar*x2bar + _b[fin4]
```