

Solution to Exercise 9.5

★9.5 Consider the linear regression model with serially correlated errors,

$$y_t = \beta_1 + \beta_2 x_t + u_t, \quad u_t = \rho u_{t-1} + \varepsilon_t, \quad (9.117)$$

where the ε_t are IID, and the autoregressive parameter ρ is assumed either to be known or to be estimated consistently. The explanatory variable x_t is assumed to be contemporaneously correlated with ε_t (see Section 8.4 for the definition of contemporaneous correlation).

Recall from Chapter 7 that the covariance matrix Ω of the vector \mathbf{u} with typical element u_t is given by (7.32), and that Ω^{-1} can be expressed as $\Psi\Psi^\top$, where Ψ is defined in (7.60). Express the model (9.117) in the form (9.20), without taking account of the first observation.

Let Ω_t be the information set for observation t with $E(\varepsilon_t | \Omega_t) = 0$. Suppose that there exists a matrix \mathbf{Z} of instrumental variables, with $\mathbf{Z}_t \in \Omega_t$, such that the explanatory vector \mathbf{x} with typical element x_t is related to the instruments by the equation

$$\mathbf{x} = \mathbf{Z}\boldsymbol{\pi} + \mathbf{v}, \quad (9.118)$$

where $E(v_t | \Omega_t) = 0$. Derive the explicit form of the expression $(\Psi^\top \bar{\mathbf{X}})_t$ defined implicitly by equation (9.24) for the model (9.117). Find a matrix \mathbf{W} of instruments that satisfy the predeterminedness condition in the form (9.30) and that lead to asymptotically efficient estimates of the parameters β_1 and β_2 computed on the basis of the theoretical moment conditions (9.31) with your choice of \mathbf{W} .

As usual, we denote by \mathbf{y} the n -vector with typical element y_t . Using expression (7.60) for Ψ , we see that $(\Psi^\top \mathbf{y})_t = y_t - \rho y_{t-1}$ for $t = 2, \dots, n$. Thus, ignoring the first observation, we can write the model (9.117) in the form of equation (9.20) as follows:

$$y_t - \rho y_{t-1} = (1 - \rho)\beta_1 + \beta_2(x_t - \rho x_{t-1}) + \varepsilon_t, \quad (\text{S9.06})$$

since $(\Psi^\top \mathbf{u})_t = u_t - \rho u_{t-1} = \varepsilon_t$ by the specification (9.117).

According to equation (9.24), $E((\Psi^\top \mathbf{X})_t | \Omega_t) = (\Psi^\top \bar{\mathbf{X}})_t$. In what follows, we ignore the constant in (S9.06), since it is clear that the instruments \mathbf{W} must in any case include a constant. Thus we restrict our attention to $E((\Psi^\top \mathbf{x})_t | \Omega_t)$, which, by (9.118), is

$$\begin{aligned} E((\Psi^\top \mathbf{x})_t | \Omega_t) &= E(x_t - \rho x_{t-1} | \Omega_t) \\ &= E(\mathbf{Z}_t \boldsymbol{\pi} + v_t - \rho x_{t-1} | \Omega_t) \\ &= \mathbf{Z}_t \boldsymbol{\pi} - \rho x_{t-1}. \end{aligned}$$

The last equality uses the facts that $E(v_t | \Omega_t) = 0$ by assumption and that $x_{t-1} \in \Omega_t$, since x_{t-1} is predetermined for observation t . It follows that $(\Psi^\top \bar{x})_t = \mathbf{Z}_t \boldsymbol{\pi} - \rho x_{t-1}$, and so $\Psi^\top \bar{x} = \mathbf{Z} \boldsymbol{\pi} - \rho \mathbf{x}_{-1}$, where $(\mathbf{x}_{-1})_t = x_{t-1}$.

We now see that $\Psi^\top \bar{x} \in \mathcal{S}(\mathbf{Z}, \mathbf{x}_{-1})$. Since $\mathbf{Z}_t \in \Omega_t$ and $\mathbf{x}_{-1} \in \Omega_t$, the instruments \mathbf{W} , except for the constant, may be taken to be such that $\Psi^\top \mathbf{W} = [\mathbf{Z} \ \mathbf{x}_{-1}]$. Formally, $\mathbf{W} = (\Psi^\top)^{-1}[\mathbf{Z} \ \mathbf{x}_{-1}]$.

The moment conditions (9.31) are such that it is not necessary to use the inverse of Ψ^\top explicitly, since \mathbf{W} enters only through the transpose of $\Psi^\top \mathbf{W}$. Using (S9.06) in order to get an explicit expression for $\Psi^\top(\mathbf{y} - \mathbf{X}\boldsymbol{\beta})$ in (9.31), we find that the efficient estimates are obtained by using the theoretical moment conditions that $\varepsilon_t = y_t - \rho y_{t-1} - (1 - \rho)\beta_1 - \beta_2 x_t + \rho \beta_2 x_{t-1}$ should be orthogonal to a constant term, \mathbf{Z}_t , and x_{t-1} . Thus if $\mathbf{A} \equiv [\iota \ \mathbf{Z} \ \mathbf{x}_{-1}]$, the empirical moments are

$$\mathbf{A}^\top (\mathbf{y} - \rho \mathbf{y}_{-1} - (1 - \rho)\beta_1 - \beta_2 \mathbf{x} + \rho \beta_2 \mathbf{x}_{-1}),$$

and the optimal weighting matrix is proportional to $(\mathbf{A}^\top \mathbf{A})^{-1}$.