

Solution to Exercise 8.5

***8.5** Under the usual assumptions of this chapter, including (8.16), show that the plim of

$$\frac{1}{n}Q(\beta_0, \mathbf{y}) = \frac{1}{n}(\mathbf{y} - \mathbf{X}\beta_0)^\top \mathbf{P}_W(\mathbf{y} - \mathbf{X}\beta_0)$$

is zero if $\mathbf{y} = \mathbf{X}\beta_0 + \mathbf{u}$. Under the same assumptions, along with the asymptotic identification condition that $\mathbf{S}_{\mathbf{X}^\top \mathbf{W}}(\mathbf{S}_{\mathbf{W}^\top \mathbf{W}})^{-1}\mathbf{S}_{\mathbf{W}^\top \mathbf{X}}$ has full rank, show further that $\text{plim } n^{-1}Q(\beta, \mathbf{y})$ is strictly positive for $\beta \neq \beta_0$.

When $\mathbf{y} = \mathbf{X}\beta_0 + \mathbf{u}$, the probability limit of $n^{-1}Q(\beta, \mathbf{y})$ is

$$\begin{aligned} \text{plim}_{n \rightarrow \infty} \frac{1}{n}Q(\beta_0, \mathbf{y}) &= \text{plim}_{n \rightarrow \infty} \frac{1}{n} \mathbf{u}^\top \mathbf{P}_W \mathbf{u} && \text{(S8.02)} \\ &= \left(\text{plim}_{n \rightarrow \infty} \frac{1}{n} \mathbf{u}^\top \mathbf{W} \right) \left(\text{plim}_{n \rightarrow \infty} \frac{1}{n} \mathbf{W}^\top \mathbf{W} \right)^{-1} \left(\text{plim}_{n \rightarrow \infty} \frac{1}{n} \mathbf{W}^\top \mathbf{u} \right). \end{aligned}$$

A standard assumption for the IV estimator to be consistent is (8.16), according to which $\text{plim } n^{-1}\mathbf{W}^\top \mathbf{u} = \mathbf{0}$. It is also assumed that $\text{plim } n^{-1}\mathbf{W}^\top \mathbf{W} = \mathbf{S}_{\mathbf{W}^\top \mathbf{W}}$, which is a nonsingular matrix. Under these assumptions, the second line of (S8.02) implies that

$$\text{plim}_{n \rightarrow \infty} \frac{1}{n}Q(\beta_0, \mathbf{y}) = \mathbf{0}^\top \mathbf{S}_{\mathbf{W}^\top \mathbf{W}}^{-1} \mathbf{0} = 0.$$

This answers the first part of the question.

When $\mathbf{y} = \mathbf{X}\beta_0 + \mathbf{u}$, we see that

$$\mathbf{y} - \mathbf{X}\beta = \mathbf{X}\beta_0 + \mathbf{u} - \mathbf{X}\beta = \mathbf{u} + \mathbf{X}(\beta_0 - \beta).$$

Therefore, the criterion function can be written as

$$Q(\beta_0, \mathbf{y}) = \mathbf{u}^\top \mathbf{P}_W \mathbf{u} + 2\mathbf{u}^\top \mathbf{P}_W \mathbf{X}(\beta_0 - \beta) + (\beta_0 - \beta)^\top \mathbf{X}^\top \mathbf{P}_W \mathbf{X}(\beta_0 - \beta).$$

We have just seen that the plim of n^{-1} times the first term here is zero. By a similar argument, so is the plim of n^{-1} times the second term. Therefore,

$$\text{plim}_{n \rightarrow \infty} \frac{1}{n}Q(\beta_0, \mathbf{y}) = \text{plim}_{n \rightarrow \infty} \frac{1}{n}(\beta_0 - \beta)^\top \mathbf{X}^\top \mathbf{P}_W \mathbf{X}(\beta_0 - \beta).$$

This can be written as the product of three probability limits:

$$\left(\text{plim}_{n \rightarrow \infty} \frac{1}{n}(\beta_0 - \beta)^\top \mathbf{X}^\top \mathbf{W} \right) \left(\text{plim}_{n \rightarrow \infty} \frac{1}{n} \mathbf{W}^\top \mathbf{W} \right)^{-1} \left(\text{plim}_{n \rightarrow \infty} \frac{1}{n} \mathbf{W}^\top \mathbf{X}(\beta_0 - \beta) \right),$$

Using the definitions of $\mathbf{S}_{\mathbf{X}^\top \mathbf{W}}$, $\mathbf{S}_{\mathbf{W}^\top \mathbf{X}}$, and $\mathbf{S}_{\mathbf{W}^\top \mathbf{W}}$, we conclude that

$$\text{plim}_{n \rightarrow \infty} \frac{1}{n}Q(\beta_0, \mathbf{y}) = (\beta_0 - \beta)^\top \mathbf{S}_{\mathbf{X}^\top \mathbf{W}}(\mathbf{S}_{\mathbf{W}^\top \mathbf{W}})^{-1}\mathbf{S}_{\mathbf{W}^\top \mathbf{X}}(\beta_0 - \beta).$$

If the matrix $\mathbf{S}_{\mathbf{X}^\top \mathbf{W}}(\mathbf{S}_{\mathbf{W}^\top \mathbf{W}})^{-1}\mathbf{S}_{\mathbf{W}^\top \mathbf{X}}$ has full rank, as we have assumed, this is a positive definite quadratic form. Therefore, the probability limit of $n^{-1}Q(\beta_0, \mathbf{y})$ must be positive whenever $\beta \neq \beta_0$.