

Solution to Exercise 15.25

★15.25 For a given choice of bandwidth, the expectation of the estimate $\hat{f}_h(x)$ of (15.61) is h^{-1} times the expectation of the random variable $k((x - X)/h)$, where X denotes the random variable of which the x_t are IID realizations. Assume that k is symmetric about the origin. Show that the bias of $\hat{f}_h(x)$ is independent of the sample size n and roughly proportional to h^2 for small h . More formally, this means that the bias is $O(h^2)$ as $h \rightarrow 0$.

Show also that the variance of $\hat{f}_h(x)$ is of order $(nh)^{-1}$ as $n \rightarrow \infty$ and $h \rightarrow 0$. Why do these facts imply that the bandwidth h that minimizes the expectation of the squared error of $\hat{f}_h(x)$ must be of order $n^{-1/5}$ as $n \rightarrow \infty$?

The expectation of the kernel density estimator (15.61) is

$$\begin{aligned} E(\hat{f}_h(x)) &= h^{-1} E\left(k\left(\frac{x - X}{h}\right)\right) \\ &= h^{-1} \int_{-\infty}^{\infty} k\left(\frac{x - X}{h}\right) f(X) dX \\ &= \int_{-\infty}^{\infty} k(W) f(x - hW) dW, \end{aligned} \quad (\text{S15.49})$$

where $W \equiv (x - X)/h$. A second-order Taylor expansion of $f(x - hW)$ around the point $h = 0$ yields

$$f(x - hW) = f(x) - hWf'(x) + \frac{1}{2}h^2W^2f''(x) + O(h^3). \quad (\text{S15.50})$$

Substituting this into (S15.49), we find that $E(\hat{f}_h(x))$ is the sum of three integrals and a term of order h^3 . The first integral is

$$\int_{-\infty}^{\infty} k(W) f(x) dW = f(x) \int_{-\infty}^{\infty} k(W) dW = f(x).$$

The second integral is the negative of

$$\int_{-\infty}^{\infty} k(W) hWf'(x) dW = hf'(x) \int_{-\infty}^{\infty} Wk(W) dW = 0,$$

because the mean of W is 0. The third integral is

$$\int_{-\infty}^{\infty} k(W) \frac{1}{2}h^2W^2f''(x) = \frac{1}{2}h^2f''(x) \int_{-\infty}^{\infty} W^2k(W) dW = \frac{1}{2}h^2f''(x),$$

because the variance of W is 1. Thus we conclude that

$$E(\hat{f}_h(x)) - f(x) = \frac{1}{2}h^2f''(x) + O(h^3), \quad (\text{S15.51})$$

which is evidently $O(h^2)$.

Because the kernel density estimator (15.61) is an average of n IID random variables, its variance is n^{-1} times the common variance of these variables. That is,

$$\text{Var}(\hat{f}_h(x)) = \frac{1}{nh^2} \text{Var}(k(x - X/h)). \quad (\text{S15.52})$$

The variance on the right-hand side above can be computed as follows. We have

$$\begin{aligned} \text{Var}(k(x - X/h)) &= \int_{-\infty}^{\infty} k^2((x - X)/h) f(X) dX - \text{E}^2(k(x - X)/h) \\ &= h \int_{-\infty}^{\infty} k^2(W) f(x - hW) dW - \text{E}^2(k(x - X)/h). \end{aligned} \quad (\text{S15.53})$$

If we substitute the second-order Taylor expansion (S15.50) into the integral in expression (S15.53), we once again obtain three integrals and a term of order h^3 . The first integral is

$$\int_{-\infty}^{\infty} k^2(W) f(x) dW = f(x) \int_{-\infty}^{\infty} k^2(W) dW. \quad (\text{S15.54})$$

The second integral is the negative of

$$h \int_{-\infty}^{\infty} k^2(W) W f'(x) dW = h f'(x) \int_{-\infty}^{\infty} W k^2(W) dW = 0,$$

because the kernel function $k(W)$ is symmetric. The third integral is

$$\int_{-\infty}^{\infty} k^2(W) \frac{1}{2} h^2 W^2 f''(x) = \frac{1}{2} h^2 f''(x) \int_{-\infty}^{\infty} W^2 k^2(W) dW.$$

Since $\text{E}(k(x - X/h))$ is h times $\text{E}(\hat{f}_h(x))$, we find from (S15.51) and (S15.53), along with the above results, that

$$\text{Var}(k(x - X/h)) = h f(x) \int_{-\infty}^{\infty} k^2(W) dW - h^2 f^2(x) + O(h^3),$$

and so, from (S15.52), we see that

$$\text{Var}(\hat{f}_h(x)) = (nh)^{-1} \left(f(x) \int_{-\infty}^{\infty} k^2(W) dW - O(h) \right). \quad (\text{S15.55})$$

Since we are considering the limit where $h \rightarrow 0$, the term above that is $O(h)$ is asymptotically negligible relative to the other term, which is of order $(nh)^{-1}$, as we wished to show.

The squared error is the variance plus the square of the bias. Therefore, the results (S15.51) and (S15.55) imply that the approximate squared error is

$O(h^4) + O(n^{-1})O(h^{-1})$. If we differentiate this with respect to h and set the derivative to 0 to find the optimal value of h , we find that

$$O(n^{-1})O(h^{-2}) + O(h^3) = 0.$$

Multiplying by h^2 and rearranging, this becomes

$$O(n^{-1}) = -O(h^5),$$

from which we conclude that $h = O(n^{-1/5})$ if it is chosen optimally, as we were required to show.